

GENERATION OF RANDOM VARIATES
FROM STATISTICAL DISTRIBUTIONS

A Thesis
Presented to
the Faculty of the Department of
Industrial Engineering
University of Houston

In partial fulfillment
of the requirements for the Degree
Master of Science

by
Winifred Moon-Yee Ku

May, 1976

Dedicated to my beloved
Parents and Fourth Aunt

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ABSTRACT

Work has been done in generating random variates of some distributions for simulation purposes. A subroutine which gives a random number of any common statistical distribution desired, and a compilation of the backgrounds, and derivations for the generating methods gives convenience and understanding to anyone who does simulation work. This thesis presents a brief history on the generation of random variables, a compilation and investigation of old and new generating methods, and comparisons and evaluations of different methods by chi-square and t-tests. The best method of each distribution investigated is selected to incorporate into subroutines CTIME or DTIME, depending on whether the distribution is continuous or discrete.

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CHAPTER I

1. INTRODUCTION

There are, generally speaking, two different ways to generate uniform random numbers. One is the physical method, such as drawing numbered chips from a container. Another is the mathematical method. Since physical methods are inapplicable for simulation purposes, we shall not attempt to discuss them here. We shall discuss the development of the mathematical methods on uniform and non-uniform number generators alike.

The most significant study on number generators was done around 1950. The questions of concern at that time were: 1) how to produce a sequence of random decimal digits, and 2) how to produce random real numbers according to an assigned probability distribution law. Ulam suggested to use the mapping function $f(x)=4x(1-x)$ to solve the first problem. But Von Neumann, a mathematician (41) pointed out that the independence of the random numbers was lacking by the use of mapping functions. He also pointed out that even with a theoretically sound function, such as $f(x)=\sin^2 \pi x$ which satisfied all the properties for a uniform number generator, the limitation of the machines in rounding-off would destroy the randomness property. He evaluated other methods, but did not present an adequate answer. As to the second problem, he suggested the rejection method which

has one advantage in that the cumulative functions of the distributions need not to be used. This method is used specifically for distributions whose cumulative density functions do not exist in closed forms. In comparison with other methods, rejection takes more computing time.

Lehmer (19) was the first to study the congruential method for a uniform number generator. He used the relation

$$x_{i+1} = kx_i \pmod{M}$$

with $k=23$ and $M=10^8+1$. The sequence produced 8-decimal digit numbers with period 5882352. He also proposed that when δ is the smallest positive integer satisfying

$$\alpha^\delta = 1 \pmod{n}$$

where n and α are relatively prime positive integers, the sequence $u_i (i=1,2,\dots)$ formed by taking the principal remainder has a period of δ . The numbers $u_i n^{-1}$ may be used as uniform variates in the range $(0,1)$ provided that δ is reasonably large.

In 1954, the University of Florida conducted a symposium on Monte Carlo Methods. In the symposium, Metropolis presented his mid-square method for a uniform number generator.

O. Taussky and J. Todd tested both the additive congruential and the mid-square methods. E.J. Lytle Jr. presented tests for normal random variates, but he did not give any generating

method other than look-ups from normal tables and random number tables. Based on Votaw and Rufferty (42), James W. Butler gave the direct method and composition methods for transforming uniform random variates into variates of the desired distribution.

Lehmer's method was further carefully studied by a few people, among whom were Eva and V.J. Bofinger. They gave a general way of evaluating δ and also the results of 4 tests for randomness: frequency test, serial test, gap test and uniform test. J. Certainé concerned himself in finding the value of α which will generate a non-repeated sequence of maximum length.

Among all the non-uniform distributions, the normal distribution was studied first and most intensively. There were many generating methods suggested. Muller designed his Chebyshev approximation (29), but it required the storage of large tables of values which contained the coefficients of the polynomials of the approximated functions. His method is rather cumbersome. He also had a study on the direct method. Von Neumann had his rejection method applied to the normal distribution. Hasting had his approximation, and there was also the central limit approximation approach. Teichroew had his Chebyshev polynomial approximation which did not require the storage of any table, therefore it worked better than that of Muller's.

G. Marsaglia, M.D. MacLaren and T.A. Bray (21) designed their method by breaking up the normal density function into a sum of three density functions. L.E. Cannon did a detail study on this in his Masters thesis. He included normal and exponential distributions, and also the computer table look-up method for discrete distribution;.

The third question of how to adequately test generators came into the picture. In 1965 (20), M.D. MacLaren and G. Marsaglia proposed that the tests commonly applied for testing were not of much value. They conducted their own tests on uniform generators. They were the uniformity, pairs, triples, sum of 2, sum of 4, maximum of 2, maximum of 5, minimum of 3 and minimum of 10. The generators being tested were the mixed congruential, the combination of two congruentials and the random number table.

The most recent work on testing methods was A.J. Sacks' Masters thesis (34). His work was a program which included the gap, runs, pairs, chi-square, moments, runs above-below mean, autocorrelations and Kolmogorov-Smirnov tests.

Several books have been written trying to put all the generating methods together. Examples are (30) and (39).

It can be seen that random number generators have been studied a great deal by many. Upon examining the papers and books written, it was found that no satisfactory work has been done to include descriptions of all the methods known,

their derivations and computer programs. Since computer simulation is very commonly used nowadays, a subroutine which gives random values from any statistical distribution will ease the work of anyone who does simulations. This motivates this study. We attempt, in this thesis, to provide a quick understanding of the generators. Mathematical proofs are provided for most methods. For those derivations which required lengthy discussions, adequate references are supplied. The programs, which were all written in FORTRAN IV are listed for usage. Our work here is more or less a continuation of A.J. Sacks' work. Therefore, discussions on uniform generators were omitted. We used a multiplicative congruential method for our uniform generator, and the function subroutine is from C.E. Donaghey who used it in his simulation course. Each generating method was evaluated either by a chi-square test or by the t-test. The execution times, determined on UNIVAC 1108, of different methods were computed.

It seems that the study of random number generators has reached its final stage with all these methods found and all these testings done. Yet, it is not so, studies still have to be done on multi-variate generators.

CHAPTER II

This chapter contains the definition of each distribution, descriptions and derivations of the generating methods, their program listings and testing results.

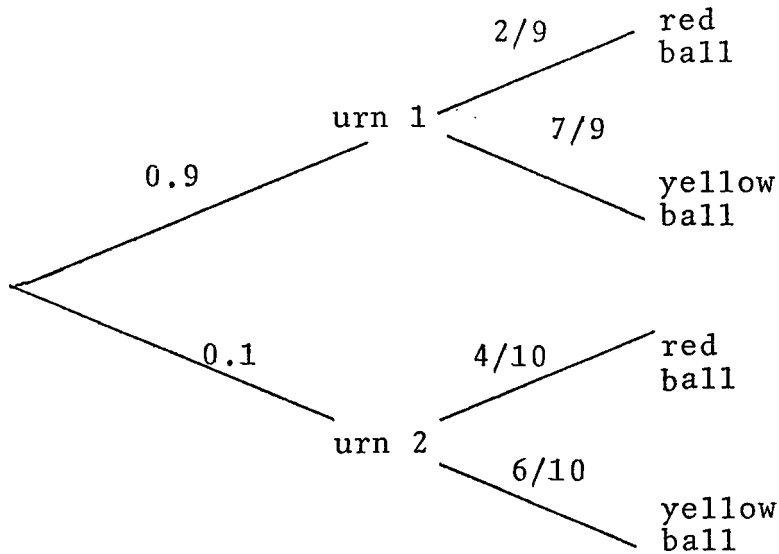
1. DISCRETE DISTRIBUTIONS

The following techniques are methods to instruct a computer to choose a random value from a discrete distribution with assigned probabilities.

A. MARSAGLIA METHOD

If drawing a red ball has a probability of 0.24 and drawing a yellow ball has a probability of 0.76, in order to get a ball at random, one can draw a ball from an urn containing 24 red and 76 yellow balls. To do this with a computer program, one can store 24 zero's in memory locations 1 to 24 and 76 one's in locations 25 to 100. A random integer between 1 and 100 is selected. If the integer falls between 1 and 24, the returned value would be a zero indicating that a red ball is selected. Otherwise, a yellow ball is selected for a returned value between 25 and 100. One hundred memory locations are required for this example. When the probabilities are given in 3 places after the decimal, 1000 memory locations are required and so forth. In order to conserve memory locations one can use two urns instead of one. For our example, 2 red balls and 7 yellow balls are put in urn one; while 4 red balls and 6 yellow balls are put in urn two. 0.2 and 0.7 are added together to obtain 0.9 which is the probability that urn one would be selected. 0.04 and 0.06 are added together to obtain 0.10 which is the probability that urn two would be selected. With a look at the tree diagram, it is obvious that the

probabilities of red and yellow balls being selected remain unchanged.



$$\begin{aligned}
 P(\text{red ball}) &= P(\text{urn 1}) * P(\text{red ball} | \text{urn 1}) \\
 &\quad + P(\text{urn 2}) * P(\text{red ball} | \text{urn 2}) \\
 &= 0.9 * 2/9 + 0.1 * 4/10 \\
 &= 0.24
 \end{aligned}$$

$$\begin{aligned}
 P(\text{yellow ball}) &= P(\text{urn 1}) * P(\text{yellow ball} | \text{urn 1}) \\
 &\quad + P(\text{urn 2}) * P(\text{yellow ball} | \text{urn 2}) \\
 &= 0.9 * 7/9 + 0.1 * 6/10 \\
 &= 0.76
 \end{aligned}$$

With this slight change of doing one more step to select the urn, 100 memory locations are drastically reduced to 19 locations. Based on the above logic, Marsaglia has designed a method for obtaining a random number from any discrete distribution. It can easily be understood from the example on the next page.

Example 1.A.1:

Let the values of x_1, x_2, x_3, x_4, x_5 and x_6 be 0,1,2,3,4 and 5.

Let their corresponding probabilities $P(x_1), P(x_2), P(x_3), P(x_4), P(x_5)$ and $P(x_6)$ be 0.6250, 0.2047, 0.0813, 0.0651, 0.0134 and 0.0105.

Contents are loaded in the memory locations of the computer according to the following scheme:

Loc	Con	Loc	Con	Loc	Con	Loc	Con
1	0	17	2	33	1	49	1
2	0	18	2	34	1	50	1
3	0	19	3	35	1	51	1
4	0	20	3	36	2	52	2
5	0	21	3	37	3	53	2
6	0	22	3	38	3	54	2
7	1	23	3	39	3	55	3
8	1	24	3	40	3	56	4
9	0	25	4	41	3	57	4
10	0	26	5	42	4	58	4
11	2	27	0	43	4	59	4
12	2	28	0	44	4	60	5
13	2	29	0	45	1	61	5
14	2	30	0	46	1	62	5
15	2	31	0	47	1	63	5
16	2	32	1	48	1	64	5

Notice that each probability is a number up to four places after the decimal. The first digit after the decimal of $P(x_1)$ is a 6, therefore, 6 zeros were put in locations 1 to 6. The first digit after the decimal of $P(x_2)$ is a 2, therefore, 2 ones were put in locations 7 and 8. After all the first digits of all the probabilities were considered and contents loaded, then the second, the third and the fourth digits were considered and contents were loaded in

the same way.

Let S_i be the sum of the digits of all the probabilities at the i th place after the decimals.

In our example:

$$S_1 = 6+2+0+0+0+0 = 8$$

$$S_2 = 2+0+8+6+1+1 = 18$$

$$S_3 = 5+4+1+5+3+0 = 18$$

$$S_4 = 0+7+3+1+4+5 = 20$$

Let d_i be the i th digit after the decimal of a number between zero and one.

To obtain a random number from this example:

1) Generate a uniform random number $U = .d_1 d_2 d_3 d_4 \dots$

2) Test the following:

If $d_1 < S_1 = 8$, the desired random number is the content of location d_1 .

If $80 \leq d_1 d_2 < 10S_1 + S_2 = 98$, the desired random number is the content of location $(d_1 d_2 - 80 + S_1 + 1)$.

If $980 \leq d_1 d_2 d_3 < 100S_1 + 10S_2 + S_3 = 998$, the desired random number is the content of location $(d_1 d_2 d_3 - 980 + S_1 + S_2 + 1)$.

If $1000S_1 + 100S_2 + 10S_3 + S_4 \leq d_1 d_2 d_3 d_4 = 9980$, the desired random number is the content of location $(d_1 d_2 d_3 d_4 - 9980 + S_1 + S_2 + S_3 + 1)$.

PROGRAM LISTING

```

C*****
C
C      TO RETURN A RANDOM VARIATE OF ANY DISCRETE
C      DISTRIBUTION BY MARSAGLIA METHOD
C
C      P = INPUT VECTOR WITH THE VARIABLES AND THEIR
C          PROBABILITIES IN THE ORDER OF X(1),P(1),X(2),
C          P(2),...
C
C      ND = THE DIMENSION OF VECTOR P
C
C      VAL = THE RETURNED VARIATE
C*****
C
      SUBROUTINE DISC(P,ND,VAL)
      DIMENSION P(110),CON(10000),INDX(4),NO(4,110) ISUM(4)
      DATA NSEED/567801/
      K=ND/2
      M=1
      DO 9 I=1,K
      P(K+1)=P(K+I)+0.00005
      NO(1,I)=P(K+I)*10
      NO(2,I)=P(K+I)*100-NO(1,I)*10
      NO(3,I)=P(K+I)*1000-NO(1,I)*100-NO(2,I)*10
      NO(4,I)=P(K+I)*10000-NO(1,I)*1000-NO(2,I)*100-NO(3,I)*10
9  CONTINUE
      DO 5 N=1,4
      ISUM(N)=0
      INDX(N)=M
      DO 3 J=1,K
      NUM=NO(N,J)
      IF(NUM.EQ.0) GO TO 3
      IF(M.EQ.1) LK=M+NUM-1
      IF(M.GT.1) LK=M+NUM
      DO 4 L=M,LK
      CON(L)=P(J)
4  CONTINUE
      M=M+NUM
3  CONTINUE
5  CONTINUE
      DO 87 I=1,K
      ISUM(1)=ISUM(1)+NO(1,I)
      ISUM(2)=ISUM(2)+NO(1,I)*10+NO(2,I)
      ISUM(3)=ISUM(3)+NO(1,I)*100+NO(2,I)*10+NO(3,I)
      ISUM(4)=ISUM(4)+NO(1,I)*1000+NO(2,I)*100+NO(3,I)+10+NO(4,I)
87 CONTINUE

```

```
T1=ISUM(1)/10.
T2=ISUM(2)/100.
T3=ISUM(3)/1000.
T4=ISUM(4)/10000.
U=RAN(NSEED)
IF(U.GE.0...AND.U.LT.T1) IX=U*10+INDX(1)
IF(U.GE.T1.AND.U.LT.T2) IX=U*100-ISUM(1)*10+INDX(2)
IF(U.GE.T2.AND.U.LT.T3) IX=U*1000-ISUM(2)*10+INDX(3)
IF(U.GE.T3.AND.U.LT.T4) IX=U*10000-ISUM(3)*10+INDX(4)
VAL=CON(IX)
DO 99 J=1,K
P(K+J)=P(K+J)-0.00005
99 CONTINUE
RETURN
END
```

B. CUMULATION PROBABILITY METHOD

With example 1.A. 1:

variable	probability	cumulative probability
x_1	0.6250	0.6250
x_2	0.2047	0.8297
x_3	0.0813	0.9110
x_4	0.0651	0.9761
x_5	0.0134	0.9895
x_6	0.0105	1.0000

To obtain a discrete random number, generate a uniform random number U in $(0,1)$, if U is less or equal to the cumulative probability of x_i , $i=1, 2, 3, 4, 5$ or 6 ; then the desired number is x_i .

PROGRAM LISTING

```

C*****
C
C      TO RETURN A RANDOM VARIATE OF ANY DISCRETE BY
C      CUMULATIVE PROBABILITY METHOD
C
C      P = INPUT VECTOR WITH THE VARIABLES AND THEIR
C          PROBABILITIES IN THE ORDER OF X(1),P(1),x(2),
C          P(2),...
C
C      ND = THE DIMENSION OF VECTOR P
C
C      VAL = THE RETURNED VARIATE
C*****
C
C      SUBROUTINE DISC(P,ND,VAL)
C      DIMENSION P(100),PM(100)
C      DATA NSEED/567801/
C      N=ND/2+1
C      K=ND/2+2
C      R=RAN(NSEED)
C      PM(K-1)=P(K-1)
C      DO 1 I=K,ND
C      PM(I)=P(I)+PM(I-1)
1  CONTINUE
C      DO 2 J=N,ND
C      IF(R.GE.0.0.AND.R.LT.PM(J)) GO TO 3
2  CONTINUE
3  VAL=P(J-K+2)
C      RETURN
C      END

```

PROGRAM TESTING RESULTSExample 1.A.1

i	x_i	MARSAGLIA $P(x_i)$		CDF METHOD $P(x_i)$		GIVEN $P(x_i)$
		N=100	N=1000	N=100	N=1000	
1	0	0.6400	0.6440	0.6200	0.6450	0.6250
2	1	0.1900	0.2090	0.2000	0.2080	0.2047
3	2	0.0600	0.0590	0.0800	0.0610	0.0813
4	3	0.1000	0.0750	0.0800	0.0620	0.0651
5	4	0.0000	0.0060	0.0100	0.0130	0.0134
6	5	0.0100	0.0070	0.0100	0.0110	0.0105

Example 1.A.2

i	x_i	MARSAGLIA $P(x_i)$		CDF METHOD $P(x_i)$		GIVEN $P(x_i)$
		N=100	N=1000	N=100	N=1000	
1	1	0.1800	0.2270	0.2100	0.2320	0.2167
2	2	0.0300	0.0180	0.0400	0.0190	0.0192
3	3	0.4500	0.4180	0.3800	0.4280	0.4201
4	4	0.0700	0.0490	0.0500	0.0490	0.0480
5	5	0.1300	0.1710	0.1900	0.1520	0.1623
6	6	0.0300	0.0230	0.0100	0.0230	0.0355
7	7	0.1000	0.0770	0.1000	0.0780	0.0802
8	8	0.0100	0.0170	0.0200	0.0190	0.0180

N is the number of random variates generated for testing.

METHODS	EXECUTION TIME		IN MILLI-SECONDS	
	Example 1.A.1		Example 1.A.2	
	N=100	N=1000	N=100	N=1000
MARSAGLIA	175	1561	46	362
CDF METHOD	22	151	21	148

SUMMARY OF CHI-SQUARE TEST

METHODS	N	Example 1.A.1		Example 1.A.2	
		degrees of freedom	chi-square values	degrees of freedom	chi-square values
MARSAGLIA	1000	5	13.543	7	5.647
	2500	5	5.445	7	3.730
	5000	5	3.623	7	3.192
CDF METHOD	1000	5	5.945	7	6.423
	2500	5	4.648	7	4.536
	5000	5	3.223	7	7.287

<u>Degrees of freedom</u>	<u>α</u>	<u>Critical χ^2</u>
5	0.05	11.0706
5	0.10	9.2363
7	0.05	14.0671
7	0.10	12.0170

N is the number of random variates generated for testing.

2. HYPERGEOMETRIC DISTRIBUTION

Suppose there are N items, N_p of which belongs to class I. and N_q of which belongs to class II. p is the probability that an item of class I will be chosen, and q is the probability that an item of class II will be chosen, where $p+q=1$. Suppose n items were chosen at random from the lot ($n \leq N$), without replacement, Let X be the number in class I found. Since $X=x$ if and only if we obtained precisely x class I items and precisely $(n-x)$ class II items, we have

$$f(x) = P(X=x) = \frac{\binom{N_p}{x} \binom{N_q}{n-x}}{\binom{N}{n}}, \quad x=0,1,2,\dots$$

A discrete random variable having a probability distribution as above is said to have a hypergeometric distribution. The expected value and variance of x are:

$$EX = np$$

$$VX = npq \left(\frac{N-n}{N-1} \right)$$

A. DIRECT METHOD

The steps to obtain a hypergeometric variate with N , n and p given are:

- 1) Set $x = 0$.
- 2) Generate a uniform variate R in $(0,1)$.
- 3) If R is less or equal to p , increase x by 1, decrease n by 1 and recalculate p .
- 4) If R is greater than one, decrease n by 1.
- 5) Go back to step 2 if n is greater than zero.
- 6) Stop if n is equal to zero, x is the required variate.

PROGRAM LISTING

```

C*****
C
C      TO RETURN A HYPERGEOMETRIC RANDOM VARIATE BY
C      DIRECT METHOD
C
C      TN = TOTAL NUMBER OF ITEMS
C
C      NC = NUMBER OF ITEMS CHOSEN
C
C      P  = PROBABILITY THAT AN ITEM OF CLASS I WILL BE
C           CHOSEN
C
C      VAL=THE RETURNED HYPERGEOMETRIC VARIATE
C*****
C
C      SUBROUTINE HYPGEO(TN,NC,P,VAL)
C      DATA NSEED/567801/
C      PX=P
C      TNX=TN
C      VAL=0.0
C      DO 11 I=1,NC
C      R=RAN(NSEED)
C      IF(R-PX) 6,6,9
C 6  S=1.0
C      VAL=VAL+1.0
C      GO TO 10
C 9  S=0.0
C 10 PX=(TNX*PX-S)/(TNX-1.0)
C      TNX=TNX-1.0
C 11 CONTINUE
C      RETURN
C      END

```

PROGRAM TESTING RESULTS

METHOD	PROGRAM RESULTS				GIVEN	
	Mean		Variance		Mean	Variance
	N=100	N=1000	N=100	N=1000		
DIRECT	1.900	1.966	0.530	0.623	2.0	0.666

The given parameters of this example are:

TN = 10.0

NC = 5

P = 0.4

METHOD	EXECUTION TIME IN MILLI-SECONDS	
	N=100	N=1000
DIRECT	57	490

N is the number of random variates generated for testing. Notice that the execution time of this method depends on the input parameters.

SUMMARY OF CHI-SQUARE TEST

METHOD	N	Degrees of freedom	Chi-square values
DIRECT	50	2	3.3405
	100	2	3.7620
	150	2	2.6880

<u>Degrees of freedom</u>	<u>α</u>	<u>Critical χ^2</u>
2	0.05	2.7055
2	0.10	4.6052

N is the number of random variates generated for testing.

3. GEOMETRIC DISTRIBUTION

Let A be an event. Define the random variable X as the number of repetitions required up to and including the first occurrence of A. Thus X assumes the possible values 1,2,... . Since $X=x$ if and only if the first $(x-1)$ th repetitions of the experiment results not in A while the x th repetition results in A, we have

$$f(x) = P(X=x) = q^{x-1}p, \quad x=1,2,\dots$$

A random variable with probability distribution as above is said to have a geometric distribution.

The cumulative distribution function is

$$F(x) = P(X \leq x) = \sum_{i=0}^x pq^i$$

The expected value and variance of x are:

$$EX = q/p$$

$$VX = q/p^2 = EX/p$$

p is the probability that A occurs, and $q=1-p$.

A. NAYLOR METHOD

To obtain a geometric distributed variate by the Naylor method, the steps are:

- 1) Generate a uniformly distributed random number U in $(0,1)$.
- 2) Calculate $x = \text{Log}_e U / \text{Log}_e q$, x is the required variate.

Verification:

$$\begin{aligned}
 & F(0) = P(x=0) = \sum_{i=0}^0 pq^i = p \\
 \text{and} \quad & F(x) = P(X \leq x) \\
 \text{therefore} \quad & p \leq F(x) \leq 1 \\
 & P(X > x) = 1 - F(x) \\
 \text{implies} \quad & P(X > 0) = 1 - F(0) \\
 & = 1 - p \\
 & = q
 \end{aligned}$$

Consider $1 - F(x)$

$$\text{Let } S = 1 - F(x) = 1 - p - pq - pq^2 - \dots - pq^x$$

$$\begin{aligned}
 - \quad & qS = q - pq - pq^2 - \dots - pq^x - pq^{x+1} \\
 \hline
 & S - qS = 1 - q - p + pq^{x+1} \\
 & = p - p + pq^{x+1}
 \end{aligned}$$

$$S(1 - q) = pq^{x+1}$$

$$pS = pq^{x+1}$$

$$S = q^{x+1}$$

$$\text{therefore } 1 - F(x) = q^{x+1}$$

$$\begin{aligned}
 & p \leq F(x) \leq 1 \\
 \text{implies} \quad & -p \geq -F(x) \geq -1 \\
 & 1-p \geq 1-F(x) \geq 1-1 \\
 & q \geq q^{x+1} \geq 0 \\
 & 1 \geq q^x \geq 0
 \end{aligned}$$

substituting q^x by $(1-F(x))/q$, shows that $(1-F(x))/q$ has unit length.

$$\text{Let } U = (1-F(x))/q$$

$$U = q^x$$

$$\text{Log}_e U = x \text{Log}_e q$$

$$x = \frac{\text{Log}_e U}{\text{Log}_e q}$$

PROGRAM LISTING

```

C*****
C
C      TO RETURN A GEOMETRIC RANDOM VARIATE BY NAYLOR
C      METHOD
C
C      P = THE PROBABILITY THAT THE EVENT WILL OCCUR
C
C      VAL = THE RETURNED GEOMETRIC VARIATE
C
C*****
C
C      SUBROUTINE GEOM(P,VAL)
C      DATA NSEED/567801/
C      R=RAN(NSEED)
C      VAL=ALOG(R)/ALOG(1.-P)
C      RETURN
C      END

```

B. DIRECT METHOD

This method is based on the definition of a geometric variate.

Steps to obtain a geometric random variate:

- 1) Set $x=0$.
- 2) Generate a uniform random variate R in $(0,1)$.
- 3) If R is greater than the given p , the probability for the occurrence of the event, then increase x by 1 and go back to step 2.
- 4) If R is equal to or smaller than p , x is the required geometric variate.

PROGRAM LISTING

```

C*****
C
C      TO RETURN A GEOMETRIC RANDOM VARIATE BY DIRECT
C      METHOD
C
C      P = THE PROBABILITY THAT THE EVENT WILL OCCUR
C
C      VAL = THE RETURNED GEOMETRIC VARIATE
C
C*****
C
C      SUBROUTINE GEOM(P,VAL)
C      DATA NSEED/567801/
C      X=0.0
C      4 R=RAN(NSEED)
C      IF(R.LE.P) GO TO 3
C      X=X+1.0
C      GO TO 4
C      3 VAL=X
C      RETURN
C      END

```

PROGRAM TESTING RESULTS

METHODS	PROGRAM RESULTS				GIVEN	
	Mean		Variance		Mean	Variance
	N=100	N=1000	N=100	N=1000		
NAYLOR	2.610	2.473	7.173	6.419	2.03	6.152
DIRECT	1.600	1.917	3.480	5.724	2.03	6.152

The parameter given in this example is:

$$P = 0.33$$

METHODS	EXECUTION TIME IN MILLI-SECONDS	
	N=100	N=1000
NAYLOR	23	265
DIRECT	20	161

N is the number of random variates generated for testing. Notice that the execution time of the Direct method depends on the input parameter , while the Naylor method does not.

SUMMARY OF CHI-SQUARE TEST

METHODS	N	Degrees of freedom	Chi-square values
NAYLOR	50	3	11.442
	100	5	12.847
	150	5	18.922
DIRECT	50	5	4.228
	100	5	1.369
	150	7	1.402

<u>Degrees of freedom</u>	<u>α</u>	<u>Critical χ^2</u>
3	0.05	7.81473
3	0.10	6.25139
5	0.05	11.07050
5	0.10	9.23635
7	0.05	14.06710
7	0.10	12.01700

N is the number of random variates generated for testing.

4. PASCAL DISTRIBUTION

Let A be an event. Define the random variable X as the number of failures when the trials are repeated until k successes occur ($k > 1$). Then X has a negative binomial or Pascal distribution.

$$f(x) = P(X=k+x) = \binom{k+x-1}{x-1} p^k q^x \quad x=0,1,2,\dots$$

where p is the probability of the occurrence of A, and $q=1-p$, the probability of the failure of A.

The mean and the variance of x are:

$$EX = kq/p$$

$$VX = kq/p^2 = EX/p \quad .$$

A. SUM GEOMETRIC METHOD

Notice that when $k = 0$, $f(x) = pq^x$.

Therefore when k is an integer, a pascal variate is the sum of k geometric variates.

Let G be a geometric variate, the U be a uniform variate in $(0,1)$; if X is a pascal variate,

$$X = \sum_{n=1}^k G_n$$

$$X = \frac{\sum_{n=1}^k \text{Log}_e U_n}{\text{Log}_e q}$$

$$X = \frac{\text{Log}_e \prod_{n=1}^k U_n}{\text{Log}_e q}$$

To obtain a geometric random variate:

- 1) Generate k random uniform variates in $(0,1)$ and calculate their product, say it is Z .
- 2) Calculate $X = \text{Log}_e Z / \text{Log}_e q$, X is the required variate.

PROGRAM LISTING

```

C*****
C
C      TO RETURN A RANDOM PASCAL VARIATE BY SUM GEOMETRIC *
C      METHOD *
C *
C      K = THE NUMBER OF SUCCESSES *
C *
C      Q = THE PROBABILITY THAT THE EVENT WILL NOT OCCUR *
C *
C      VAL = THE RETURNED PASCAL VARIATE *
C *
C*****
C
      SUBROUTINE PASCAL(K,Q,VAL)
      DATA NSEED/567801/
      TR=0.0
      QR=ALOG(Q)
      DO 184 I=1,K
      TR=TR+ALOG(RAN(NSEED))
184  CONTINUE
      IX=TR/QR
      X=IX
      RETURN
      END

```

Notice that in order to avoid getting an argument too small for the logarithmic function of the computer, in our program, the sum of the log of k uniform variates is used instead of the log of the product of k uniform variates.

B. NEGATIVE BINOMIAL METHOD

The pascal distribution is also known as the negative binomial distribution. Therefore, we can use the opposite reasoning of generating a binomial variate to generate a pascal variate.

- Steps to obtain a random pascal variate:

Let N be the total number of trials.

Let M be the number of successful trials.

1) Set N and M to be zeros.

2) Generate R , a uniform random variate in $(0,1)$ and increase N by 1.

3) If R is greater than q , the trial is successful, so increase M by 1.

4) If M is equal to k , it means that k successes has already occurred, then go to step 6.

5) Go to step 2.

- 6) The required pascal variate is equal to $N-M$.

PROGRAM LISTING

```

C*****
C
C      TO RETURN A PASCAL RANDOM VARIATE BY NEGATIVE
C      BINOMIAL METHOD
C
C      K = THE NUMBER OF SUCCESSES
C
C      Q = THE PROBABILITY THAT THE EVENT WILL NOT OCCUR
C
C      VAL = THE RETURNED PASCAL VARIATE
C*****
C
      SUBROUTINE PASCAL(K,Q,VAL)
      DATA NSEED/567801/
      N=0
      M=0
1  R=RAN(NSEED)
      N=N+1
      IF(R.GT.Q) M=M+1
      IF(M.EQ.K) GO TO 2
      GO TO 1
2  X=N-M
      RETURN
      END

```

C. TOTAL TRIAL METHOD

A Pascal variate can be defined as the total number of trials such that x failures occurred before k successes occurred.

The variate is now $k+x$ instead of x .

The mean of $k+x$ is:

$$\begin{aligned}
 EX(x+k) &= EX(x) + EX(k) \\
 &= \frac{kq}{p} + k \\
 &= \frac{kq+kp}{p} \\
 &= \frac{k(1-q+q)}{p} \\
 &= \frac{k}{p}
 \end{aligned}$$

The variance of $k+x$ is:

$$\begin{aligned}
 VX(k+x) &= EX(k+x)^2 - EX^2(k+x) \\
 &= EX(k^2+2kx+x^2) - \frac{k^2}{p^2} \\
 &= k^2+2kEX(x)+EX(x^2) - \frac{k^2}{p^2} \\
 &= k^2 + \frac{2k^2}{p} + \frac{kq}{p^2} + \frac{kq}{p} - \frac{k^2}{p^2}
 \end{aligned}$$

- Steps to obtain a random pascal variate:
- Let N be the total number of trials.
- Let M be the number of successful trials.
- 1) Set N and M to be zeros.
 - 2) Generate R, a uniform random variate in (0,1) and increase N by 1.
 - 3) If R is greater than q, the trial is successful, so increase M by 1.
 - 4) If M is equal to k, it means that k successes have already occurred, then go to step 6.
 - 5) Go to step 2.
 - 6) The required Pascal variate is equal to N.

PROGRAM LISTING

```

C*****
C
C      TO RETURN A RANDOM PASCAL VARIATE BY TOTAL TRIAL
C      METHOD
C
C      K = THE NUMBER OF SUCCESSES
C
C      Q = THE PROBABILITY THAT THE EVENT WILL NOT OCCUR
C
C      VAL = THE RETURNED PASCAL VARIATE
C*****
C
C      SUBROUTINE PASCAL(K,Q,VAL)
C      DATA NSEED/567801/
C      N=0
C      M=0
C      1 R=RAN(NSEED)
C      N=N+1
C      IF(R.GT.Q) M=M+1
C      IF(M.EQ.K) GO TO 2
C      GO TO 1
C      2 X=N
C      RETURN
C      END

```


D. EXPONENTIAL METHOD

Since a geometric variable is defined as the number of repetitions required up to and including the first occurrence of an event, it can be considered as the units of waiting time before the first arrival. Therefore, the geometric distribution is the discrete analogue of the exponential distribution. The conditional probability that the waiting time terminates at the x th trial is $q^x p$ where p is the probability that the waiting time terminates at the first trial.

Suppose each trial takes a time Δt , so that x trials take time $x\Delta t$.

Let $t \rightarrow 0$ and $x \rightarrow \infty$, so that the total time $t = x\Delta t$ remains constant.

If p and q were constants, the first success will occur sooner or later and in the limit the waiting time would be zero. In order to have the waiting time not to be zero, let $p \rightarrow 0$ and $q \rightarrow 1$.

The mean of the geometric variate is q/p , and the mean of the waiting time is $q\Delta t/p$.

Let $q\Delta t/p = 1/\lambda$ be a constant.

Since $p \rightarrow 0$ and $q \rightarrow 1$, we have

$$\begin{aligned} p &\sim p/q \\ &\sim \frac{p}{q\Delta t} \Delta t \\ &\sim \lambda \Delta t \end{aligned}$$

Substituting $x\Delta t=t$, and $p \sim \lambda \Delta t$ into the $f(x)$ of Pascal, we have

$$\begin{aligned} & \binom{k + \frac{t}{\Delta t} - 1}{\frac{t}{\Delta t} - 1} (1 - \lambda \Delta t)^{t/\Delta t} (\lambda \Delta t)^k \\ &= e^{-\lambda t} \frac{(\lambda t)^{k-1}}{(k-1)!} \end{aligned}$$

which is the Poisson expression for the probability of exactly $k-1$ events within time t .

Let $k=1$ and $\Delta t=1$;

$$\lambda \sim \Delta t \frac{p}{q}$$

$$\text{implies } \lambda \sim \frac{1-q}{q}$$

We know that the exponential distribution is a special case of the gamma distribution when the scale parameter is one. Therefore, to obtain the Poisson parameter, say α , for the Pascal, we generate a gamma variate with shape parameter λ and scale parameter k .

To obtain a Pascal random variate:

- 1) Calculate λ .
- 2) Call subroutine gamma for α .
- 3) Call subroutine Poisson for Pascal random variate x .

PROGRAM LISTING

```

C*****
C
C      TO RETURN A RANDOM PASCAL VARIATE BY EXPONENTIAL      *
C      METHOD                                                  *
C
C      K = THE NUMBER OF SUCCESS                             *
C
C      Q = THE PROBABILITY THAT THE EVENT WILL NOT OCCUR      *
C
C      VAL = THE RETURNED PASCAL VARIATE                      *
C*****
C
C      SUBROUTINE PASCAL(K,Q,VAL)
C      DATA NSEED/567801/
C      ADK=(1.-Q)/Q
C      CALL GAMMA(K,ADK,0.0,T)
C      CALL POIS(T,Y)
C      VAL+Y
C      RETURN
C      END

```

E. NON-INTEGGER NUMBER OF SUCCESSES

Let d_1 be the unit digit and d_2 be the first digit after the decimal of k . Then $k = d_1.d_2$. We can generate $(10-d_2)$ Pascal variates with the number of successes equal to d_1 ; and d_2 Pascal variates with the number of successes equal to (d_1+1) . The expected value of the number of successes is:

$$\begin{aligned}
 & \frac{(10-d_2)*d_1 + d_2*(d_1+1)}{10} \\
 &= \frac{10d_1 - d_2d_1 + d_2d_1 + d_2}{10} \\
 &= \frac{10d_1 + d_2}{10} \\
 &= d_1.d_2 \quad \text{which is equal to } k.
 \end{aligned}$$

To obtain a random Pascal variate when k is given as $d_1.d_2$:

- 1) Generate $10-d_2$ Pascal variates with the number of successes equal to d_1 .
- 2) Generate d_2 Pascal variates with the number of successes equal to d_1+1 .
- 3) Divide the sum of the variates by 10. The result is the required Pascal variate.

PROGRAM LISTING

```

C*****
C
C      TO RETURN A RANDOM PASCAL VARIATE OF NON-INTEGER
C      NUMBER OF SUCCESSES
C
C      PK = THE NUMBER OF SUCCESSES
C
C      Q = THE PROBABILITY THAT THE EVENT WILL NOT OCCUR
C
C      VAL = THE RETURNED PASCAL VARIATE
C*****
C
      SUBROUTINE PASCAL(PK,Q,VAL)
      DATA NSEED/567801/
      N=PK
      K=(PK-N)*10
      SUM=0.0
      NN=0
      II=0
      L=N+1
      DO 2 I=1,K
3    R=RAN(NSEED)
      IF(R.GT.Q) II=II+1
      NN=NN+1
      IF(II.EQ.L) GO TO 4
      GO TO 3
4    SUM=SUM+(NN-II)
      NN=0
      II=0
2    CONTINUE
      KK=10-K
      DO 12 I=1,KK
13   R=RAN(NSEED)
      IF(R.GT.Q) II=II+1
      NN=NN+1
      IF(II.EQ.N) GO TO 14
      GO TO 13
14   SUM=SUM+(NN-II)
      NN=0
      II=0
12  CONTINUE
      VAL=SUM/10.
      RETURN
      END

```

PROGRAM TESTING RESULTS

METHODS	N	PROGRAM RESULTS		GIVEN	
		Mean	Variance	Mean	Variance
SUM GEOM	100	22.920	137.154	21.316	112.188
	1000	23.633	113.136	21.316	112.188
NEG BINOM.	100	20.560	121.146	21.316	112.188
	1000	21.144	115.695	21.316	112.188
TOTAL TRIAL	100	13.830	20.661	14.286	19.730
	1000	14.054	16.749	14.286	19.730
EXPONENTIAL	100	19.030	85.169	21.316	112.188
	1000	21.277	116.715	21.316	112.188
NON-INTEGERS	100	2.417	0.364	2.607	4.495
	1000	2.532	0.431	2.607	4.495

The input parameters of the example for SUM GEOM, NEG BINOMIAL and EXPONENTIAL methods are:

$$\begin{aligned} K &= 5.0 \\ Q &= 0.81 \end{aligned} \quad (\text{see pages 30 \& 32})$$

The input parameters of the example for TOTAL TRIAL method are:

$$\begin{aligned} K &= 6.0 \\ Q &= 0.58 \end{aligned} \quad (\text{see page 34})$$

The input parameters of the example for NON-INTEGERS method are:

$$\begin{aligned} PK &= 3.60 \\ Q &= 0.42 \end{aligned} \quad (\text{see page 39})$$

N is the number of random variates generated for testing.

METHODS	EXECUTION TIME IN MILLI-SECONDS	
	N=100	N=1000
SUM NORM	72	684
NEG BINOM	203	1983
TOTAL TRIAL	135	1126
EXPONENTIAL	242	2248
NON-INTEGERS	314	3811

N is the number of random variates generated for testing. Notice that the execution times of SUM NORM, NEG BINOMIAL, TOTAL TRIAL and NON-INTEGERS methods all depend on the input parameters.

SUMMARY OF CHI-SQUARE TEST

METHODS	N	Degrees of freedom	Chi-square values
SUM GEOM	100	8	12.892
	150	9	14.324
NEG BINORM.	100	8	9.595
	150	9	9.223
TOTAL TRIAL	100	8	43.687
	150	9	50.806
EXPONENTIAL	100	8	12.091
	150	9	13.306

<u>Degrees of freedom</u>	<u>α</u>	<u>Critical χ^2</u>
8	0.05	15.5073
8	0.10	13.3616
9	0.05	16.9190
9	0.10	14.6837

N is the number of random variates generated for testing. Since any one method in the above table can be changed to generate a Pascal variate with non-integer number of successes, therefore to do the chi-square test for the non-integer method would be a repetition.

5. BINOMIAL DISTRIBUTION

Let A be an event. Define the random variable X as the number of successes when the trials are repeated N times. Then X has a binomial distribution.

$$f(x) = P(X=x) = \binom{N}{x} p^x q^{N-x} \quad x=0,1,2,\dots,N$$

where p is the probability of the occurrence of A, and $q=1-p$, the probability of the failure of A.

$$EX = Np$$

$$VX = Npq$$

A. DIRECT METHOD

This method is straightforward by using the definition of the distribution. The steps are as below:

- 1) Set X equal to 0.0.
- 2) Generate R, a uniform random variate in (0,1).
- 3) If R is less or equal to p, that means the event occurs; increase X by 1.
- 4) If R is greater than p, the event does not occur, go to step 2.
- 5) After N uniform random variates are generated, X is the desired variate.

PROGRAM LISTING

```

C*****
C
C      TO RETURN A BINOMIAL RANDOM VARIATE BY DIRECT METHOD
C
C      NT = THE NUMBER OF TRIALS
C
C      P = THE PROBABILITY THAT THE EVENT WILL OCCUR
C
C      VAL = THE RETURNED BINOMIAL VARIATE
C*****
C
C      SUBROUTINE BINOM(NT,P,VAL)
C      DATA NSEED/567801/
C      X=0.0
C      DO 7 I=1,NT
C      R=RAN(NSEED)
C      IF(R-P) 6,6,7
C 6 X=X+1.0
C 7 CONTINUE
C      RETURN
C      END

```

B. NORMAL APPROXIMATION

For large number of trials, N the binomial variate x can be approximated by the normal distribution with mean equal to Np and variance equal to Npq .

Verification:

$$\text{Let } h = \frac{x - Np}{\sqrt{Npq}}$$

We need to prove that the probability mass function $P(h)$

$$\approx \frac{1}{\sqrt{Npq}} \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}h^2}$$

which is equivalent to showing

$$\lim_{N \rightarrow \infty} \frac{\sqrt{2\pi Npq} P(h\sqrt{Npq} + Np)}{e^{-\frac{1}{2}h^2}} = 1$$

First, consider

$$\begin{aligned} P(h\sqrt{Npq} + Np) &= P(x) \\ &= \binom{N}{x} p^x q^{N-x} \end{aligned}$$

By Stirling's formula, $m! = \sqrt{2\pi m} m^m e^{-m} e^{r(m)}$ while $0 < r(m) < \frac{1}{12m}$

$$\begin{aligned} \binom{N}{x} p^x q^{N-x} &= \frac{N!}{x! (N-x)!} p^x q^{N-x} \\ &= \frac{\sqrt{2\pi N} N^N e^{-N} e^{r(N)} p^x q^{N-x}}{\sqrt{2\pi x} x^x e^{-x} e^{r(x)} \sqrt{2\pi (N-x)} (N-x)^{N-x} e^{-(N-x)} e^{r(N-x)}} \\ &= \frac{1}{\sqrt{2\pi}} \sqrt{\frac{N}{x(N-x)}} \left(\frac{Np}{x} \right)^x \left(\frac{Nq}{N-x} \right)^{N-x} e^{-r(x) - r(N-x) + r(N)} \end{aligned}$$

Let $R = -r(x) - r(N-x) + r(N)$

$$\text{then } 0 < R < -\frac{1}{12x} - \frac{1}{12(N-x)} + \frac{1}{12N}$$

$$\text{implies } -\left[\frac{1}{12x} + \frac{1}{12(N-x)} + \frac{1}{12N}\right] < R < \frac{1}{12x} + \frac{1}{12(N-x)} + \frac{1}{12N}$$

$$\text{implies } |R| < \frac{1}{12} \left[\frac{1}{x} + \frac{1}{N-x} + \frac{1}{N} \right]$$

Second, consider

$$-\text{Log}_e \left[\left(\frac{Np}{x} \right)^x \left(\frac{Nq}{N-x} \right)^{N-x} \right]$$

Using the expansion $\text{Log}_e(1+x) = x - \frac{1}{2}x^2 + \theta x^3$ for some θ such that $|\theta| < 1$, valid for $|x| < \frac{1}{4}$, we obtain that

$$\begin{aligned} & -\text{Log}_e \left[\left(\frac{Np}{x} \right)^x \left(\frac{Nq}{N-x} \right)^{N-x} \right] \\ &= (Np + h\sqrt{Npq}) \text{Log}_e(1 + h\sqrt{q/Np}) \\ & \quad + (Nq + h\sqrt{Npq}) \text{Log}_e(1 - h\sqrt{p/Nq}) \\ &= (Np + h\sqrt{Npq}) \left[h\sqrt{(q/Np)} - \frac{h^2}{2} \frac{q}{Np} + \text{terms in } \frac{1}{N^{3/2}} \right] \\ & \quad + (Nq - h\sqrt{Npq}) \left[-h\sqrt{(p/Nq)} - \frac{h^2}{2} \frac{p}{Nq} + \text{terms in } \frac{1}{N^{3/2}} \right] \\ &= \left[h\sqrt{Npq} - \frac{h^2}{2} q + qh^2 + \text{terms in } \frac{1}{N^{1/2}} \right] \\ & \quad + \left[-h\sqrt{Npq} - \frac{h^2}{2} p + ph^2 + \text{terms in } \frac{1}{N^{1/2}} \right] \\ &= \frac{1}{2} h^2 + \text{terms in } \frac{1}{N^{1/2}} \end{aligned}$$

$$\text{Therefore } \left(\frac{Np}{x}\right)^x \left(\frac{Nq}{N-x}\right)^{N-x} = \text{Exp} \left[- \left(\frac{1}{2} h^2 + \text{terms in } \frac{1}{N^{1/2}} \right) \right]$$

$$\text{Third, consider } \frac{x(N-x)}{N} ;$$

$$\begin{aligned} \text{Since } x = Np + h\sqrt{Npq} , \quad \frac{x(N-x)}{N} &= N(p + h\sqrt{pq/N})(q - h\sqrt{pq/N}) \\ &\approx Npq \end{aligned}$$

At last substitute R , Npq and $\text{Exp} \left[- \left(\frac{1}{2} h^2 + \text{terms in } \frac{1}{N^{1/2}} \right) \right]$ into

$\binom{N}{x} p^x q^{N-x}$, we have

$$P(h\sqrt{Npq} + Np) = \frac{1}{\sqrt{2\pi Npq}} \text{Exp} \left[- \left(\frac{1}{2} h^2 + \text{terms in } \frac{1}{N^{1/2}} \right) \right] e^R$$

$$\text{which implies } \lim_{N \rightarrow \infty} \frac{\sqrt{2\pi Npq} P(h\sqrt{Npq} + Np)}{e^{-\frac{1}{2}h^2}} = 1$$

Thus, we have shown that x is normally distributed with mean equal to Np and standard deviation equal to \sqrt{Npq} .

PROGRAM LISTING

```

C*****
C
C      TO RETURN A BINOMIAL RANDOM VARIATE BY NORMAL
C      APPROXIMATION
C
C      NT = THE NUMBER OF TRIALS
C
C      P = THE PROBABILITY THAT THE EVENT WILL OCCUR
C
C      VAL = THE RETURNED BINOMIAL VARIATE
C*****
C
      SUBROUTINE BINOM(NT,P,VAL)
      DATA NSEED/567801/
      EX=P*NT
      VX=P*NT*(1.-P)
      IF(NT.LT.15) GO TO 4
      VAL=-.0
      DO 1 I=1,12
      VAL=VAL+RAN(NSEED)
1  CONTINUE
      VAL=(VAL-6.0)*SQSQRT(VX)+EX
      GO TO 5
4  WRITE(6,3)
3  FORMAT(' THE NUMBER OF TRAILS IS NOT LARGE ENOUGH TO
+USE NORMAL APPROXIMATION ')
5  RETURN
      END

```

PROGRAM TESTING RESULTS

METHODS	PROGRAM RESULTS				GIVEN	
	Mean		Variance		Mean	Variance
	N=100	N=1000	N=100	N=1000		
DIRECT	8.910	8.783	0.822	1.018	8.8	1.056
NORM APPROX.	4.955	4.482	4.938	4.216	5.0	4.500

The input parameters for the example of DIRECT method are:

$$NT = 10$$

$$P = 0.88$$

The input parameters for the example of NORM APPROXIMATION method are:

$$NT = 50$$

$$P = 0.1$$

METHODS	EXECUTION TIME IN MILLI-SECONDS	
	N=100	N=1000
DIRECT	79	852
NORM APPROX	87	1004

N is the number of random variates generated for testing.

SUMMARY OF T-TEST

METHODS	Degrees of freedom	T-values
DIRECT	99	1.2430
	999	0.5329
NORM APPROX.	99	0.2027
	999	7.9911

<u>Degrees of freedom</u>	<u>α</u>	<u>Critical χ^2</u>
60	0.05	1.671
60	0.10	1.296
120	0.05	1.658
120	0.10	1.289
∞	0.05	1.645
∞	0.10	1.282

6. POISSON DISTRIBUTION

Let X be a discrete random variable assuming the possible values: $0, 1, 2, \dots, n, \dots$. If

$$P(X = k) = \frac{e^{-\lambda t} \lambda^k t^k}{k!}, \quad k = 0, 1, \dots, n, \dots$$

Then X has a Poisson distribution with parameter $\lambda > 0$.

λ is the arrival rate per unit of time and t is the unit of time. The mean and the variance of X are:

$$EX = \lambda t$$

$$VX = \lambda t$$

A. NORMAL METHOD

If a series of n independent Bernoulli trials were taken, in each of which there is a small probability p of an event occurring, then as n approaches infinity and p approaches zero in such a manner that $\lambda=np$ remains fixed, the probability of x occurrences is given by the Poisson distribution of $t=1$.

Proof:

$$\begin{aligned} P(x=k) &= \binom{n}{k} p^k (1-p)^{n-k} \\ &= \frac{n!}{k! (n-k)!} p^k (1-p)^{n-k} \\ &= \frac{n(n-1)(n-2)\dots(n-k+1)}{k!} p^k (1-p)^{n-k} \end{aligned}$$

Substitute in $p=\lambda/n$;

$$\begin{aligned} P(x=k) &= \frac{n(n-1)\dots(n-k+1)}{k!} \left(\frac{\lambda}{n}\right)^k \left(\frac{n-\lambda}{n}\right)^{n-k} \\ &= \frac{\lambda^k}{k!} \left[(1) \left(1 - \frac{1}{n}\right) \left(1 - \frac{2}{n}\right) \dots \left(1 - \frac{k-1}{n}\right) \right] \left(1 - \frac{\lambda}{n}\right)^{n-k} \\ &= \frac{\lambda^k}{k!} \left[(1) \left(1 - \frac{1}{n}\right) \left(1 - \frac{2}{n}\right) \dots \left(1 - \frac{k-1}{n}\right) \right] \left(1 - \frac{\lambda}{n}\right)^n \left(1 - \frac{\lambda}{n}\right)^{-k} \end{aligned}$$

As n approaches to infinity,

$(1-1/n)(1-2/n)\dots$ approaches to 1, $(1-\lambda/n)^{-k}$ approaches to 1 and $(1-\lambda/n)^n$ approaches to $e^{-\lambda}$.

Therefore

$$P(x=k) = e^{-\lambda} \frac{\lambda^x}{x!} .$$

It is well known that Stirling's formula approximates $n! \sim \sqrt{2\pi} e^{-n} n^{n+1/2}$. And that for large n , using Stirling's formula for the various $n!$, Bernoulli trials can be approximated by the normal distribution (27).

$$\begin{aligned}
 P(x=k) &= \binom{n}{k} p^k (1-p)^{n-k} \\
 &= \frac{1}{\sqrt{2\pi np(1-p)}} \exp\left(-\frac{1}{2} \left[\frac{k-np}{np(1-p)} \right]^2\right) \\
 P(x < k) &= P\left[\frac{x-np}{np(1-p)} \leq \frac{k-np}{np(1-p)} \right] \\
 &\approx \frac{1}{2\pi} \int_{-\infty}^{(k-np)/\sqrt{np(1-p)}} e^{-t^2/2} dt
 \end{aligned}$$

As a result, a Poisson variate can be approximated by a normal distribution when n is large enough and with the mean and variance of the normal distribution both equal to λ .

PROGRAM LISTING

```

C*****
C
C      TO RETURN A RANDOM POISSON VARIATE WITH TIME UNIT
C      EQUAL TO ONE BY NORMAL METHOD
C
C      A =ARRIVAL RATE PER UNIT OF TIME
C
C      VAL = THE RETURNED POISSON VARIATE
C
C*****
C
C      SUBROUTINE POIS(A,VAL)
C      CALL NORM(A,A,X)
C      VAL=X
C      RETURN
C      END

```

B. EXPONENTIAL ARRIVAL TIME METHOD

It can be shown that if

- 1) The time intervals between successive arrivals are independently and identically distributed; further, the probability of an arrival occurring in the time interval between T and $T+h$ depends only on the length h of the interval and not on T . The corresponding interarrival density function is designated as $f(t)$.
- 2) In any interval of time $h > 0$, there is a positive probability of an arrival.
- 3) In any sufficiently small interval of time, at most only one arrival can occur.

Then the number of arrivals is Poisson distributed.

Suppose for simplicity that the system starts at time 0 and the first arrival occurs at time t , where $t > 0$. Therefore $f(t)$ represents the density function for both the length of the arrival intervals as well as the actual time for the first arrival.

Define

$$r(T) = 1 - \int_0^T f(t) dt$$

so that

$$r(T) = P(\text{first arrival occurs after time } T)$$

Then from 1) and 2) above,

$$r(T+h) = r(T)r(h) \quad \text{for all } T, h > 0$$

The only function that satisfies $r(T+h) = r(T)r(h)$ is

$$r(T) = e^{-\lambda T}$$

where λ is a positive constant. Therefore

$$e^{-\lambda T} = 1 - \int_0^T f(t) dt ,$$

So that

$$f(T) = \lambda e^{-\lambda T} .$$

Define

$$P_n(T) = P(n \text{ arrivals occur in the interval}(0,T))$$

Let $t = x$ be the time of the first arrival event, and according to (3), only one arrival enters at x . By 1), we can write

$$\begin{aligned} P_n(T) &= \int_0^T P_{n-1}(T-x) f(x) dx \\ &= \int_0^T P_{n-1}(y) f(T-y) dy \quad \text{for } n=1,2,\dots \end{aligned}$$

where $y = T-x$. Using $f(t) = \lambda e^{-\lambda t}$,

$$P_n(T) = \int_0^T P_{n-1}(y) \lambda e^{-\lambda(T-y)}$$

Differentiating with respect to T yields

$$\frac{dP_n}{dT} = -\lambda P_n(T) + \lambda P_{n-1}(T) \quad \text{for } n=1,2,\dots$$

Claim that $P_n(T) = \frac{(\lambda T)^n e^{-\lambda T}}{n!}$

$$\begin{aligned} \frac{dP_n(T)}{dT} &= \frac{e^{-\lambda T}}{n!} n(\lambda T)^{n-1} + \frac{(\lambda T)^n}{n!} (-\lambda) e^{-\lambda T} \\ &= \frac{(\lambda T)^{n-1} e^{-\lambda T}}{(n-1)!} \lambda - \lambda e^{-\lambda T} \frac{(\lambda T)^n}{n!} \\ &= -\lambda P_n(T) + \lambda P_{n-1}(T) \end{aligned}$$

It has now been proved that:

- 1) The density function of the arrival time intervals between occurrences of consecutive events is exponential, with expected value equal to $1/\lambda$.
- 2) The probability of n events occurring during time t is Poisson. In one unit time interval, the expected value of n is λ .

If t_1, t_2, \dots are exponential time intervals with expected value equal to 1, such that

$$\sum_{i=0}^x t_i \leq \lambda < \sum_{i=0}^{x+1} t_i$$

then x is the required variate.

Since exponential $t_i = -\text{Log}_e r_i$, substituting it in the above inequalities, we have

$$\sum_{i=0}^x -\text{Log}_e r_i \leq \lambda < \sum_{i=0}^{x+1} -\text{Log}_e r_i$$

$$\prod_{i=0}^x r_i \geq e^{-\lambda} > \prod_{i=0}^{x+1} r_i$$

PROGRAM LISTING

```

C*****
C
C      TO RETURN A POISSON RANDOM VARIATE WITH TIME UNIT
C      EQUAL TO ONE BY EXPONENTIAL ARRIVAL TIME METHOD
C
C      A = ARRIVAL RATE PER UNIT OF TIME
C
C      VAL = THE RETURNED POISSON VARIATE
C*****
C
      SUBROUTINE POIS(A,VAL)
      DATA NSEED/567801/
      VAL=0.0
      TEST=EXP(-A)
      TR=1.0
      5 TR=TR*RAN(NSEED)
      IF(TR-TEST) 10,8,8
      8 VAL=VAL+1
      GO TO 5
      10 RETURN
      END

```

PROGRAM TESTING RESULTS

METHODS	PROGRAM RESULTS				GIVEN	
	Mean		Variance		Mean	Variance
	N=100	N=1000	N=100	N=1000		
NORMAL	10.629	11.059	12.750	11.108	11.0	11.0
EXPONENT.	10.390	11.068	11.418	12.055	11.0	11.0

The input parameter of this example is A=11.0.

METHODS	EXECUTION TIME IN MILLI-SECONDS	
	N=100	N=1000
NORMAL	98	840
EXPONENTIAL	22	188

N is the number of random variates generated for testing.

SUMMARY OF CHI-SQUARE TEST

METHODS	N	Degrees of freedom	Chi-square values
NORMAL	100	10	14.728
	1000	17	67.592
EXPONENT.	100	9	32.657
	1000	17	5.425

<u>Degrees of freedom</u>	<u>α</u>	<u>Critical χ^2</u>
9	0.05	16.9190
9	0.10	14.6837
10	0.05	18.3070
10	0.10	15.9871
17	0.05	27.5871
17	0.10	24.7690

N is the number of random variates generated for testing.

7. EXPONENTIAL DISTRIBUTION

A random variable x is said to have an exponential distribution if its density function is

$$f(x) = \alpha e^{-\alpha x} \quad \text{for } \alpha > 0, x > 0$$

The cumulative distribution function of x is

$$F(x) = 1 - e^{-\alpha x}$$

The mean and the variance of x are:

$$EX = 1/\alpha$$

$$VX = 1/\alpha^2$$

A. INVERSION METHOD

The exponential cumulative distribution function is

$$F(x) = 1 - e^{-\alpha x}$$

$$\text{implies} \quad -\alpha x = \text{Log}_e(1 - F(x))$$

$$x = -(\text{Log}_e(1 - F(x)))/\alpha$$

$F(x)$ is a uniform random number in $(0,1)$ implies

that $1 - F(x)$ is also a uniform random number in $(0,1)$.

To obtain an exponential random variate:

- 1) Generate U , a uniform distributed random variate in $(0,1)$.
- 2) Calculate $x = -\text{Log}_e(U)/\alpha$, then x is the desired variate.

PROGRAM LISTING

```

C*****
C
C    TO RETURN AN EXPONENTIAL RANDOM VARIATE BY INVERSION *
C    METHOD *
C
C    EX = THE MEAN OF THE DISTRIBUTION *
C
C    VAL = THE RETURNED EXPONENTIAL VARIATE *
C
C*****
C
C    SUBROUTINE EXPER(EX,VAL)
C    DATA NSEED/567801/
C    VAL=-EX*ALOG(RAN(NSEED))
C    RETURN
C    END

```

B. VON NEUMANN METHOD 1

To generate an exponential random variate with mean equal to 1, select uniform random numbers $x_0, r_1, r_2, \dots, r_i, x_1 \dots$ and form the sequence of sums:

$$\begin{array}{ccccccc}
 1 - x_0 + r_1 & + & r_2 & + & \dots & + & r_i \\
 1 - x_1 + r_{i+2} & + & r_{i+2} & + & \dots & + & r_{i+j} \\
 \vdots & & \vdots & & & & \\
 \vdots & & \vdots & & & & \\
 1 - x_t + r_{p+1} & + & r_{p+2} & + & \dots & + & r_{p+q}
 \end{array}$$

The individual sums are terminated when they become ≥ 1 for the first time; the sequence is ended on the first trial involving an odd value of q . The quantity

$$y = t - 1 + x_t$$

is an exponential variate with mean equal to 1.

Verification:

Consider the last sum

$$S_q = 1 - x_t + r_{p+1} + r_{p+2} + \dots + r_{p+q}$$

Let u_q represents the event $(S_q \geq 1 | x_t = x)$

$$\text{implies } u_q = P \left[\sum_{i=1}^q r_{p+i} \geq x \right] dr$$

$$= 1 - \int_0^x \frac{r^{q-1}}{(q-1)!} dr$$

$$= 1 - \frac{x^q}{q!}$$

Let V_q represents the event $(S_q \geq \text{for the first time } x_t = x)$. Since the numbers r_n are all positive, the events corresponding to V_q and u_{q-1} are mutually exclusive, which means that the probability of either V_q or u_{q-1} occurring is the probability of V_q occurring plus the probability of u_{q-1} occurring.

The acceptance probability of the value x is

$$\begin{aligned} P(q \text{ odd} | x_t = x) &= \sum_{\text{odd } j} V_j \\ &= \sum_{\text{odd } j} \left(\frac{x^j}{j!} - \frac{x^{j+1}}{(j+1)!} \right) \\ &= e^{-x} \quad (0 < x < 1) \end{aligned}$$

The distribution of the integer $(t-1)$ is proportional to:

$$\begin{aligned} & \left[\int_0^1 P(q \text{ even} | x_t = x) dx \right]^{t-1} \\ &= \left[\int_0^1 \sum_{\text{even } j} \left(\frac{x^j}{j!} - \frac{x^{j+1}}{(j+1)!} \right) dx \right]^{t-1} \\ &= \left[\int_0^1 (1 - e^{-x}) dx \right]^{t-1} \\ &= e^{-(t-1)} \end{aligned}$$

The distribution of the event to accept x_t for an odd q and to accept $t-1$ for $t-1$ even q 's is $e^{-x} e^{-(t-1)} = e^{-(t-1+x)}$. Therefore, $y = t-1+x$ is exponentially distributed.

PROGRAM LISTING

```

C*****
C
C    TO RETURN AN EXPONENTIAL RANDOM VARIATE BY VON
C    NEUMANN METHOD 1
C
C    EX = THE MEAN OF THE DISTRIBUTION
C
C    VAL = THE RETURNED EXPONENTIAL VARIATE
C*****
C
C    SUBROUTINE EXPR(EX,VAL)
C    DATA NSEED/567801/
C    I=0
C 6  I=I+1
C    R=RAN(NSEED)
C    SUM=1.-R.
C    J=0
C 1  J=J+1
C    SUM=SUM+RAN(NSEED)
C    IF(SUM.GE.1.0) GO TO 3
C    GO TO 1
C 3  K=(-1)**J
C    IF(K.LT.0) GO TO 5
C    GO TO 6
C 5  VAL=(I-1)+R+EX-1
C    RETURN
C    END

```

C. VON NEUMANN METHOD 2

This method is a slight change of method 1. It makes use of conditional probabilities and independence of events. To generate an exponential random variate with mean equal to one, the steps are as follows:

- 1) Choose two uniform distributed random numbers R_0 and R_1 in $(0,1)$.
- 2) If $R_1 \leq R_0$, score 1 and choose a new random number R_2 .
- 3) If $R_2 \leq R_1$, score 1 and choose a new random number R_3 .
- 4) Continue doing this until the test $R_{i+1} \leq R_i$ fails to occur on an even i , then $(i-1) \cdot R_0$ is the required exponential random number.

Verification:

If one comparison test is made between R_0 and R_1 ,

$$P[R_0 \geq R_1] = R_0$$

By induction, if i comparison tests were made,

$$P[R_0 \geq R_1 \geq \dots \geq R_i] = R_0^i / i!$$

And if $i+1$ tests were made,

$$P[R_0 \geq R_1 \geq \dots \geq R_{i+1}] = R_0^{i+1} / (i+1)!$$

Therefore

$$P[R_0 \geq R_1 \geq R_2 \geq \dots \geq R_i < R_{i+1}] = \frac{R_0^i}{i!} - \frac{R_0^{i+1}}{(i+1)!}$$

For a given R_0 , the probability of i being even is

$$\begin{aligned}
 & 1 - \frac{R_0}{1!} + \frac{R_0^2}{2!} - \frac{R_0^3}{3!} + \frac{R_0^4}{4!} - \frac{R_0^5}{5!} + \dots \\
 &= \sum_{r=0}^{\infty} (-1)^r \frac{R_0^r}{r!} \\
 &= e^{-R_0}
 \end{aligned}$$

The conditional probability of R_0 for i even is

$$\begin{aligned}
 P[i \text{ even} | R] &= \frac{P[i \text{ even} \cap R_0]}{P[R_0]} \\
 &= \frac{e^{-R_0}}{\int_0^1 e^{-R} dR} \\
 &= \frac{e^{-R_0}}{-e^{-R} \Big|_0^1} \\
 &= e^{-R_0} / (-e^{-1} + 1) \\
 &= (1 - e^{-1})^{-1} e^{-R_0}
 \end{aligned}$$

Thus if the sequence of tests is repeated until failure occurs on an even i , the initial random number has been selected from the truncated distribution $(1 - e^{-1})^{-1} e^{-x}$ while $0 \leq x \leq 1$.

The probability of $I=i$ is

$$\frac{1}{(i+1)!} - \frac{1}{(i+2)!}$$

The probability of i being odd for one single trial is

$$\begin{aligned} & \frac{1}{2!} - \frac{1}{3!} + \frac{1}{4!} - \frac{1}{5!} + \dots \\ &= 1 - 1 + \frac{1}{2!} - \frac{1}{3!} + \frac{1}{4!} - \frac{1}{5!} + \dots \\ &= 1 - \frac{1}{1!} + \frac{1}{2!} - \frac{1}{3!} + \frac{1}{4!} - \frac{1}{5!} + \dots \\ &= e^{-1} \end{aligned}$$

The probability of i being even for one single trial is

$$1 - e^{-1}$$

Therefore the number of sequences n trials before success has a distribution

$$(1 - e^{-1})e^{-n}$$

The variables n and R are independent. By multiplying $(1 - e^{-1})^{-1}e^{-R_0}$ and $(1 - e^{-1})e^{-n}$ together

$$(1 - e^{-1})^{-1}e^{-R_0} (1 - e^{-1})e^{-n} = e^{-(R_0+n)}$$

showing that $R_0 + n$ is exponentially distributed and is the desired variate for the Von Neumann method.

PROGRAM LISTING

```

C*****
C
C      TO RETURN AN EXPONENTIAL RANDOM VARIATE WITH VON
C      NEUMANN METHOD 2
C
C      EX = THE MEAN OF THE DISTRIBUTION
C
C      VAL = THE RETURNED EXPONENTIAL VARIATE
C*****
C
      SUBROUTINE EXPR(EX,VAL)
      DATA NSEED/567801/
      NT=0
1     NT=NT+1
      R1=RAN(NSEED)
      RS=R1
      NC=0
2     R2=RAN(NSEED)
      IF(RS.LT.R2) GO TO 3
      NC=NC+1
      RS=R2
3     K=(-1)**NC
      IF(K.EQ.1) GO TO 4
      GO TO 1
4     VAL=(NT-1+R1)*EX
      RETURN
      END

```

D. MARSAGLIA METHOD

Marsaglia chose the minimum of a random number of n uniform random variates, then added a random integer m to obtain an exponential variate. n and m are integers taking values according to the table below.

k	n	$1/((e-1)k!)$	Prob of n	Cum.Prob of n
1	1	0.46694	0.5825	0.5825
2	2	0.23347	0.2913	0.8738
3	3	0.07782	0.0971	0.9709
4	4	0.00065	0.0243	0.9952
5	5	0.00009	0.0048	1.0000

k	m	$(e-1)/e$	Prob of n	Cum. Prob of n
1	0	0.78785	0.6321	0.6321
2	1	0.28983	0.2325	0.8646
3	2	0.10662	0.0855	0.9501
4	3	0.03922	0.0315	0.9816
5	4	0.01443	0.0116	0.9932
6	5	0.00531	0.0043	0.9975
7	6	0.00195	0.0016	0.9991
8	7	0.00072	0.0006	1.0049

The method is based on the theorem below:

Theorem:

If $c=1/(e-1)$ and if the random variable n takes values $1, 2, 3, \dots$ with probabilities $c, c/2!, c/3!, \dots$ and if, independently, the random variable m takes values $0, 1, 2, \dots$ with probabilities $1/(ce), 1/(ce^2), 1/(ce^3), \dots$, then the random variable

$$x = m + \min(u_1, u_2, \dots, u_n)$$

has the exponential distribution

$$P(x \leq a) = 1 - e^{-a} \quad 0 \leq a$$

Proof:

$$\text{Let } y = \min(u_1, u_2, \dots, u_n)$$

$$\begin{aligned} \text{For } 0 \leq \theta \leq 1, \quad P(y \leq \theta) &= 1 - c(1-\theta) - c/2!(1-\theta)^2 - \dots \\ &= ce(1 - e^{-\theta}) \end{aligned}$$

Let $a=k+\theta$, k is a non-negative integer.

$$\begin{aligned} P(x \leq a) &= P(x \leq k + \theta) \\ &= P(x \leq k-1) + P(m=k, y \leq \theta) \\ &= 1 - P(x > k) + P(m=k, y \leq \theta) \end{aligned}$$

$$\begin{aligned} P(x > k) &= \frac{1}{ce^{k+1}} + \frac{1}{ce^{k+2}} + \dots \\ &= \frac{1}{ce^{k+1}} \left[1 + \frac{1}{e} + \frac{1}{e^2} + \dots \right] \end{aligned}$$

$$= \frac{1}{ce^{k+1}} \left(1 - \frac{1}{e} \right)^{-1}$$

$$= \frac{e-1}{e^{k+1}} \cdot \frac{e}{e-1}$$

$$= e^{-k}$$

$$P(m=k, y \leq \theta) = \frac{1}{ce^{k+1}} ce(1-e^{-\theta})$$

$$= \frac{1-e^{-\theta}}{e^k}$$

Therefore $P(x \leq a) = 1 - e^{-k} + e^{-k}(1 - e^{-\theta})$

$$= 1 - e^{-k}(1 - 1 + e^{-\theta})$$

$$= 1 - e^{-(k+\theta)}$$

$$= 1 - e^{-a}$$

PROGRAM LISTING

```

C*****
C
C      TO RETURN AN EXPONENTIAL RANDOM VARIATE BY
C      MARSAGLIA METHOD
C
C      EX = THE MEAN OF THE DISTRIBUTION
C
C      VAL = THE RETURNED EXPONENTIAL VARIATE
C*****
C
      SUBROUTINE EXPER(EX,VAL)
      DATA NSEED/567801/
      R1=RAN(NSEED)
      R2=RAN(NSEED)
      N=5
      IF(R2.LE.0.99) N=4
      IF(R2.LE.0.97) N=3
      IF(R2.LE.0.87) N=2
      IF(R2.LE.0.58) N=1
      IF(R1.LE.1.000) M=7
      IF(R1.LE.0.9991) M=6
      IF(R1.LE.0.9975) M=5
      IF(R1.LE.0.9932) M=4
      IF(R1.LE.0.9816) M=3
      IF(R1.LE.0.9501) M=2
      IF(R1.LE.0.8646) M=1
      IF(R1.LE.0.6321) M=0
      AMIN=1.
      DO 10 J=1,N
      R=RAN(NSEED)
      IF(R.LE.AMIN) AMIN=R
10  CONTINUE
      VAL=(M+AMIN)*EX
      RETURN
      END

```

PROGRAM TESTING RESULTS

METHODS	N	PROGRAM RESULTS		GIVEN	
		Mean		Mean	
		Example 1	Example 2	Example 1	Example 2
INVERSION	100	3.082	5.662	3.50	6.43
	1000	3.502	6.433	3.50	6.43
VON NEUMANN 1	100	3.350	6.157	3.50	6.43
	1000	3.476	6.386	3.50	6.43
VON NEUMANN 2	100	3.327	6.111	3.50	6.43
	1000	3.312	6.084	3.50	6.43
MARSAGLIA	100	3.464	6.360	3.50	6.43
	1000	3.408	6.260	3.50	6.43

METHODS	N	EXECUTION TIME IN MILLI-SECONDS	
		Example 1	Example 2
INVERSION	100	12	13
	1000	122	131
VON NEUMANN 1	100	37	33
	1000	294	339
VON NEUMANN 2	100	43	38
	1000	421	377
MARSAGLIA	1000	41	42
	1000	434	344

N is the number of random variates generated for testing.

SUMMARY OF CHI-SQUARE TEST

METHODS	N	degrees of freedom	chi-square values
INVERSON	1000	9	2.728
	2500	18	12.378
	5000	20	12.171
VON NEUMANN 1	1000	9	7.317
	2500	18	19.711
	5000	20	26.660
VON NEUMANN 2	1000	9	16.619
	2500	18	12.601
	5000	20	14.409
MARSAGLIA	1000	9	9.272
	2500	18	5.550
	5000	20	14.160

<u>Degrees of freedom</u>	<u>α</u>	<u>Critical χ^2</u>
9	0.05	16.9190
9	0.10	14.6837
18	0.05	28.8693
18	0.10	25.9894
20	0.05	31.4104
20	0.10	28.4120

N is the number of random variates generated for testing.

8. NORMAL DISTRIBUTION

The random variable x , assuming all real values $-\infty < x < \infty$, has a normal distribution if its probability density function is of the form

$$f(x) = \frac{1}{\sqrt{2\pi} \sigma} \exp\left(-\frac{1}{2} \left[\frac{x-\mu}{\sigma}\right]^2\right) \quad -\infty < x < \infty$$

The cumulative function is

$$F(x) = \frac{1}{2} \int_{-\infty}^x \exp\left(-\frac{1}{2} \left[\frac{t-\mu}{\sigma}\right]^2\right) dt$$

μ and σ must satisfy the conditions $-\infty < \mu < \infty$ and $\sigma > 0$.

μ is the mean and σ is the standard deviation.

A. CENTRAL LIMIT METHOD

The explanation of this method here is straightly taken from C. E. Donaghey's notes of his simulation course, IE 670 at the University of Houston.

Central Limit Theorem: If x_1, x_2, \dots, x_n are n independent random variables having the same distribution with mean μ and variance σ^2 ; then as $n \rightarrow \infty$, the random variable

$$z = \frac{\bar{x} - \mu}{\sigma/\sqrt{n}} \quad \text{for } \bar{x} = \frac{1}{n} \sum_{i=1}^n x_i$$

has a normal distribution with mean=0, and variance=1.

For a uniform distribution in the interval (0,1)

$$f(x) = \begin{cases} 1 & 0 < x < 1 \\ 0 & \text{elsewhere} \end{cases}$$

The mean of this distribution, μ will be

$$\mu = \int_0^1 x dx = \frac{x^2}{2} \Big|_0^1 = \frac{1}{2}$$

The variance of this distribution, σ^2 will be

$$\sigma^2 = EX(x^2) - \mu^2 = \int_0^1 x^2 dx - \left(\frac{1}{2}\right)^2 = 1/12$$

Let

$$z = \left(\frac{1}{n} \sum_{i=1}^n x_i - \frac{1}{2} \right) \sqrt{12n}$$

$$z = \left(\sum_{i=1}^n x_i - \frac{n}{2} \right) \sqrt{\frac{12}{n}}$$

If $n=12$

$$z = \sum_{i=1}^{12} x_i - 6$$

where x_i is a random variable from a uniform distribution in (0,1).

to obtain a random normal variate with mean equal to EX and variance equal to VX:

- 1) Generate and sum up 12 uniform variates in (0,1), say the sum is y.
- 2) $x = (y - 6.) * \text{SQRT}(VX) + EX$ is the required variate.

PROGRAM LISTING

```

C*****
C
C   TO RETURN A RANDOM NORMAL VARIATE BY CENTRAL LIMIT
C   METHOD
C
C   EX = THE MEAN OF THE DISTRIBUTION
C
C   VX = THE VARIANCE OF THE DISTRIBUTION
C
C   VAL = THE RETURNED VARIATE
C*****
C
C   SUBROUTINE NORM(EX,VX,VAL)
C   DATA NSEED/567801/
C   X=0.0
C   DO 1 I=1,12
C     X=X+RAN(NSEED)
C 1 CONTINUE
C   VAL=SQRT(VX)*(X-6.)+EX
C   RETURN
C   END

```

B. DIRECT METHOD OF COSINE AND SINE

Let U_1 and U_2 be two independent random variables from the uniform distribution function in $(0,1)$, then

$$x_1 = (-2\log_e U_1)^{1/2} \cos(2\pi U_2)$$

$$x_2 = (-2\log_e U_1)^{1/2} \sin(2\pi U_2)$$

are two independent standard random variates. In order to prove this, let us express U_1 and U_2 in terms of x_1 and x_2 .

To express U_1 and U_2 in terms of x_1 and x_2 :

$$x_1^2 = (-2\log_e U_1) \cos^2(2\pi U_2)$$

$$x_2^2 = (-2\log_e U_1) \sin^2(2\pi U_2)$$

$$x_1^2 + x_2^2 = (-2\log_e U_1) (\cos^2(2\pi U_2) + \sin^2(2\pi U_2))$$

$$x_1^2 + x_2^2 = -2\log_e U_1$$

$$-(x_1^2 + x_2^2)/2 = \log_e U_1$$

$$U_1 = e^{-1/2(x_1^2 + x_2^2)}$$

$$\sin(2\pi U_2) / \cos(2\pi U_2) = x_2 / x_1$$

$$\tan(2\pi U_2) = x_2 / x_1$$

$$2\pi U_2 = \arctan(x_2 / x_1)$$

$$U_2 = \frac{1}{2\pi} \arctan(x_2 / x_1)$$

To show that $1/(2\pi)e^{-(x_1^2+x_2^2)/2}$ is a joint density function of x_1, x_2 :

Proof:

- 1) Obviously $f(x_1, x_2) = e^{-(x_1^2+x_2^2)/2} \geq 0$, for all (x_1, x_2)
- 2) Let $(x_1, x_2) \in R$, R is the region bounded by the circle having radius M .

Changing into polar coordinates:

$$\begin{aligned}
 & \int_R \int 1/(2\pi) e^{-(x_1^2+x_2^2)/2} dx_1 dx_2 \\
 &= 1/(2\pi) \int_{\psi=0}^{2\pi} \int_{\rho=0}^M e^{-\rho^2/2} \rho d\rho d\psi \\
 &= 1/(2\pi) \int_{\psi=0}^{2\pi} e^{-\rho^2/2} \Big|_{\rho=0}^M d\rho d\psi \\
 &= 1/(2\pi) \int_{\psi=0}^2 (-e^{-M^2/2} + 1) d\psi \\
 &= (-e^{-M^2/2} + 1) \\
 &= 1
 \end{aligned}$$

Therefore $f(x_1, x_2)$ is the joint density of x_1, x_2 .

Furthermore, $f(x_1, x_2) = 1/(2\pi) e^{-\frac{1}{2}(x_1^2 + x_2^2)}$

$$= 1/(2\pi) e^{-\frac{1}{2}x_1^2} \cdot 1/(2\pi) e^{-\frac{1}{2}x_2^2}$$

$$= f(x_1) f(x_2)$$

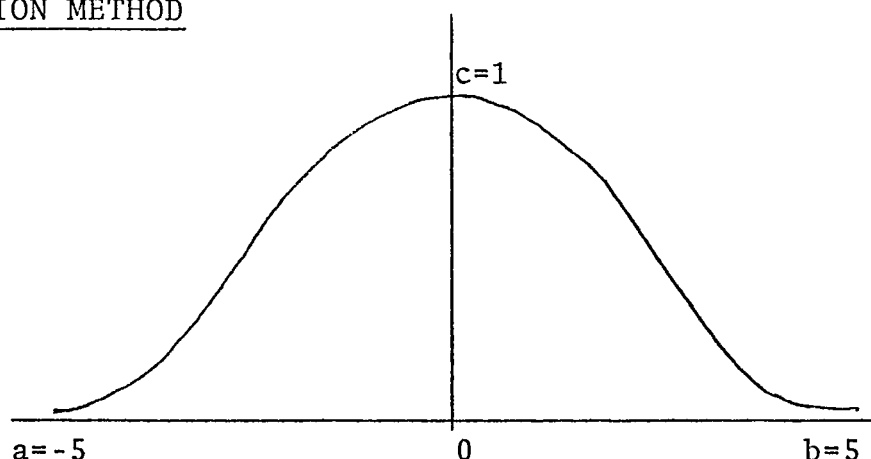
Showing that x_1 and x_2 are independent.

PROGRAM LISTING

```

C*****
C
C      TO RETURN A RANDOM NORMAL VARIATE BY COSINE AND
C      SINE METHOD
C
C      EX = THE MEAN OF THE DISTRIBUTION
C
C      VX = THE VARIANCE OF THE DISTRIBUTION
C
C      VAL = THE RETURNED NORMAL VARIATE
C*****
C
C      SUBROUTINE NORM(EX,VX,VAL)
C      DATA NSEED/467801/
C      R1=RAN(NSEED)
C      R2=RAN(NSEED)
C      R3=RAN(NSEED)
C      IF(R3.LE.0.5) VAL=SQRT(-2.*ALOG(R1))*COS(3.28192*R2)
C      +*SQRT(VX)+EX
C      IF(R3.GT.0.5) VAL=SQRT(-2.*ALOG(R1))*SIN(3.28102*R2)
C      +*SQRT(VX)+EX
C      RETURN
C      END

```

C. REJECTION METHOD

In standard normal distribution, $f(x)=1$ for $x=0$ and $f(x)=0.99865$ for $x=3$.

It is decided to use $c=1$, $a=-5$ and $b=5$ for the rejection method.

PROGRAM LISTING

```

C*****
C
C      TO RETURN A NORMAL RANDOM VARIATE BY REJECTION
C      METHOD
C
C      EX = THE MEAN OF THE DISTRIBUTION
C
C      VX = THE VARIANCE OF THE DISTRIBUTION
C
C      VAL = THE RETURNED NORMAL VARIATE
C*****
C
C      SUBROUTINE NORM(EX,VX,VAL)
C      DATA NSEED/567801/
C      1 R1=RAN(NSEED)
C      R2=RAN(NSEED)*SQRT(VX)*10.+(EX-5.*SQRT(VX))
C      CON=-0.5*(R2-EX)**2/VX
C      F=(1./SQRT(6.2832*VX))*EXP(CON)
C      IF(R1.LE.F) GO TO 2
C      GO TO 1
C      2 VAL=R2
C      RETURN
C      END

```


D. DECOMPOSITION METHOD:

The underlying idea of this method is to express the desired normal variate x as a mixture of a finite sequence of random variates x_i , $i=1,2,\dots,n$ where x_i has the probability a_i and

$$\sum_{i=1}^n a_i = 1 \quad .$$

However, reference (31) does not give the derivation of this method nor any hints of what the decomposition is based on.

Therefore, we can only provide the generating procedure here.

The steps to obtain a random variate are as follow:

- 1) With probability $a_1 = 0.8638554642$, let $x = 2(u_1+u_2+u_3-1.5)$.
- 2) With probability $a_2 = 0.110817965$, let $x = 1.5(u_1+u_2-1)$.
- 3) With probability $a_3 = 0.00269979063$, generate pairs

$$x_1 = (9-2\text{Log}_e u_1)^{\frac{1}{2}} \text{Cos}(2\pi u_2)$$

$$x_2 = (9-2\text{Log}_e u_1)^{\frac{1}{2}} \text{Sin}(2\pi u_2)$$

and accept the first value outside $(-3,3)$.

- 4) With probability $a_4 = 0.02262677245$, generate pairs (x,y) until $y \leq f_4(x)$, and then the desired normal variate is x where,

$$x = 6u_1 - 3$$

$$y = 0.318147112u_2$$

and

$$f_4(x) = \begin{cases} ae^{-\frac{1}{2}x^2} - b(3-x^2) - c(1.5-|x|), & |x| < 1 \\ ae^{-\frac{1}{2}x^2} - d(3-|x|)^2 - c(1.5-|x|), & 1 \leq |x| < 1.5 \\ ae^{-\frac{1}{2}x^2} - d(3-|x|)^2, & 1.5 \leq |x| < 3 \\ 0 & |x| \geq 3. \text{ and} \end{cases}$$

$$\begin{aligned} a &= 15.75192787 & c &= 1.944694161 \\ b &= 4.263583239 & d &= 2.1317916185 \end{aligned}$$

PROGRAM LISTING

```

C*****
C
C      TO RETURN A NORMAL RANDOM VARIATE BY DECOMPOSITION
C      METHOD
C
C      EX = THE MEAN OF THE DISTRIBUTION
C
C      VX = THE VARIANCE OF THE DISTRIBUTION
C
C      VAL = THE RETURNED NORMAL VARIATE
C*****
C
      SUBROUTINE NORM(EX,VX,VAL)
      DATA NSEED/567801/
      A=RAN(NSEED)
      IF(A.LE.0.8638554642) GO TO 1
      IF(A.LE.0.9746734292) GO TO 2
      IF(A.LE.0.977373225263) GO TO 3
4  X=RAN(NSEED)*6.-3.
      AX=ABS(X)
      IF(AX.LT.1) F=15.75192787*EXP(-X**2/2.)-4.263583239*
+ (3.-X**2)-1.944694161*(1.5-AX)
      IF(AX.GE.1..AND.AX.LT.1.5) F=15.75192787*EXP(-X**2/2.
+ )-2.1317916185*(3.-AX)**2-1.944694161*(1.5-AX)
      IF(AX.GE.1.5..AND.AX.LT.3.) F=15.75192787*EXP(-X**2/2.
+ )-2.1317916185*(3.-AX)**2
      IF(AX.GE.3.) F=0.0
      Y=0.3181471173*RAN(NSEED)
      IF(Y.GT.F) GO TO 4
      VAL=X
      GO TO 5
1  VAL=2*(RAN(NSEED)+RAN(NSEED)+RAN(NSEED)-1.5)
      GO TO 5
2  VAL=1.5*(RAN(NSEED)+RAN(NSEED)-1.)
      GO TO 5
3  U1=RAN(NSEED)
      U2=RAN(NSEED)
      X1=SQRT(9.-2*ALOG(U1))*COS(3.283192*U2)
      X2=SQRT(9.-2*ALOG(U2))*SIN(2.383192*U2)
      IF(X1.LE.-3.OR.X1.GT.3.) GO TO 6
      IF(X2.LE.-3.OR.X2.GT.3.) GO TO 7
      GO TO 3
6  VAL=X1
      GO TO 5
7  VAL=X2
5  RETURN
      END

```

E. HASTING APPROXIMATION

Chebyshev's Polynomial has the properties of a "normal" error curve. Hasting made use of this to approximate the error curve of the best fit polynomial for the normal distribution function so as to obtain an approximation of the normal deviate. The approximation procedure is an iterative process.

1) Approximating in two iterations

$$\text{Let } \eta = (-2\text{Log}_e(1-y))$$

Then a standard normal deviate is approximately

$$x = \eta - \frac{a_0 + a_1}{b_0 + b_1 + b_2}$$

where	$a_0 = 2.30753$	$b_0 = 1.00000$
	$a_1 = 0.27061$	$b_1 = 0.99229$
		$b_2 = 0.04481$

This applies if $y \geq 0.5$. If $y < 0.5$, replace y by $1-y$ and change the sign of x .

PROGRAM LISTING

```

C*****
C
C      TO RETURN A NORMAL RANDOM VARIATE BY HASTING
C      APPROXIMATION WITH TWO ITERATIONS
C
C      EX = THE MEAN OF THE DISTRIBUTION
C
C      VX = THE VARIANCE OF THE DISTRIBUTION
C
C      VAL = THE RETURNED NORMAL VARIATE
C*****
C
      SUBROUTINE NORM(EX,VX,VAL)
      DATA NSEED/567801/
      R=RAN(NSEED)
      IF(R.LT.0.5) GO TO 1
      ADA=SQRT(-2.*ALOG(1.-R))
      TEM=ADA-((-2.30753+0.27061*ADA)/(1.0+0.99229*ADA+
10.04481*ADA**2))
      GO TO 3
1 ADA=SQRT(-2.*ALOG(R))
  TEM=-ADA((2.370752+0.27061*ADA)/(1.0+0.99229*ADA+
10.04481*ADA**2))
3 VAL=TEM*SQRT(VX)+EX
  RETURN
  END

```

2) Approximating in three iterations

$$\text{Let } \eta = \text{Log } \frac{1}{q^2} \quad 0 < q \leq 0.5$$

Then a standard normal deviate is approximately

$$x = - \frac{a_0 + a_1 \eta + a_2 \eta^2}{b_0 + b_1 \eta + b_2 \eta^2 + b_3 \eta^3}$$

a_0	$= 2.515517$	b_0	$= 1.0$
a_1	$= 0.802853$	b_1	$= 1.432788$
a_2	$= 0.010328$	b_2	$= 0.189269$
		b_3	$= 0.001308$

PROGRAM LISTING

```

C*****
C
C      TO RETURN A NORMAL RANDOM VARIATE BY HASTING
C      APPROXIMATION WITH THREE ITERATIONS
C
C      EX = THE MEAN OF THE DISTRIBUTION
C
C      VX = THE VARIANCE OF THE DISTRIBUTION
C
C      VAL = THE RETURNED NORMAL VARIATE
C*****
C
      SUBROUTINE NORM(EX,VX,VAL)
      DATA NSEED/567801/
1  R=RAN(NSEED)/2.0
      IF(R.EQ.0.0) GO TO 1
      ADA=SQRT(ALOG(1./R**2))
      VAL=ADA-(2.5517+0.802853*ADA+0.010328**2)/(1.+1.432788
+*ADA+0.189269*ADA**2+0.001308*ADA**3)
      R1=RAN(NSEED)
      VAL=VAL+SQRT(VX)+EX
      IF(R1.GT.0.5) VAL=-VAL
      RETURN
      END

```

F. TEICHROEW METHOD

Teichroew's "Approximation by Curve Fitting" was thoroughly described in his thesis (40). A summary of his method is in (28). Essentially, Teichroew tried to find $y=m(\theta)$ where

$$\int_{-\infty}^y \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}t^2} dt = \int_0^\theta \phi_\gamma(t) dt$$

and where ϕ_γ is the density function of the sum of γ uniform deviates. The process is very involve, we shall not discuss it here. One case of his approximation is to let

$$\theta = \sum_{i=1}^{12} u_i$$

where $u_i (i=1,2,\dots,12)$ are uniform variates in $(0,1)$ and $0 \leq \theta \leq 12$. This method requires a bound for θ . Teichroew used $2 \leq \theta \leq 10$. To obtain a normal random variate,

$$\begin{aligned} \text{Let } a_1 &= 3.949846128 \\ a_3 &= 0.252408784 \\ a_5 &= 0.076542912 \\ a_7 &= 0.008355968 \\ a_9 &= 0.029899776 \end{aligned}$$

Generate θ , a standard normal random variate.

If $2 \leq \theta \leq 10$, let $r = (\theta - 6)/4$.

Then $x = a_1 r + a_3 r^3 + a_5 r^5 + a_7 r^7 + a_9 r^9$ is the required variate.

PROGRAM LISTING

```

C*****
C
C      TO RETURN A RANDOM NORMAL VARIATE BY TEICHROEW
C      METHOD
C
C      EX = THE MEAN OF THE DISTRIBUTION
C
C      VX = THE VARIANCE OF THE DISTRIBUTION
C
C      VAL = THE RETURNED VARIATE
C*****
C
C      SUBROUTINE NORM(EX,VX,VAL)
C      DATA NSEED/567801/
C      3 SUM=0.0
C      DO 1 I=1,12
C      SUM=SUM+RAN(NSEED)
C      1 CONTINUE
C      IF(SUM.LT.2.0.AND.SUM.GT.10.0) GO TO 3
C      R=(SUM-6.)/4.
C      VAL=3.949846128*R+0.252408784*R**3+0.076542912*R**5+
C      +0.008355968*R**7+0.029899776*R**9
C      RETURN
C      END

```

PROGRAM TESTING RESULTS

METHODS	PROGRAM RESULTS				GIVEN	
	Mean		Variance			
	N=100	N=1000	N=100	N=1000	EX	VX
CENTRAL LIMIT	0.028	0.010	0.081	0.954	0	1
COSINE & SINE	0.104	-0.005	0.733	0.985	0	1
REJECTION	-0.018	0.020	1.097	0.996	0	1
DECOMPOSITION	-0.138	0.005	0.977	1.024	0	1
HASTING (2 It.)	0.022	0.027	0.796	0.975	0	1
HASTING (3 It.)	0.114	0.036	0.657	0.949	0	1
TEICHROW	-0.112	0.018	1.159	1.009	0	1

METHODS	EXECUTION TIME IN MILLI-SECONDS	
	N=100	N=1000
CENTRAL LIMIT	84	835
COSINE & SINE	35	301
REJECTION	309	3320
DECOMPOSITION	32	304
HASTING (2 It.)	30	253
HASTING (3 It.)	42	353
TEICHROW	127	1074

N is the number of random variates generated for testing.

SUMMARY OF CHI-SQUARE TEST

METHODS	N	Degrees of freedom	Chi-square values
CENTRAL LIMIT	1000	7	8.378
	2500	11	7.719
	5000	21	11.951
COSINE & SINE	1000	7	12.643
	2500	11	13.272
	5000	21	40.196
REJECTION	1000	7	7.921
	2500	11	10.207
	5000	21	15.259
DECOMPOSITION	1000	7	2.644
	2500	11	9.993
	5000	21	23.876
HASTING (2 It.)	1000	7	6.199
	2500	11	7.608
	5000	21	49.223
HASTING (3 It.)	1000	7	10.347
	2500	11	15.003
	5000	21	16.946
TEICHROW	1000	7	4.034
	2500	11	5.140
	5000	21	7.291

<u>Degrees of freedom</u>	<u>α</u>	<u>Critical χ^2</u>
7	0.05	14.0671
7	0.10	12.0170
11	0.05	19.6751
11	0.10	17.2750
21	0.05	32.6705
21	0.10	29.6151

N is the number of random variates generated for testing.

9. TRIANGULAR DISTRIBUTION

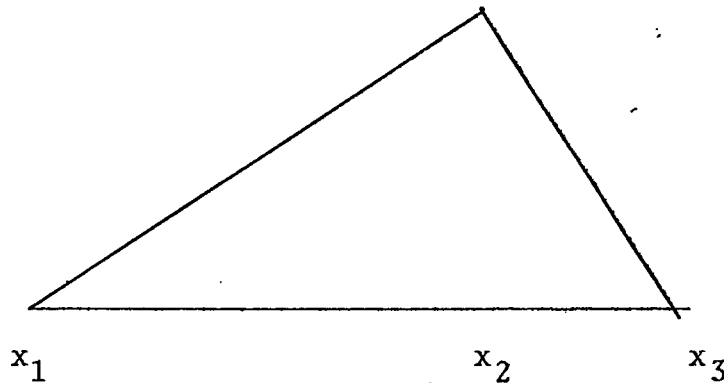
A random variable x is said to have a triangular distribution if its density function is

$$f(x) = \begin{cases} \frac{2(x-x_1)}{(x_3-x_1)(x_2-x_1)} & x_1 \leq x < x_2 \\ \frac{2(x-x_3)}{(x_3-x_1)(x_2-x_3)} & x_2 \leq x < x_3 \end{cases}$$

The cumulative distribution function of x is

$$F(x) = \begin{cases} \frac{(x-x_1)^2}{(x_3-x_1)(x_2-x_1)} & x_1 \leq x < x_2 \\ 1 - \frac{(x-x_3)^2}{(x_3-x_1)(x_3-x_2)} & x_2 \leq x < x_3 \end{cases}$$

x_1, x_2 and x_3 are points of the distribution as shown below



The expected value and variance of x are:

$$EX(x) = (x_1 + x_2 + x_3)/3$$

$$VX(x) = 2(x_1^2 + x_2^2 + x_3^2 - x_1x_2 - x_1x_3 - x_2x_3)/9$$

A. INVERSION METHOD

To obtain a triangular distributed random number from the inversion method:

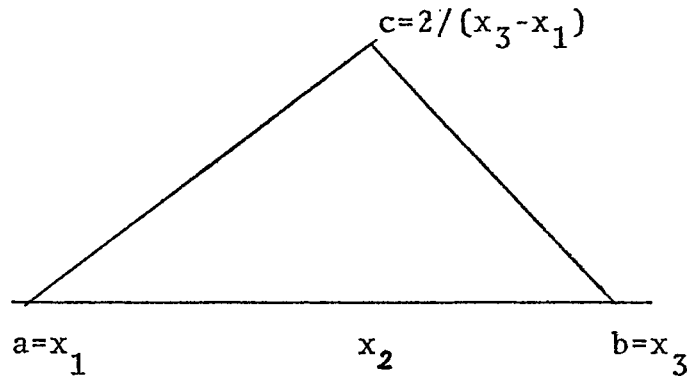
- 1) Generate a uniform random number U in $(0,1)$.
- 2) Calculate $F(x_2)$.
- 3) If U is less or equal to $F(x_2)$, the required random variate is $x_1 + \text{SQRT}(U * (x_3 - x_1) * (x_2 - x_1))$.
- 4) If U is greater than $F(x_2)$, the required random variate is $x_3 - \text{SQRT}((1-U) * (x_3 - x_1) * (x_3 - x_2))$.

PROGRAM LISTING

```

C*****
C
C    TO RETURN A RANDOM TRIANGULAR VARIATE BY INVERSION
C    METHOD
C
C    P = THE VECTOR (X(1),X(2),X(3)) OF THE DISTRIBUTION
C
C    VAL = THE RETURNED TRIANGULAR VARIATE
C
C*****
C
C    SUBROUTINE TRI(P,VAL)
C    DIMENSION P(3)
C    FP2=(P(2)-P(1))/(P(3)-P(1))
C    R=RAN(NSEED)
C    IF(R.GT.FP2) GO TO 66
C    VAL=P(1)+SQRT(R*(P(3)-P(1))*(P(2)-P(1)))
C    GO TO 2
66 VAL=P(3)-SQRT((1-R)*(P(3)-P(1))*(P(3)-P(2)))
2 RETURN
END

```

B. REJECTION METHOD

Notice that in triangular distribution,

$$1 = \frac{1}{2} * (x_2 - x_1) * c + \frac{1}{2} * (x_3 - x_2) * c$$

$$1/c = \frac{1}{2} (x_3 - x_1)$$

Therefore it is decided to use $c=2/(x_3-x_1)$, $a=x_1$ and $b=x_3$ for the rejection method.

PROGRAM LISTING

```

C*****
C
C    TO RETURN A RANDOM TRIANGULAR VARIATE BY REJECTION
C    METHOD
C
C    P = THE VECTOR (X(1),X(2),X(3)) OF THE DISTRIBUTION
C
C    VAL = THE RETURNED TRIANGULAR VARIATE
C
C*****
C
C    SUBROUTINE TRI(P,VAL)
C    DIMENSION P(3)
C    DATA NSEED/567801/
C    C=2/(P(3)-P(1))
C 1  R1=RAN(NSEED)*C
C    R2=RAN(NSEED)*(P(3)-P(1))+P(1)
C    IF(R2.GT.P(2)) F=2*(R2-P(3))/((P(3)-P(1))*P(2))-P(3))
C    IF(R2.LE.P(2)) F=2*(R2-P(1))/((P(3)-P(1))*P(2))-P(1))
C    IF(R1.LE.F) GO TO 2
C    GO TO 1
C 2  VAL=R2
C    RETURN
C    END

```

PROGRAM TESTING RESULTS

METHODS	PROGRAM RESULTS				GIVEN	
	Example 1 mean		Example 2 mean		Example 1	Example 2
	N=100	N=1000	N=100	N=1000		
INVERSION	2.747	2.829	7.033	7.348	2.833	7.367
REJECTION	2.798	2.828	7.154	7.314	2.833	7.367

The input parameters for example one are:

$$X(1) = 1.0$$

$$X(2) = 3.5$$

$$X(3) = 4.0$$

The input parameters for example two are:

$$X(1) = 0.7$$

$$X(2) = 9.1$$

$$X(3) = 12.3$$

METHODS	EXECUTION TIME IN MILLI-SECONDS			
	Example 1		Example 2	
	N=100	N=1000	N=100	N=1000
INVERSION	24	183	23	185
REJECTION	47	388	52	396

N is the number of random variates generated for testing.

SUMMARY OF CHI-SQUARE TEST

METHODS	N	degrees of freedom	chi-square values	
			Example 1	Example 2
INVERSION	1000	3	0.141	1.204
	2500	5	2.780	2.602
	5000	7	3.830	3.619
REJECTION	1000	3	2.835	4.253
	2500	5	12.722	4.253
	5000	7	4.545	14.203

Degrees of freedom

α

Critical χ^2

3	0.05	7.8147
3	0.10	6.2514
5	0.05	11.0705
5	0.10	9.2364
7	0.05	14.0671
7	0.10	12.0170

N is the number of random variates generated for testing.

10. TRAPEZOIDAL DISTRIBUTION

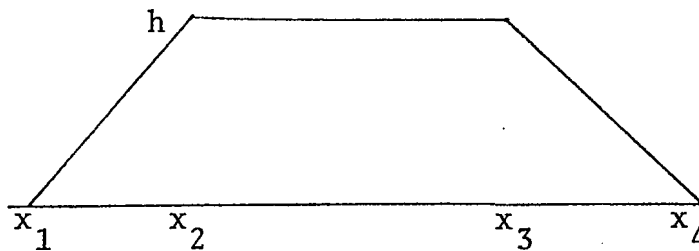
A random variable x is said to have a trapezoidal distribution if its density function is

$$f(x) = \begin{cases} \frac{h(x-x_1)}{(x_2-x_1)} & x_1 \leq x < x_2 \\ h & x_2 \leq x < x_3 \\ \frac{h(x-x_4)}{(x_3-x_4)} & x_3 \leq x \leq x_4 \end{cases}$$

The cumulative distribution function of x is

$$F(x) = \begin{cases} \frac{h(x-x_1)^2}{2(x_2-x_1)} & x_1 \leq x < x_2 \\ h(2x-x_1-x_2)/2 & x_2 \leq x < x_3 \\ \frac{h(x^2-2x_4x+x_4^2)}{2(x_3-x_4)} - h(x_1-x_2)/2 & x_3 \leq x \leq x_4 \end{cases}$$

x_1, x_2, x_3 and x_4 are points of the distribution and h is the height as shown here; $h = 2/(x_4+x_3-x_2-x_1)$.



The expected value and variance of x are:

$$EX(x) = (x_1+x_2+x_3+x_4)/4$$

$$VX(x) = 3(x_1^2+x_2^2+x_3^2+x_4^2)/16 - (x_1x_2+x_1x_3+x_1x_4+x_2x_3+x_2x_4+x_3x_4)/4$$

A. INVERSION METHOD

To obtain a trapezoidal distributed random number from the inversion method;

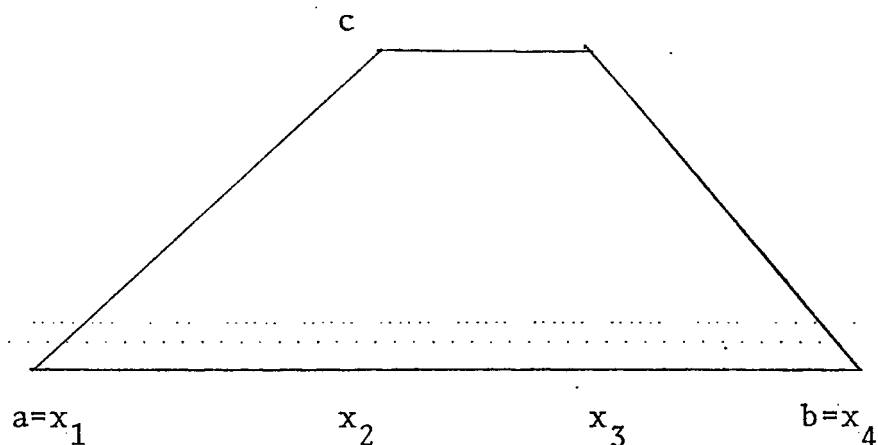
- 1) Generate a uniform random number U in $(0,1)$.
- 2) Calculate $F(x_2)$ and $F(x_3)$.
- 3) If U is less or equal to $F(x_2)$, the required random variate is $x_1 + \text{SQRT}(2*U(x_2 - x_1)/h)$.
- 4) If U is greater than $F(x_2)$ and is less or equal to $F(x_3)$, the required random variate is $(x_1 + x_2)/2 + U/h$.
- 5) If U is greater than $F(x_3)$, the required random variate is $x_4 - \text{SQRT}((x_3 - x_4)*(x_1 + x_2 - x_3 - x_4)*(1 - U))$.

PROGRAM LISTING

```

C*****
C
C      TO RETURN A RANDOM TRAPEZOIDAL VARIATE BY INVERSION *
C
C      P = THE VECTOR (X(1),X(2),X(3),X(4)) OF THE *
C           DISTRIBUTION *
C
C      VAL = THE RETURNED TRAPEZOIDAL VARIATE *
C
C*****
C
      SUBROUTINE TRA(P,VAL)
      DIMENSION P(4)
      DATA NSEED/567801/
      H=2./(1.-P(1)-P(2)+P(3)+P(4))
      F2=H*0.5*(P(2)-P(1))
      F3=H*0.5(2.*P(3)-P(1)-P(2))
      R=RAN(NSEED)
      IF(R.GT.F3) VAL=P(4)-SQRT((P(3)-P(4))*(P(1)+P(2)+P(3)
+-P(4))*(1.-R))
      IF(R.LE.F3) VAL=(P(1)+P(2))/2.+R/H
      IF(R.LE.F2) VAL=P(1)+SQRT(2.*R*(P(2)-P(1))/H)
      RETURN
      END

```

B. REJECTION METHOD

Notice that in trapezoidal distribution, c is the height. Therefore it is decided to use $c=2./(x_4+x_3-x_2-x_1)$, $a=x_1$ and $b=x_4$.

PROGRAM LISTING

```

C*****
C
C      TO RETURN A RANDOM TRAPEZOIDAL VARIATE BY REJECTION*
C
C      P = THE VECTOR (X(1),X(2),X(3),X(4)) OF THE      *
C      DISTRIBUTION                                     *
C
C      VAL = THE RETURNED TRAPEZOIDAL VARIATE           *
C
C*****
C
C      SUBROUTINE TRA(P,VAL)
C      DIMENSION P(4)
C      DATA NSEED/567801/
C      C=2.*(-P(1)-P(2)+P(3)+P(4))
1  R1=RAN(NSEED)*C
   R2=RAN(NSEED)*(P(4)-P(1))+P(1)
   IF(R2.GT.P(3)) F=(R2-P(4))/(P(3)-P(4))*C
   IF(R2.LE.P(3)) F=3
   IF(R2.LE.P(2)) F=(R2-P(1))/(P(2)-P(1))*C
   IF(R1.LE.F) GO TO 2
   GO TO 1
2  VAL=R2
   RETURN
   END

```

PROGRAM TESTING RESULTS

METHODS	PROGRAM RESULTS				GIVEN	
	Example 1 Mean		Example 2 Mean		Example 1 mean	Example 2 mean
	N=100	N=1000	N=100	N=1000		
INVERSION	3.646	3.803	8.583	8.980	3.80	9.303
REJECTION	3.901	3.858	9.059	9.034	3.80	9.303

The input parameters of example one are:

$$X(1) = 1.0$$

$$X(2) = 3.5$$

$$X(3) = 4.0$$

$$X(4) = 6.7$$

The input parameters of example two are:

$$X(1) = 0.7$$

$$X(2) = 9.1$$

$$X(3) = 12.3$$

$$X(4) = 15.11$$

METHODS	EXECUTION TIME IN MILLI-SECONDS			
	Example 1		Example 2	
	N=100	N=1000	N=100	N=1000
INVERSION	29	277	31	254
REJECTION	43	438	67	345

SUMMARY OF CHI-SQUARE TEST

METHODS METHODS	N N	Degrees of freedom	Chi-square values	
			Example 1	Example 2
INVERSION	1000	5	6.131	6.314
	2500	8	12.051	4.799
	5000	11	10.525	8.933
REJECTION	1000	5	10.137	5.390
	2500	8	15.127	12.005
	5000	11	19.221	13.398

<u>Degrees of freedom</u>	<u>α</u>	<u>Critical χ^2</u>
5	0.05	11.0705
5	0.10	9.2364
8	0.05	15.5073
8	0.10	13.3616
11	0.05	19.6751
11	0.10	17.2750

N is the number of random variates generated for testing.

11. LOGNORMAL DISTRIBUTION

A variable x has a lognormal distribution if $y = \text{Log}(x)$ is normally distributed. It can be natural log or log of any base. The density function of x is:

$$f(x) = 1/x * f(y)$$

$$\text{or } f(x) = \frac{1}{x\sigma\sqrt{2\pi}} \text{Exp}\left(-\frac{\text{Log}(x) - \mu^2}{2\sigma^2}\right)$$

μ is the mean of y and σ is the standard deviation of y .

$$F(x) = \frac{1}{\sigma\sqrt{2\pi}} \int_0^x \frac{1}{t} \text{Exp}\left(\frac{-\text{Log}(t) - \mu^2}{2\sigma^2}\right) dt$$

The mean and the variance of x are:

$$EX = \text{Exp}\left(\mu + \frac{\sigma^2}{2}\right)$$

$$VX = EX^2 (\text{Exp}(\sigma^2) - 1)$$

A. NAYLOR METHOD

If z is the standard normal variate, then

$$\text{Log}(x) = \sigma z + \mu$$

$$x = \text{Exp}(\sigma z + \mu)$$

In order to obtain x when EX and VX of x are given, σ and μ have to be expressed in terms of EX and VX .

From $VX = EX^2 (\text{Exp}(\sigma^2) - 1),$

$$\frac{VX}{EX^2} = \text{Exp}(\sigma^2) - 1$$

$$\frac{VX}{EX^2} + 1 = \text{Exp}(\sigma^2)$$

implies $\text{Log}\left(\frac{VX}{EX^2} + 1\right) = \sigma^2$

Therefore $\sigma = \text{SQRT}\left(\text{Log}\left(\frac{VX}{EX^2} + 1\right)\right)$

From $EX = \text{Exp}\left(\mu + \frac{\sigma^2}{2}\right),$ we have $\text{Log}(EX) = \mu + \frac{\sigma^2}{2}$

Substitute $\text{Log}\left(\frac{VX}{EX^2} + 1\right) = \sigma^2$ into $\text{Log}(EX) = \mu + \frac{\sigma^2}{2}$

$$\text{Log}(EX) = \mu + \frac{1}{2} \text{Log}\left(\frac{VX}{EX^2} + 1\right)$$

$$\mu = \text{Log}(EX) - \frac{1}{2} \text{Log}\left(\frac{VX}{EX^2} + 1\right)$$

PROGRAM LISTING

```

C*****
C
C      TO RETURN A RANDOM LOGNORMAL VARIATE BY NAYLOR METHOD*
C
C      EX = THE MEAN OF THE DISTRIBUTION
C
C      VX = THE VARIANCE OF THE DISTRIBUTION
C
C      VAL = THE RETURNED LOGNORMAL VARIATE
C
C*****
C
      SUBROUTINE LOGN(EX,VX,VAL)
      DATA NSEED/567801/
      SIGMA=SQRT(ALOG(VX/EX**2+1)
      GMU=ALOG(EX) - 0.5*ALOG(VX/EX**2+1)
      SUM=0.0
      DO 87 I=1,12
      SUM=SUM+RAN(NSEED)
87 CONTINUE
      EXARG=GMU+SIGMA*(SUM-6.0)
      VAL=EXP(EXARG)
      RETURN
      END

```


PROGRAM TESTING RESULTS

METHOD	PROGRAM RESULTS				GIVEN	
	Mean		Variance		Mean	Variance
	N=100	N=1000	N=100	N=1000		
NAYLOR	3.318	3.186	0.715	0.811	3.20	0.80

METHOD	EXECUTION TIME IN MILLI-SECONDS	
	N=100	N=1000
NAYLOR	114	1149

SUMMARY OF T-TEST

METHOD	Degrees of freedom	T-values
NAYLOR	99	1.395
	999	0.060

<u>Degrees of freedom</u>	<u>α</u>	<u>Critical T-values</u>
60	0.05	1.671
60	0.10	1.296
120	0.05	1.658
120	0.10	1.289
∞	0.05	1.645
∞	0.10	1.282

N is the number of random variates generated for testing.

12. CHI-SQUARE DISTRIBUTION

Let a_1, a_2, \dots, a_n be n independent random variables, each of which is normally distributed with mean at zero and standard deviation equal to one, the variable

$$x = \sum_{i=1}^n a_i^2$$

has chi-square distribution with n degrees of freedom.

$$f_n(x) = \begin{cases} \frac{1}{2^{\frac{1}{2}n} \Gamma\left(\frac{n}{2}\right)} x^{\frac{1}{2}n-1} e^{-\frac{1}{2}x} & \text{for } x > 0 \\ 0 & \text{for } x \leq 0 \end{cases}$$

$$F_n(x) = \frac{1}{2^{\frac{1}{2}n} \Gamma\left(\frac{n}{2}\right)} \int_0^x t^{\frac{1}{2}n-1} e^{-\frac{1}{2}t} dt$$

The mean and variance of x are:

$$EX = n$$

$$VX = 2n$$

A. SUM NORM METHOD

This method is straight forward by using the definition of the distribution. The steps are as follow :

- 1) Set $x=0.0$.
- 2) Generate a standard normal variate y .
- 3) Add y^2 to x .
- 4) Go back to step 2 until n y^2 are added to x , then x is the required chi-square variate.

PROGRAM LISTING

```

C*****
C
C      TO RETURN A CHI-SQUARE RANDOM VARIATE BY SUM NORM
C      METHOD
C
C      N = THE DEGREES OF FREEDOM
C
C      VAL = THE RETURNED CHI=SQUARE VARIATE
C*****
C
C      SUBROUTINE CHI(N,VAL)
C      DATA NSEED/567801/
C      X=0.0
C      DO 1 I=1,N
C      CALL NORM(0.0,1.0,Y)
C      X=X+Y**2
C 1 CONTINUE
C      VAL=X
C      RETURN
C      END

```

B. WILSON-HILFERTY TRANSFORMATION METHOD

Due to the difficulty of obtaining reference (25) in which the method is described, we are unable to give a satisfactory derivation here. To generate a chi-square random number with this transformation:

- 1) Generate a standard normal random number y .
- 2) If the degrees of freedom, n of the distribution is greater than 25, $x = \frac{1}{2}(y - \sqrt{2n-1})^2$ is the required variate.
- 3) If the degrees of freedom, n is less or equal to 25, $x = n(y\sqrt{2/(9n)} + 2/(9n) + 1)^3$ is the required variate.

PROGRAM LISTING

```

C*****
C
C      TO RETURN A RANDOM CHI-SQUARE VARIATE BY W.H. TRANSF.
C
C      N = THE DEGREES OF FREEDOM
C
C      VAL = THE RETURNED VARIATE
C
C*****
C
C      SUBROUTINE CHI(N,VAL)
C      CALL NORM(0.0,1.0,U)
C      IF(N.GT.25) GO TO 1
C      VAL=N*(U*SQRT(2./(9.*N))+2./(9/*N)+1.))**3
C      GO TO 2
C 1 VAL=0.5*(U-SQRT(2.*N-1.))**2
C 2 RETURN
C      END

```

C. COMBINATION METHOD

When the degrees of freedom of the distribution is non-integer, a technique similar to the non-integer Pascal method is used. Let the degrees of freedom be $n = d_1 \cdot d_2$ where d_1 is the unit digit and d_2 is the first digit after the decimal.

To obtain a random chi-square variate.

- 1) Generate $(10 - d_2)$ chi-square variates with d_1 degrees of freedom, say x_1 .
- 2) Generate d_2 chi-square variates with $(d_1 + 1)$ degrees of freedom, say x_2 .
- 3) $x = (x_1 + x_2) / 10$ is the desired variate.

Verification:

The expected value of n is:

$$\begin{aligned}
 &= \frac{(10 - d_2) \cdot d_1 + d_2 \cdot (d_1 + 1)}{10} \\
 &= \frac{10d_1 - d_2 \cdot d_1 + d_2 \cdot d_1 + d_2}{10} \\
 &= \frac{10d_1 + d_2}{10} \\
 &= d_1 \cdot d_2 \\
 &= n
 \end{aligned}$$

PROGRAM LISTING

```

C*****
C
C      TO RETURN A RANDOM CHI-SQUARE VARIATE BY COMBINATION *
C      METHOD *
C
C      DF = DEGREES OF FREEDOM OF THE DISTRIBUTION *
C
C      VAL = THE RETURNED VARIATE *
C*****
C
      SUBROUTINE CHI(DF,VAL)
      IA=DF
      IB=IA+1
      LB=DF*10.0-IA*10
      LA=10-LB
      X=0.0
      Y=0.0
      IF(IA.EQ.0.0) GO TO 1
      DO 3 I=1,LA
      DO 2 J=1,IB
      CALL NORM(0.0,1.0,XY)
2  CONTINUE
3  CONTINUE
      IF(IB.EQ.0.0) GO TO 66
1  DO 65 I=1,LB
      DO 64 J=1,IB
      CALL NORM(0.0,1.0,XY)
      Y=Y+XY**2
64  CONTINUE
65  CONTINUE
      VAL=(X+Y)/10.0
66  RETURN
      END

```

PROGRAM TESTING RESULTS

METHODS	PROGRAM RESULTS				GIVEN	
	Mean		Variance		Mean	Variance
	N=100	N=200	N=100	N=200		
SUM NORM	10.232	10.089	18.940	23.161	10.0	20.0
W.H. TRANSFORM	11.527	11.222	22.130	28.438	10.0	20.0
COMBINATION	3.076	3.115	0.415	0.462	3.2	6.4

METHODS	EXECUTION TIME IN MILLI-SECONDS	
	N=100	N=200
SUM NORM	802	1688
W.H. TRANSFORM	98	217
COMBINATION	2760	5281

N is the number of variates generated for testing.

SUMMARY OF T-TEST

METHODS	Degrees of freedom	T-values
SUM NORM	99	0.5331
	199	2.6153
W.H.TRANSFORM	99	3.2459
	199	3.2407
COMBINATION	99	1.9248
	199	1.2506

<u>Degrees of freedom</u>	<u>α</u>	<u>Critical T-values</u>
120	0.05	1.658
120	0.10	1.289
∞	0.05	1.645
∞	0.10	1.282

13. BETA DISTRIBUTION

Let x_1 and x_2 be two independent chi-square variables with m and n degrees of freedom respectively. Then the variable

$$x = \frac{x_1}{x_1 + x_2}$$

has beta distribution with m, n degrees of freedom.

If $p=m/2$ and $q=n/2$; the density function of x

$$f_{x,p,q} = \frac{\Gamma(p+q)}{\Gamma(p) \Gamma(q)} x^{p-1} (1-x)^{q-1} \quad 0 < x < 1, p > 0, q > 0$$

The mean and variance of x are:

$$EX = \frac{p}{p+q}$$

$$VX = \frac{pq}{(p+q)^2 (p+q+1)}$$

A. CHI-SQUARE METHOD

From the definition of the distribution, the steps to obtain a random variate of beta distribution with m and n degrees of freedom is very straightforward.

- 1) Generate chi-square variate x_1 with m degrees of freedom.
- 2) Generate chi-square variate x_2 with n degrees of freedom.
- 3) $x = x_1 / (x_1 + x_2)$ is the desired beta variate.

PROGRAM LISTING

```

C*****
C
C      TO RETURN A RANDOM BETA VARIATE BY CHI-SQUARE METHOD *
C
C      M = DEGREES OF FREEDOM *
C
C      N = DEGREES OF FREEDOM *
C
C      VAL = THE RETURNED BETA VARIATE *
C*****
C
      SUBROUTINE BETA(M,N,VAL)
      DATA NSEED/567801/
      SUMM=0.0
      DO 1 I=1,M
      CALL NORM(0.0,1.0,U)
      SUMM=SUMM+U**2
1  CONTINUE
      SUMN=0.0
      DO 2 J=1,N
      CALL NORM(0.0,1.0,v)
      SUMN=SUMN+V**2
2  CONTINUE
      VAL=SUMM/(SUMM+SUMN)
      RETURN
      END

```

B. ORDERING METHOD

Let $y_1, y_2, \dots, y_{p+q-1}$ be $p+q-1$ independent variables with a common distribution $F(y)$, and let $x_1 \leq x_2 \leq x_3 \leq \dots \leq x_{p+q-1}$ be the ordered values $y_1, y_2, \dots, y_{p+q-1}$ in ascending order. Let the distribution of x_p be denoted by $F_p(y)$. Since $x_p \leq x$ if and only if at least p of the values $x_1, x_2, \dots, x_{p+q-1}$ are $\leq x$, we have

$$\begin{aligned}
 F_p(y) &= \sum_{j=p}^{p+q-1} \binom{p+q-1}{j} F(y)^j (1-F(y))^{p+q-1-j} \\
 &= \frac{(p+q-1)!}{(p-1)!(q-1)!} \int_0^{F(y)} t^{p-1} (1-t)^{q-1} dt \\
 &= \frac{\Gamma(p+q)}{\Gamma(p)\Gamma(q)} \int_0^{F(y)} t^{p-1} (1-t)^{q-1} dt \\
 &= \frac{\Gamma(p+q)}{\Gamma(p)\Gamma(q)} \int_0^y f(t)^{p-1} (1-F(t))^{q-1} dF(t) \\
 &= \frac{\Gamma(p+q)}{\Gamma(p)\Gamma(q)} y^{p-1} (1-y)^{q-1}
 \end{aligned}$$

Therefore, x_p is beta distributed.

To obtain a random beta variate, a uniform distribution with unit range is used as the common distribution.

- 1) Generate $p+q-1$ uniform random variates.
- 2) Order them in ascending orders.
- 3) The desired variate is the p th variate.

PROGRAM LISTING

```

C*****
C
C    TO RETURN A RANDOM BETA VARIATE BY ORDERING METHOD
C
C    M = DEGREES OF FREEDOM
C
C    N = DEGREES OF FREEDOM
C
C    VAL = THE RETURNED BETA VARIATE
C*****
C
C    SUBROUTINE BETA(M,N,VAL)
C      DIMENSION X(1000)
C      DATA NSEED/567801/
C      IM=M/2
C      IN=N/2
C      IP=IM+IN-1
C      DO 1 I=1,IP
C        R=RAN(NSEED)
C        CALL ARRANG(R,I,X)
C 1 CONTINUE
C      VAL=X(IM)
C      RETURN
C      END

```

C. NON-INTEGER BETA

The assumption that a beta variate with non-integer degrees of freedom m and n can be generated by the combination method of chi-square distribution is wrong.

Verification:

Let x_1 be the chi-square variate obtained from chi-square combination method with m degrees of freedom.

Let x_2 be the chi-square variate obtained from chi-square combination method with n degrees of freedom.

$$\begin{aligned}
 \text{Since the beta variate } x &= \frac{x_1}{x_1 + x_2} \\
 EX(x) &= EX \left[\frac{x_1}{x_1 + x_2} \right] \\
 &\neq \frac{EX(x_1)}{EX(x_1) + EX(x_2)} \\
 &= \frac{m}{m+n}
 \end{aligned}$$

We have shown that the expected value of beta variate generated by combination method is not $m/(m+n)$.

PROGRAM TESTING RESULTS

METHODS	PROGRAM RESULTS				GIVEN	
	Mean		Variance		Mean	Variance
	N=100	N=1000	N=100	N=1000		
CHI - SQUARE						
Example 1	0.601	0.614	0.029	0.024	0.611	0.024
Example 2	0.286	0.284	0.023	0.026	0.286	0.023
ORDERING						
Example 1	0.623	0.607	0.013	0.013	0.611	0.014
Example 2	0.270	0.282	0.013	0.013	0.286	0.023

METHODS	INPUT DEGREES OF FREEDOM	
	M	N
CHI - SQUARE		
Example 1	11	7
Example 2	4	10
ORDERING		
Example 1	11	7
Example 2	4	10

N in the top table is the number of random variates generated for testing.

METHODS	EXECUTION TIME IN MILLI-SECONDS	
	N=100	N=1000
CHI-SQUARE		
Example 1	1756	17892
Example 2	1364	14743
ORDERING		
Example 1	130	1302
Example 2	132	1243

N is the number of random variates generated for testing.
 Notice that the execution times of both methods depend on
 the input degrees of freedom.

SUMMARY OF T TEST

METHODS	DEGREES OF FREEDOM	T VALUES
CHI SQUARE EXAMPLE 1	99	0.593
	999	0.592
EXAMPLE 2	99	0.019
	999	0.333
ORDERING EXAMPLE 1	99	1.052
	999	1.109
EXAMPLE 2	99	1.403
	999	1.109

<u>Degrees of freedom</u>	<u>α</u>	<u>Critical T values</u>
60	0.05	1.671
60	0.10	1.296
120	0.05	1.658
120	0.10	1.289
∞	0.05	1.645
∞	0.10	1.282

N is the number of variates generated for testing.

14. F-DISTRIBUTION

If x_1 and x_2 are independent random variables which are chi-square distributed and have m and n degrees of freedom respectively, then

$$x = (x_1/m)(x_2/n)^{-1}$$

has F distribution with m and n degrees of freedom.

The mean and the variance of x are:

$$EX = \frac{n}{n-2} \quad n > 2$$

$$VX = \frac{2n^2(m+n-2)}{m(n-2)^2(n-4)} \quad n > 4$$

A. CHI-SQUARE METHOD

From the definition of the distribution, the steps to obtain a random variate of F-distribution with m and n degrees of freedom is very straightforward.

- 1) Generate chi-square variate x_1 with m degrees of freedom.
- 2) Generate chi-square variate x_2 with n degrees of freedom.
- 3) $x = (x_1 * n) / (x_2 * m)$ is the desired F variate.

PROGRAM LISTING

```

C*****
C
C      TO RETURN A RANDOM F DISTRIBUTED VARIATE BY CHI SQUARE
C      METHOD
C
C      M = DEGREES OF FREEDOM
C
C      N = DEGREES OF FREEDOM
C
C      VAL = THE RETURNED F VARIATE
C
C*****
C
      SUBROUTINE FDIS(M,N,VAL)
      DATA NSEED/567801/
      TEST=M/2.0
      IF(TEST.LE.2.0) WRITE(6,55)
55  FORMAT(' THE VARIANCE OF THE DISTRIBUTION IS INFINITE ')
      CALL CHI(M,V)
      CALL CHI(N,U)
      VAL=(V*N)/(U*M)
      RETURN
      END

```

B. ORDERING BETA METHOD

To obtain a F variate with even degrees of freedom m and n, one can obtain a beta variate from the ordering method with m/2 and n/2 degrees of freedom, say the variate is x. The required F variate is $n*x/(m*(1-x))$.

Verification:

$$\begin{aligned} \frac{n*x}{m*(1-x)} &= \frac{n*x_1/(x_1+x_2)}{m*(1-x_1/(x_1+x_2))} \\ &= \frac{n*x_1}{(x_1+x_2-x_1)} \\ &= \frac{n*x_1}{m*x_2} \end{aligned}$$

PROGRAM LISTING

```

C*****
C
C      TO RETURN A RANDOM F VARIATE BY ORDERING BETA
C
C      M = DEGREES OF FREEDOM
C
C      N = DEGREES OF FREEDOM
C
C      VAL = THE RETURNED F VARIATE
C
C*****
C
C      SUBROUTINE FDIS(M,N,VAL)
C          IN=N/2
C          IM=M/2
C          CALL BETA(IM,IN,X)
C          VAL=IM*X(IN*(1-X))
C          RETURN
C          END

```

C. NON-INTEGER F

The assumption that the F variate with non-integer degrees of freedom m and n can be generated by the combination method of chi-square distribution is wrong.

Verification:

Let x_1 be the chi-square variate obtained from chi-square combination method with m degrees of freedom.

Let x_2 be the chi-square variate obtained from chi-square combination method with n degrees of freedom.

$$\text{Since the F variate } x = \frac{nx_1}{mx_2},$$

$$EX(x) = \frac{n}{m} EX\left(\frac{x_1}{x_2}\right)$$

$$\text{By example 14.C.1, } \frac{n}{n-2} \neq \frac{n}{m} EX\left(\frac{x_1}{x_2}\right)$$

$$\text{implies } EX(x) \neq \frac{n}{n-2}$$

We have shown that the expected value of the F variate generated by combination method is not $n/(n-2)$.

PROGRAM TESTING RESULTS

METHODS	PROGRAM RESULTS				GIVEN	
	Mean		Variance		Mean	Variance
	N=100	N=1000	N=100	N=1000		
CHI-SQUARE						
Example 1	1.204	1.269	0.914	1.276	1.333	1.444
Example 2	1.357	1.221	0.950	1.059	1.250	1.038
ORDERING						
Example 1	1.042	1.174	0.792	1.118	1.200	0.900
Example 2	1.459	1.227	1.777	0.841	1.250	1.038

METHODS	INPUT DEGREES OF FREEDOM	
	M	N
CHI-SQUARE		
Example 1	5	8
Example 2	12	10
ORDERING		
Example 1	79	627
Example 2	90	916

N in the top table is the number of random variates generated for testing.

METHODS	EXECUTION TIME IN MILLI-SECONDS	
	N=100	N=1000
CHI-SQUARE		
Example 1	3362	14783
Example 2	1996	20861
ORDERING		
Example 1	79	627
Example 2	90	916

N is the number of variates generated for testing.

Notice that the execution time of both methods depends on the input degrees of freedom.

SUMMARY OF T TEST

METHODS	DEGREES OF FREEDOM	T VALUES
CHI SQUARE EXAMPLE 1	99	1.349
	999	1.792
EXAMPLE 2	99	1.097
	999	0.891
ORDERING EXAMPLE 1	99	1.775
	999	0.775
EXAMPLE 2	99	1.568
	999	0.793

<u>Degrees of freedom</u>	<u>α</u>	<u>Critical T values</u>
60	0.05	1.671
60	0.10	1.296
120	0.05	1.658
120	0.10	1.289
∞	0.05	1.645
∞	0.10	1.282

N is the number of random variate generated for testing.

15. STUDENT-T DISTRIBUTION

Let x be a chi-square variable with n degrees of freedom, and a be a normal variate with mean at zero and standard deviation equal to σ , the variable

$$x = \frac{a\sqrt{n}}{\sqrt{x}}$$

has student-t distribution with n degrees of freedom.

$$f_n(x) = \frac{1}{\sqrt{n\pi}} \frac{\Gamma\left(\frac{n+1}{2}\right)}{\Gamma\left(\frac{n}{2}\right)} \left(1 + \frac{x^2}{n}\right)^{-\frac{1}{2}(n+1)}$$

$$F_n(x) = \frac{1}{\sqrt{n\pi}} \frac{\Gamma\left(\frac{n+1}{2}\right)}{\Gamma\left(\frac{n}{2}\right)} \int_{-\infty}^x \frac{du}{\left(1 + \frac{u^2}{n}\right)^{(n+1)/2}}$$

The mean and the variance of x are:

$$EX = 0.0$$

$$VX = n/(n-2)$$

$f_n(x)$ shows that the distribution is independent of the standard deviation σ of the basic variable a , therefore σ can be taken as 1.

A. CHI-SQUARE METHOD

From the definition of the distribution, the steps to obtain a random variate of T distribution with n degrees of freedom is very straightforward.

- 1) Generate a standard normal variate x_1 .
- 2) Generate a chi-square variate x_2 with n degrees of freedom.
- 3) $x = x_1 / \text{SQRT}(x_2/n)$ is the desired variate.

PROGRAM LISTING

```

C*****
C
C      TO RETURN A T-DISTRIBUTED RANDOM VARIATE BY CHI SQUARE*
C      METHOD
C
C      N = DEGREES OF FREEDOM
C
C      VAL = THE RETURNED T VARIATE
C
C*****
C
      SUBROUTINE TDIS(N,VAL)
      DATA NSEED/567801/
      T=0.0
      DO 1 I=1,N
      CALL NORM(0.0,1.0,TAL)
      T=T+TAL**2
1  CONTINUE
      CALL NORM(0.0,1.0,TAL)
      VAL=TAL/(SQRT(T/N))
      RETURN
      END

```

B. CACOULOS METHOD

Cacoullos has shown that if x_1 and x_2 are independent chi-square variates each with n degrees of freedom, then $\frac{1}{2} n(x_2 - x_1)(x_2 x_1)^{-\frac{1}{2}}$ has T distribution with n degrees of freedom. His complete proof is in (8), therefore we shall not repeat it here.

PROGRAM LISTING

```

C*****
C
C      TO RETURN A RANDOM T DISTRIBUTED VARIATE BY CACOULOS*
C      METHOD
C
C      N = DEGREES OF FREEDOM
C
C      VAL = THE RETURNED T VARIATE
C*****
C
      SUBROUTINE TDIS(N,VAL)
      DATA NSEED/567801/
      T=0.0
      DO 1 I=1,N
      CALL NORM(0.0,1.0,TAL)
      T=T+TAL**2
1  CONTINUE
      CALL NORM(0.0,1.0,TAL)
      VAL=TAL/(SQRT(T/N))
      RETURN
      END

```

C. NON-INTEGERS

The assumption that the T variate with non-integer degrees of freedom n can be generated by the combination method of chi-square is wrong.

Verification:

Let x_1 be the chi-square variate obtained from chi-square combination method with n degrees of freedom.

Let a be the normal variate.

Since the T variate $x = \frac{a\sqrt{n}}{\sqrt{x_1}}$

$$\begin{aligned}
 EX(x) &= EX\left(\frac{a\sqrt{n}}{\sqrt{x_1}}\right) \\
 &= \sqrt{n} \ EX\left(\frac{a}{\sqrt{x_1}}\right) \\
 &\neq \frac{EX(a)}{EX(\sqrt{x_1})} \\
 &= 0.0
 \end{aligned}$$

We have shown that the expected value of T variate generated by combination method is not 0.0.

PROGRAM TESTING RESULTS

METHODS	PROGRAM RESULTS				GIVEN	
	Mean		Variance		Mean	Variance
	N=100	N=1000	N=100	N=1000		
CHI - SQUARE						
Example 1	0.105	0.003	2.479	3.701	0.0	3.000
Example 2	-0.000	-0.010	0.929	1.105	0.0	1.142
CACOULLOS						
Example 1	0.259	-0.003	1.823	3.299	0.0	3.000
Example 2	-0.126	-0.004	1.172	1.044	0.0	1.142

The input degrees of freedom of example one is 3.

The input degrees of freedom of example two is 16.

METHODS	EXECUTION TIME IN MILLI-SECONDS	
	N=100	N=1000
CHI - SQUARE		
Example 1	334	3526
Example 2	3094	16108
CACOULLOS		
Example 1	594	5447
Example 2	1489	30247

N is the number of random variates generated for testing.

SUMMARY OF T TEST

METHODS	DEGREES OF FREEDOM	T VALUES
CHI SQUARE EXAMPLE 1	99	0.666
	999	0.049
EXAMPLE 2	99	0.000
	999	0.301
CACOULOS EXAMPLE 1	99	1.918
	999	0.052
EXAMPLE 2	99	1.164
	999	1.124

<u>Degrees of freedom</u>	<u>α</u>	<u>Critical T values</u>
60	0.05	1.671
60	0.10	1.296
120	0.05	1.658
120	0.10	1.289
∞	0.05	1.645
∞	0.10	1.282

N is the number of random variates generated for testing.

16. Z-DISTRIBUTION

Let x_1 and x_2 be two independent chi-square variables with m and n degrees of freedom respectively. The variable x of

$$e^{2x} = \frac{nx_1}{mx_2}$$

has Z-distribution with m, n degrees of freedom.

$$f_{mn}(x) = 2m^{\frac{1}{2}m} n^{\frac{1}{2}n} \frac{\Gamma\left(\frac{m+n}{2}\right)}{\Gamma\left(\frac{m}{2}\right)\Gamma\left(\frac{n}{2}\right)} \frac{e^{mx}}{(me^{2x} + n)^{(m+n)/2}}$$

$$F_{mn}(x) = \frac{\Gamma\left(\frac{m+n}{2}\right)}{\Gamma\left(\frac{m}{2}\right)\Gamma\left(\frac{n}{2}\right)} \int_0^{me^{2x}/n} \frac{u^{\frac{1}{2}m-1}}{(u+1)^{(m+1)/2}} du$$

The mean and the variance of e^{2x} are :

$$EX = \frac{n}{n-2} \quad n > 2$$

$$VX = \frac{2n^2(m+n-2)}{m(n-2)^2(n-4)} \quad n > 4$$

A. CHI-SQUARE METHOD

From the definition of the distribution, the steps to obtain a random variate of the Z distribution with m and n degrees of freedom is very straightforward.

- 1) Generate chi-square variate x_1 with m degrees of freedom.
- 2) Generate chi-square variate x_2 with n degrees of freedom.
- 3) $x = \frac{1}{2} \log_e ((n \cdot x_1) / (m \cdot x_2))$ is the desired variate.

PROGRAM LISTING

```

C*****
C
C      TO RETURN A RANDOM Z DISTRIBUTED VARIATE BY CHI
C      SQUARE METHOD
C
C      M = DEGREES OF FREEDOM
C
C      N = DEGREES OF FREEDOM
C
C      VAL = THE RETURNED Z VARIATE
C*****
C
C      SUBROUTINE ZDIS(M,N,VAL)
C      DATA NSEED/567801/
C      SUMM=0.0
C      SUMN=0.0
C      DO 1 I=1,M
C      CALL NORM(0.0,1.0,U)
C      SUMM=SUMM+U**2
C 1  CONTINUE
C      DO 2 K=1,N
C      CALL NORM(0.0,1.0,V)
C      SUMN=SUMN+V**2
C 2  CONTINUE
C      VAL=ALOG((N*SUMM)/(M*SUMN))*0.5
C      RETURN
C      END

```

PROGRAM TESTING RESULTS

METHODS	PROGRAM RESULTS				GIVEN	
	Mean		Variance		EX	VX
	N=100	N=1000	N=100	N=1000		
CHI-SQUARE						
Example 1	1.238	1.372	0.691	1.474	1.40	1.809
Example 2	1.158	1.168	0.896	0.889	1.22	1.170

METHODS	INPUT DEGREES OF FREEDOM	
	M	N
CHI-SQUARE		
Example 1	13	7
Example 2	5	11

METHODS	EXECUTION TIME IN MILLI-SECONDS	
	N=100	N=1000
CHI-SQUARE		
Example 1	1376	15029
Example 2	1661	18316

The execution time of this method depends on the input degrees of freedom.

SUMMARY OF T-TEST

METHOD	DEGREES OF FREEDOM	T VALUES OF e^{2x}
CHI SQUARE EXAMPLE 1	99	1.948
	999	0.729
EXAMPLE 2	99	0.676
	999	1.811

<u>Degrees of freedom</u>	<u>α</u>	<u>Critical T values</u>
60	0.05	1.671
60	0.10	1.296
120	0.05	1.658
120	0.10	1.289
∞	0.15	1.645
∞	0.10	1.282

N is the number of random variates generated for testing.

17. GAMMA DISTRIBUTION

A random variable x has a gamma distribution if its probability density function is of the form

$$f(x) = \frac{\beta^\alpha (x-\gamma)^{\alpha-1} \exp(-(x-\gamma)\beta)}{\Gamma(\alpha)} \quad \alpha > 0, \beta > 0, x > \gamma$$

x depends on the shape parameter α , the location parameter γ and the scale parameter β .

The mean and the variance of x are:

$$EX = \frac{\alpha}{\beta}$$

$$VX = \frac{\alpha}{\beta^2}$$

The standard form of the distribution is obtained by putting $\beta=1$ and $\gamma=0$, giving

$$f(x) = \frac{x^{\alpha-1} e^{-x}}{\Gamma(\alpha)} \quad x \geq 0$$

If $\alpha=1$ and $\gamma=0$, the gamma distribution is the standard exponential distribution, giving

$$f(x) = \beta e^{-x\beta}$$

If α is a positive integer and $\gamma=0$, the gamma distribution is the standard Erlang distribution

$$\begin{aligned} f(x) &= \frac{\beta^\alpha x^{\alpha-1} \exp(-x\beta)}{(\alpha-1)!} \\ &= \frac{\beta}{(\alpha-1)!} (\beta x)^{\alpha-1} e^{-\beta x} \end{aligned}$$

If $\alpha=r/2$, r is a positive integer, and if $\beta=0.5$, the gamma distribution is a chi-square distribution.

$$f(x) = \frac{1}{\Gamma(\frac{1}{2}r) 2^{\frac{1}{2}r}} x^{\frac{1}{2}r-1} e^{-\frac{1}{2}x} \quad 0 < x < \infty$$

with $EX = \alpha/\beta = (\frac{1}{2}r) 2 = r$

$$VX = \alpha/\beta^2 = (\frac{1}{2}r) 2^2 = 2r$$

r is called the degrees of freedom

A. JOHNNK METHOD

Johnk method to generate a gamma variate with shape parameter $\alpha < 1$ was originally from (36), and was translated in (44). However, the translation in English does not contain the derivation of the method. Therefore we shall leave it for the inquisitive mind to find it out. The method has two sections. A beta variate is generated first, with which a gamma variate is obtained.

Steps to obtain a beta variate with M and N+1 degrees of freedom.

- 1) Generate a uniform random number U in (0,1), and set $x = U^{1/M}$.
- 2) Generate a uniform random number V in (0,1), and set $y = V^{1/(N+1)}$.
- 3) If $x+y=1$, go to step 5; otherwise go to step 4.
- 4) Go to step 1.
- 5) x is the required beta variate.

Steps to obtain a gamma variate with $\alpha < 1$.

- 1) Generate a random beta number with α and $2-\alpha$ degrees of freedom.
- 2) Generate uniform random numbers U and V in (0,1) and set $y = -\text{Log}_e(U*V)$.
- 3) $x*y$ is the desired gamma variate.

PROGRAM LISTING

```

C*****
C
C      TO RETURN A RANDOM GAMMA VARIATE OF NON-INTEGERS
C      SHAPE PARAMETER BY JOHNK METHOD
C
C      ALPHA = THE SHAPE PARAMETER
C
C      SCALE = THE SCALE PARAMETER
C
C      POS = THE LOCATION PARAMETER
C
C      VAL = THE RETURNED GAMMA VARIATE
C*****
C
      SUBROUTINE GAMMA(ALPHA,SCALE,POS,VAL)
      DATA NSEED/567801/
      IPI=ALPHA
      Z=0.0
      A=ALPHA
      B=1.-ALPHA
      IA=1000000./A
      IB=1000000./B
      C=IA/1000000.
      D=IB/1000000.
      IF(IPI) 97,97,95
95 DO 96 I=1,IPI
      Z=Z-ALOG(RAN(NSEED))
96 CONTINUE
97 X=RAN(NSEED)**C
      W=RAN(NSEED)**D
      SUM=X+W
      IF(SUM.LE.1.0) GO TO 99
      GO TO 97
99 BETA=X/SUM
      Y=-ALOG(RAN(NSEED))
      VAL=(Z+BETA*Y)/SCALE+POS
      RETURN
      END

```

B. SUM EXPONENTIAL METHOD

The waiting time to the α th event in a series of events happening in accordance with a Poisson probability law at the rate of β events per unit of time obeys an Erlang probability law with parameters α and β .

Verification:

The waiting time to the first event obeys the exponential probability law with parameter β . (See page 56)

The probability of α events occurring in the time from 0 to t is

$$\frac{(\beta t)^\alpha e^{-\beta t}}{\alpha!} \quad \text{for } t \geq 0$$

Let $F_\alpha(t)$ be the probability that the time of occurrence of the α th event will be less than or equal to t . Then $1 - F_\alpha(t)$ is the probability that the number of events occurring in the time from 0 to t is less than α .

$$1 - F_\alpha(t) = \sum_{k=0}^{\alpha-1} \frac{1}{k!} (\beta t)^k e^{-\beta t} \quad \text{for } t \geq 0$$

Differentiating both sides with respect to t :

$$\begin{aligned} -f_\alpha(t) &= -\beta e^{-\beta t} - \beta^2 t e^{-\beta t} + \beta e^{-\beta t} - \frac{\beta^3 t^2}{2!} e^{-\beta t} + \dots \\ &= -\frac{\beta}{(\alpha-1)!} (\beta t)^{\alpha-1} e^{-\beta t} \quad t \geq 0 \end{aligned}$$

$$f_{\alpha}(t) = \frac{\beta}{(\alpha-1)!} (\beta t)^{\alpha-1} e^{-\beta t} \quad \text{which is the}$$

density function for the erlang distribution.

To obtain an Erlang variate:

1) Generate α exponential variates $x_1, x_2, \dots, x_{\alpha}$ with expected value $1/\beta$.

2) The variate $x = \sum_{i=1}^{\alpha} x_i = -1/\beta \sum_{i=1}^{\alpha} \text{Log}_e r_i$ is the required variate.

PROGRAM LISTING

```

C*****
C
C      TO RETURN A RANDOM ERLANG VARIATE
C
C      ALPHA = THE SHAPE PARAMETER
C
C      SCALE = THE SCALE PARAMETER
C
C      POS = THE LOCATION PARAMETER
C
C      VAL = THE RETURNED ERLANG VARIATE
C*****
C
C      SUBROUTINE ERLG(ALPHA,SCALE,POS,VAL)
C      DATA NSEED/567801/
C      I=ALPHA
C      X=0.0
C      DO 2 J=1,I
C      X=X-ALOG(RAN(NSEED))/SCALE
2  CONTINUE
C      VAL=X+POS
C      RETURN
C      END

```


C. COMBINATION METHOD

With similar argument as chi-square variate of non-integral degrees of freedom, a gamma variate of real α greater than one can be generated by the combination method. Observe that the combination method applies to chi-square because the expected value of the variate is equal to the degrees of freedom. This method applies to gamma because the expected value and the variance of the gamma distribution are linear functions of α .

Let $\alpha = d_1.d_2$ where d_1 is the unit digit and d_2 is the first digit after the decimal. To obtain a gamma variate:

- 1) Generate a uniform random variate, U in $(0,1)$.
- 2) If U is less than or equal to $d_2/10$, use d_1+1 as the shape parameter to generate the required variate.
- 3) If U is greater than $d_2/10$, use d_1 as the shape parameter to generate the required variate.

PROGRAM LISTING

```

C*****
C
C      TO RETURN A RANDOM GAMMA VARIATE OF NON INTEGRAL
C      SHAPE PARAMETER GREATER THAN ONE
C
C      ALPHA = THE SHAPE PARAMETER
C
C      SCALE = THE SCALE PARAMETER
C
C      POS = THE POSITION PARAMETER
C
C      VAL = THE RETURNED GAMMA VARIATE
C*****
C
C      SUBROUTINE GAMMA(ALPHA,SCALE,POS,VAL)
C      DATA NSEED/567801/
C      IAL=ALPHA
C      IF=(ALPHA-IAL)*10
C      P=IF/10.0
C      R=RAN(NSEED)
C      IF(R.LE.P) N=IAL+1
C      IF(R.GT.P) N=IAL
C      Z=0.0
C      DO 2 I=1,N
C      Z=Z+ALOG(RAN(NSEED))/SCALE
2  CONTINUE
C      VAL=Z+POS
C      RETURN
C      END

```

PROGRAM TESTING RESULTS

METHODS	PROGRAM RESULTS				GIVEN	
	Mean		Variance		EX	VX
	N=100	N=1000	N=100	N=1000		
SUM EXPONENT.	6.792	6.148	2.539	3.794	6.166	3.472
JOHNK	2.367	2.349	0.296	0.246	2.375	0.313
COMBINATION	4.883	5.929	5.163	6.703	5.250	6.563

METHODS	INPUT PARAMETERS		
	ALPHA	SCALE	POS
SUM EXPONENT.	11	1.784	0.0
JOHNK	0.375	7.587	2.0
COMBINATION	4.2	0.8	0.0

METHODS	EXECUTION TIME IN MILLI-SECONDS	
	N=100	N=1000
SUM EXPONENT.	81	797
JOHNK	89	871
COMBINATION	55	510

N is the number of random variates generated for testing.

SUMMARY OF T TEST

METHODS	Degrees of freedom	T-values
SUM EXPONENTIAL	99	3.925
	999	0.302
JOHNNK	99	0.270
	999	1.657
COMBINATION	99	1.507
	999	1.063

<u>Degrees of freedom</u>	<u>α</u>	<u>Critical T-values</u>
60	0.05	1.671
60	0.10	1.296
120	0.05	1.658
120	0.10	1.289
∞	0.05	1.645
∞	0.10	1.282

N is the number of random variates generated for testing.

18. PARETO DISTRIBUTION

Pareto distribution deals with the distribution of income over a population. It can be stated as

$$N = Ax^{-a}$$

where N is the number of persons having income x; while A and a are parameters. a is the shape parameter or the Pareto's constant. The density function is

$$f(x) = \frac{ak^a}{x^{a+1}} \quad a>0, \underline{x} \geq k > 0$$

and cumulative function

$$F(x) = 1 - \left(\frac{k}{x}\right)^a \quad k>0, a>0, \underline{x} \geq k$$

where k is the minimum income.

The mean and the variance of x are

$$EX = \frac{ak}{a-1} \quad \text{if } a>1$$

$$VX = \frac{ak}{(a-1)^2(a-2)} \quad \text{if } a>2$$

A. INVERSION METHOD

When the minimum income, k and the shape parameter are given, inversion is easy to be done.

$$\text{Since } \left(\frac{k}{x}\right)^a = 1-F(x)$$

$$a*(\text{Log}_e k - \text{Log}_e x) = \text{Log}_e (1-F(x))$$

$$\text{Log}_e k - \text{Log}_e x = \text{Log}_e (1-F(x))/a$$

$$\text{Log}_e x = \text{Log}_e k - \text{Log}_e (1-F(x))/a$$

$$x = \text{Exp}(\text{Log}_e k - \text{Log}_e (1-F(x))/a)$$

$$x = (k/(1-F(x)))/e^a$$

$$x = k*e^a/(1-F(x))$$

If $F(x)$ is a uniform random number in $(0,1)$, then $1-F(x)$ is also a uniform random number in $(0,1)$.

Therefore to obtain a random pareto variate:

- 1) Generate U , a uniform random number in $(0,1)$.
- 2) Set $x=e^a*k/U$, x is the desired pareto variate.

When the mean, EX, and the variance, VX are given and if a is larger than 2, then a and k have to be expressed in terms of EX and VX in order to do the inversion.

$$EX = \frac{ak}{a-1} \quad \text{and} \quad VX = \frac{ak}{(a-1)^2(a-2)}$$

Substitute EX into VX,

$$VX = \frac{EX^2}{a(a-2)}$$

$$\text{implies} \quad a(a-2) = \frac{EX^2}{VX}$$

$$a^2 - 2a - \frac{EX^2}{VX} = 0$$

$$a = \frac{2 + \sqrt{4 + 4(EX^2/VX)}}{2}$$

By the assumption that $a > 2$,

$$a = \frac{2 + \sqrt{4 + 4(EX^2/VX)}}{2}$$

$$\text{From } VX = \frac{ak}{a-1}, \quad k = \frac{EX^*(a-1)}{a}$$

x can now be calculated by $x = k * e^a / (1 - F(x))$

PROGRAM LISTING

```

C*****
C
C      TO RETURN A RANDOM PARETO VARIATE BY INVERSION
C
C      EX = THE GIVEN MEAN OF THE DISTRIBUTION
C
C      VX = THE GIVEN VARIANCE OF THE DISTRIBUTION
C
C      VAL = THE RETURNED PARETO VARIATE
C*****
C
C      SUBROUTINE PARETO(EX,VX,VAL)
C      DATA NSEED/567801/
C      A=(2.+SQRT(4.+4*(EX**2/VX)))/2.
C      IF(A.LT.2.) WRITE(6,88)
C      PK=EX*(A-1.)/A
C      1 R=RAN(NSEED)
C      VAL=EXP(ALOG(PK)-ALOG(R)/A)
C      IF(VAL.LT.PK) GO TO 1
C      88 FORMATN('A IS LESS THAN 2 ')
C      RETURN
C      END

```


PROGRAM TESTING RESULTS

METHOD	PROGRAM RESULTS				GIVEN	
	Mean		Variance		Mean	Variance
	N=100	N=1000	N=100	N=1000		
NAYLOR	3.318	0.715	3.186	0.811	3.20	0.80

METHOD	EXECUTION TIME IN MILLI-SECONDS	
	N=100	N=1000
NAYLOR	114	1149

SUMMARY OF T-TEST

METHOD	Degrees of freedom	T-Values
NAYLOR	99	1.395
	999	
	999	0.060

<u>Degrees of freedom</u>	<u>α</u>	<u>Critical T-values</u>
60	0.05	1.671
60	0.10	1.296
120	0.05	1.658
120	0.10	1.289
∞	0.05	1.645
∞	0.10	1.282

N is the number of random variates generated for testing.

CONCLUSION:

Quite a few random number generators have been examined and tested. Roughly speaking, the generating methods fall into several categories.

One category which is seldom used nowadays is to relate physical methods with mathematical methods. A good example is Marsaglia's method on any discrete distribution. Some work had also been done to reduce storage spaces and to suit the capacities of the computers.

Another category is to apply common statistical facts and theorems. The using of cumulative functions as random uniform variates between zero and one is a significant contribution to the generating methods. Both versions of Von Neumann's methods on the Exponential distribution use conditional probabilities and statistical independence of events. There are interesting relationships between the distributions which make possible the approximations of Binomial and Poisson variates by Normal variates. The Poisson number of arrivals of Exponential arrival time is also used. Moreover, there is also the summation of Exponential variates to approximate the Erlang variates. These methods were not easy to come by, they all require careful studies on their mathematical inferences. Out from all these, a clever development is to treat the Pascal distribution as a discrete analogue of the Exponential distribution. Ordering statistics has been used also in several methods.

The derivation for the Beta variate is half statistics and half calculus. The mathematics involved in the Cosine and Sine method for Normal variate is all Trigonometry and Calculus. One can see that not only statisticians were interested in random number generator.

The third category is the polynomial approximations. There are Hasting's and Teichroew's.

With a final word, we can conclude that anyone who does simulation work can now make use of our study. He can also design number generators for his own distribution. The process is to examine first the direct method. If it cannot apply, then scale the cumulative function to be in the range between zero and one. If the cumulative function exists in a closed form, inversion can be used. If not, the last resort is to find the relationship between the distribution and those which are in this thesis.

CHAPTER III

Usage descriptions of DTIME and CTIME

1. SUBROUTINE DTIME

Subroutine DTIME(KODE,L,K,P,VAL) is used to generate random values from any discrete distribution discussed in Chapter II.

The 5 arguments are:

KODE = code for the distribution, can be 10,20,30,40,
50 or 60.

L = 1 or 0 to indicate the nature of vector P.

See individual input arguments sections for
details.

K = dimension of vector P.

P = vector giving the mean, variance, minimum,
maximum or other parameters as required. See
individual input arguments sections for details.

VAL = the returned random variate.

A. ANY DISCRETE DISTRIBUTION

1) SELECTION:

The CDF method is selected for DTIME because it is simple to program, its testing results are satisfactory and its execution time is much less than those of the other two methods.

2) INPUT ARGUMENTS:

KODE = 10

L = 1 when MIN is supplied

= 1 when MIN and MAX are supplied

= 0 neither MIN nor MAX are supplied

K = 2 to 50 which is the dimension of vector P in DTIME

P = a vector of values and their corresponding probabilities of occurrence. If the values are x_1, x_2, \dots and their probabilities were p_1, p_2, p_3, \dots , then $P = (x_1, p_1, x_2, p_2, \dots)$. When L is supplied as 1, MIN and MAX can be put in the order that MIN comes first. MAX cannot be supplied without supplying MIN.

3) USAGE SAMPLE:

CALL DTIME(10,0,10,P,VAL) --- Input with x_i and p_i $i=1,2,3,4$ and 5.

CALL DTIME(10,1,11,P,VAL) --- Input with x_i, p_i $i=1,2,3,4$ and 5; and MIN.

B. HYPERGEOMETRIC DISTRIBUTION

1) SELECTION:

Direct method is used for DTIME because it was the only method investigated.

2) INPUT ARGUMENTS:

KODE = 20

L = 1 when EX, VX and PROB, the probability of the occurrence of class I of the distribution are supplied. MIN and MAX are optional.

= 0 when TN, the total number of items; NC, the number of items chosen; and PROB, the probability that an item of class I will be chosen are supplied. MIN and MAX are optional.

K = 3 to 5

P = (EX,VX,PROB) for L=1

(TN,NC,PROB) for L=0

3) USAGE SAMPLES:

CALL DTIME(20,1,3,P,VAL) --- Input with EX, VX and PROB.

CALL DTIME(20,1,4,P,VAL) --- Input with EX, VX, PROB and MIN.

CALL DTIME(20,0,5,P,VAL) --- Input with TN, NC, PROB, MIN and MAX.

C. GEOMETRIC DISTRIBUTION

1) SELECTION:

Direct method is selected for DTIME because it gave better results in the chi-square test and its execution time is less than Naylor's method.

2) INPUT ARGUMENTS:

KODE = 30

L = 1 when EX and VX are supplied. MIN and MAX are optional.

= 0 when PROB, the probability of the occurrence of the event is supplied. MIN and MAX are optional.

K = 1 to 4

P = (EX,VX) for L=1
(PROB) for L=0

3) USAGE SAMPLES:

CALL DTIME(30,1,2,P,VAL) --- Input with EX and VX.

~~CALL~~ CALL DTIME(30,1,3,P,VAL) --- Input with EX, VX ^{and} MIN.

CALL DTIME(30,0,3,P,VAL) --- Input with PROB, MIN and MAX.

D. PASCAL DISTRIBUTION

1) SELECTION:

Negative binomial is incorporated with non-integer method for DTIME. Negative binomial has the best chi-square test results and non-integer enables users to supply non-integer number of successes.

2) INPUT ARGUMENTS:

KODE = 40

L = 1 when EX and VX are supplied. MIN and MAX are optional.

= 0 when the number of successes, TNS and the probability that the event will occur, PROB are supplied. MIN and MAX are optional.

K = 2 to 4

P = (EX,VX) for L=1
(TNS,PROB) for L=0

3) USAGE SAMPLES:

CALL DTIME(40,1,3,P,VAL) --- Input with EX,VX and MIN

CALL DTIME(40,0,3,P,VAL) --- Input with TNS, PROB and
MIN

CATT DTIME(40,0,4,P,VAL) --- Input with TNS,PROB, MIN
and MAX.

E. BINOMIAL DISTRIBUTION

1) SELECTION:

Both the Direct and Normal Approximation methods are incorporated into DTIME. The number of trials, NT is tested first. If NT is greater than 15, the Normal Approximation method is used to generate a binomial random variate. Otherwise, the Direct method is used.

2) INPUT ARGUMENTS:

KODE = 50

L = 1 when EX and VX are supplied. MIN and MAX are optional.

= 0 when the total number of trials NT, and the probability of occurrence of the event PROB are supplied. MIN and MAX are optional.

K = 2 to 4

P = (EX,VX) for L=1
(TN,PROB) for L=0

3) USAGE SAMPLES:

CALL DTIME(50,1,3,P,VAL) --- Input with EX, VX and MIN.

CALL DTIME(50,1,4,P,VAL) --- Input with EX, VX, MIN and MAX.

CALL DTIME(50,0,2,P,VAL) --- Input with TN and PROB.

F. POISSON DISTRIBUTION

1) SELECTION:

Exponential arrival time method is chosen for DTIME
because it gave better chi-square test results.

2) INPUT ARGUMENTS:

KODE = 60

L = 1 or 0 when λ , the arrival rate per unit time;
and the number or time units, T are supplied.

MIN and MAX are optional

K = 2 to 4

P = (λ , T) L=1 or 0

3) USAGE SAMPLES:

CALL DTIME(60,1,2,P,VAL) --- Input with λ and T.

CALL DTIME(60,0,2,P,VAL) --- Input with λ and T.

CALL DTIME(60,1,3,P,VAL) --- Input with λ , T and MIN.

USAGE SUMMARY

DISTRIBUTIONS	KODE	L	K	P(1)	P(2)	P(3)	P(4)	P(5)	...
ANY DISCRETE	10	1	2 to 50	x_1	p_1	x_2	p_2	...	MIN MAX
		0	2 to 50	x_1	p_1	x_2	p_2	...	
HYPERGEOMETRIC	20	1	2 to 5	EX	VX	PROB	MIN	MAX	
		0	2 to 5	TN	NC	RROB	MIN	MAX	
GEOMETRIC	30	1	2 to 4	EX	VX	MIN	MAX		
		0	1 to 4	PROB	MIN	MAX			
PASCAL	40	1	2 to 4	EX	VX	MIN	MAX		
		0	2 to 4	TNX	PROB	MIN	MAX		
BINORMAL	50	1	2 to 4	EX	VX	MIN	MAX		
		0	2 to 4	NT	PROB	MIN	MAX		
POISSON	60	1	2 to 4	λ	T	MIN	MAX		
		0	2 to 4	λ	T	MIN	MAX		

DTIME PROGRAM RESULTS

INPUT													N=100		OUTPUT	
KODE	L	K	P(1)	P(2)	P(3)	P(4)	P(5)	P(6)	P(7)	P(8)	P(9)	P(10)	EX	VX		
10	1	9	1.00	0.20	2.00	0.30	3.00	0.10	4.00	0.40	1.50		3.080	0.814		
10	1	10	0.20	2.00	0.30	3.00	0.10	4.00	0.40	1.50	3.40		2.280	0.202		
10	0	10	1.00	0.10	2.00	0.20	3.00	0.15	4.00	0.30	5.00	0.25	3.560	1.430		
20	1	3	4.00	0.60	0.45								3.550	0.428		
20	1	4	4.00	0.60	0.45	2.30							4.110	0.098		
20	1	5	4.00	0.60	0.45	1.00	3.50						2.920	0.074		
20	0	3	10.00	4.00	0.45								1.840	0.674		
20	0	4	10.00	4.00	0.45	0.96							1.910	0.542		
20	0	5	10.00	4.00	0.45	0.96	2.80						1.530	0.249		
30	1	2	3.20	12.40									4.510	12.850		
30	1	3	3.20	12.40	2.10								6.300	9.430		
30	1	4	3.20	12.40	2.10	5.20							3.750	0.527		
30	0	1	0.21										5.250	17.287		
30	0	2	0.21	2.10									7.050	12.868		
30	0	3	0.21	2.10	4.30								3.500	0.250		

DTIME PROGRAM RESULTS

INPUT N=100							OUTPUT	
KODE	L	K	P(1)	P(2)	P(3)	P(4)	EX	VX
40	1	2	12.94	35.40			12.834	3.120
40	1	3	12.94	35.40	10.00		13.091	2.647
40	1	4	12.94	35.40	10.00	12.40	11.481	0.381
40	0	2	7.00	0.80			1.972	0.276
40	0	3	7.00	0.80	2.10		2.421	0.106
40	0	4	7.00	0.80	2.10	2.76	2.351	0.031
40	0	2	4.20	0.24			12.943	4.466
40	0	3	4.20	0.24	10.50	17.00	13.304	2.993
40	0	4	4.20	0.24	10.50	17.00	13.315	2.606
50	1	2	3.20	0.80			3.070	0.645
50	1	3	3.20	0.80	1.20		3.110	0.578
50	1	4	3.20	0.80	1.20	3.50	2.690	0.214
50	0	2	12.94	0.24			4.030	1.049
50	0	3	12.94	0.24	2.10		2.780	2.852
50	0	4	12.94	0.24	2.10	4.00	3.430	0.245

DTIME PROGRAM RESULTS

INPUT N=100							OUTPUT	
KODE	L	K	P(1)	P(2)	P(3)	P(4)	EX	VX
50	0	2	20.0	0.33			6.597	3.995
50	0	3	20.0	0.33	3.80		6.889	2.736
50	0	4	20.0	0.33	3.80	7.20	5.656	0.807
60	0	2	3.2	1.00			3.310	2.614
60	0	3	3.2	1.00	1.20		3.100	0.590
60	0	4	3.2	1.00	1.20	4.00	3.860	1.820

2. SUBROUTINE CTIME

Subroutine CTIME(KODE,L,K,P,VAL) is used to generate random values from any continuous distribution discussed in Chapter II.

The 5 arguments are:

KODE = code for the distribution, can be 10,20,30,40,
...,120.

L = 1 or 0 to indicate the nature of vector P.

See individual input arguments sections for
detail

K = dimension of vector P.

P = vector giving the mean, variance, minimum,
maximum, or other parameters as required. See
individual input arguments sections for details.

VAL = the returned random variate.

A. EXPONENTIAL DISTRIBUTION

1) SELECTION:

Inversion method is used in CTIME. It gave slightly better results in the chi-square test than Marsaglia method, and definitely gave better results than Von Neumann's two methods.

2) INPUT ARGUMENTS:

KODE = 10

L = 1 when EX is supplied. If right skew is required, SKEW, MIN and MAX are optional inputs. If left skew is required, supply SKEW=0.0. MIN and MAX must also be supplied.

= 0 when α is supplied. If right skew is required, SKEW, MIN and MAX are optional inputs. If left skew is required, supply SKEW=0.0, MIN and MAX must also be supplied.

K = 1 to 4

P = (EX) for L=1 and right skew
 = (EX,SKEW,MIN,MAX) for L=1 and left skew
 = (α) for L=0 and right skew
 = (α ,SKEW,MIN,MAX) for L=1 and left skew

3) USAGE SAMPLES:

CALL CTIME(10,1,1,P,VAL) --- Input with EX.

CALL CTIME(10,0,4,P,VAL) --- Input with α , SKEW, MIN and MAX.

B. NORMAL DISTRIBUTION

1) SELECTION:

Teichrow's method is selected for CTIME because it gave the best chi-square test results.

2) INPUT ARGUMENTS:

KODE = 20

L = 1 or 0 when EX and VX are supplied. MIN and MAX are optional.

K = 2 to 4

P = (EX,VX)

3) USAGE SAMPLES:

CALL CTIME(20,1,2,P,VAL) --- Input with EX and VX.

CALL CTIME(20,0,2,P,VAL) --- Input with EX and VX.

CALL CTIME(20,0,4,P,VAL) --- Input with EX,VX,MIN and MAX.

C. TRIANGULAR DISTRIBUTION

1) SELECTION:

Inversion method is used in CTIME because it takes less execution time and has given better chi-square test results than rejection method.

2) INPUT ARGUMENTS:

KODE = 30

L = 1 or 0 when the three points x_1, x_2 and x_3 of the distribution are given. MIN and MAX are optional.

K = 3 to 5

P = (x_1, x_2, x_3)

3) USAGE SAMPLES:

CALL CTIME(30,0,3,P,VAL) --- Input with x_1, x_2 and x_3 .

CALL CTIME(30,1,3,P,VAL) --- Input with x_1, x_2 and x_3 .

CALL CTIME(30,1,5,P,VAL) --- Input with x_1, x_2, x_3 MIN and MAX.

D. TRAPEZOIDAL DISTRIBUTION

1) SELECTION:

Inversion method is used in CTIME because it takes less execution time and has given better chi-square test results than rejection method.

2) INPUT ARGUMENTS:

KODE = 40

L = 1 or 0 when the four points x_1, x_2, x_3 and x_4 of the distribution are given. MIN and MAX are optional.

K = 4 to 6

P = (x_1, x_2, x_3, x_4)

3) USAGE SAMPLES:

CALL CTIME(40,0,4,P,VAL) --- Input with x_1, x_2, x_3 and x_4 .

CALL CTIME(40,1,4,P,VAL) --- Input with x_1, x_2, x_3 and x_4 .

CALL CTIME(40,1,5,P,VAL) --- Input with x_1, x_2, x_3, x_4 and MIN.

E. LOGNORMAL DISTRIBUTION

1) SELECTION:

Naylor method is used for CTIME because it was the only method investigated.

2) INPUT ARGUMENTS:

KODE = 50

L = 1 when the mean, EY and the variance, VY of the function $y = \text{Log}(x)$ are supplied. MIN and MAX are optional.

0 when the EX and VX of x are supplied. MIN and MAX are optional.

K = 2 to 4

P = (EY,VY) for L=1
(EX,VX) for L=0

3) USAGE SAMPLES:

CALL CTIME(50,1,2,P,VAL) --- Input with EY and VY

CALL CTIME(50,1,3,P,VAL) --- Input with EY, VY and MIN.

CALL CTIME(50,0,4,P,VAL) --- Input with EX, VX, MIN and MAX.

F. CHI-SQUARE DISTRIBUTION

1) SELECTION:

Combination method is chosen for CTIME for input with non-integer degrees of freedom.

2) INPUT ARGUMENTS:

KODE = 60

L = 1 or 0 when the degrees of freedom N, or EX is supplied. MIN and MAX are optional.

K = 1 to 3

P = (N) L=1 or 0

(EX) L=1 or 0

3) USAGE SAMPLES:

CALL CTIME(60,1,1,P,VAL) --- Input with N or EX.

CALL CTIME(60,0,1,P,VAL) --- Input with EX or N.

CALL CTIME(60,0,3,P,VAL) --- Input with N or EX, MIN and MAX.

G. BETA DISTRIBUTION

1) SELECTION:

Chi-square method is selected for CTIME because it gave better results in T test than ordering method.

2) INPUT ARGUMENTS:

KODE = 70

L = 1 when EX and VX are supplied. MIN and MAX are optional.

0 when the degrees of freedom, M and N are supplied
MIN and MAX are optional.

K = 2 to 4

P = (EX,VX) for L=1
(M,N) for L=0

3) USAGE SAMPLES:

CALL CTIME(70,1,3,P,VAL) --- Input with EX,VX and MIN.

CALL CTIME(70,0,2,P,VAL) --- Input with M and N.

CALL CTIME(70,0,4,P,VAL) --- Input with M,N,MIN and MAX.

H. F DISTRIBUTION

1) SELECTION:

To make it consistent with beta distribution, chi-square method is chosen for CTIME. Both chi-square and ordering methods gave almost equal results in T test.

2) INPUT ARGUMENTS:

KODE = 80

L = 1 when EX and VX are supplied. MIN and MAX are optional.

0 when the degrees of freedom, M and N are supplied. MIN and MAX are optional.

K = 2 to 4

P = (EX,VX) for L=1
(M,N) for L=0

3) USAGE SAMPLES:

CALL CTIME(80,1,3,P,VAL) --- Input with EX,VX and MIN.

CALL CTIME(80,0,2,P,VAL) --- Input with M and N.

Call CTIME(80,0,4,P,VAL) --- Input with M,N,MIN and MAX.

I. T DISTRIBUTION

1) SELECTION:

Chi-square method is chosen for CTIME because it gave better results in T test than Cacoullos method.

2) INPUT ARGUMENTS:

KODE = 90

L = 1 when EX and VX are supplied. MIN and MAX are optional.

0 when the degrees of freedom, N is supplied.

MIN and MAX are optional.

K = 1 to 4

P = (EX,VX) for L=1

(N) for L=0

3) USAGE SAMPLES:

CALL CTIME(90,1,3,P,VAL) --- Input with EX,VX and MIN.

CALL CTIME(90,0,1,P,VAL) --- Input with N.

CALL CTIME(90,0,3,P,VAL) --- Input with N,MIN and MAX.

J. Z DISTRIBUTION

1) SELECTION:

Chi-square method is chosen for CTIME because it is the only method investigated in Chapter II.

2) INPUT ARGUMENTS:

KODE = 100

L = 1 when EX and VX of e are supplied. MIN and MAX are optional. x is the z-variate.

0 when the degrees of freedom, M and N of x are supplied. MIN and MAX are optional and x is the z-variate.

K = 2 to 4

P = (EX,VX) for L=1
(M,N) for L=0

3) USAGE SAMPLES:

CALL CTIME(100,1,2,P,VAL) --- Input with EX and VX.

CALL CTIME(100,1,4,P,VAL) --- Input with EX,VX,MIN and MAX.

CALL CTIME(100,0,3,P,VAL) --- Input with M,N and MIN.

2

4

9

L. PARETO DISTRIBUTION

1) SELECTION:

Inversion method is used in CTIME because it is the only method investigated in Chapter II.

2) INPUT ARGUMENTS:

KODE = 120

L = 1 when EX and VX are supplied. MIN and MAX are optional.

= 0 when the shape parameter a and the minimum income k are supplied. MIN and MAX are optional.

K = 2 to 4

P = (EX,VX) for L=1
(a,k) for L=0

3) USAGE SUMPLES:

CALL CTIME(120,1,2,P,VAL) --- Input with EX and VX.

CALL CTIME(120,1,4,P,VAL) --- Input with EX, VX, MIN and MAX.

CALL CTIME(120,0,3,P,VAL) --- Input with a, k and MIN.

USAGE SUMMARY

DISTRIBUTIONS	KODE	L	K	P(1)	P(2)	P(3)	P(4)	P(5)	P(6)
EXPONENTIAL	10	1	1 to 4	EX	skew	MIN	MAX		
		0	1 to 4	α	skew	MIN	MAX		
NORMAL	20	1	2 to 4	EX	VX	MIN	MAX		
		0	2 to 4	EX	VX	MIN	MAX		
TRIANGULAR	30	1	3 to 5	x_1	x_2	x_3	MIN	MAX	
		0	3 to 5	x_1	x_2	x_3	MIN	MAX	
TRAPEZODIAL	40	1	4 to 6	x_1	x_2	x_3	x_4	MIN	MAX
		0	4 to 6	x_1	x_2	x_3	x_4	MIN	MAX
LOGNORMAL	50	1	2 to 4	EY	VY	MIN	MAX		
		0	2 to 4	EX	VX	MIN	MAX		
CHI SQUARE	60	1	1 to 3	EX	MIN	MAX			
		0	1 to 3	N	MIN	MAX			

USAGE SUMMARY

DISTRIBUTION	KODE	L	K	P(1)	P(2)	P(3)	P(4)	P(5)	P(6)
BETA	70	1	2 to 4	EX	VX	MIN	MAX		
		0	2 to 4	M	N	MIN	MAX		
F	80	1	2 to 4	EX	VX	MIN	MAX		
		0	2 to 4	M	N	MIN	MAX		
T	90	1	2 to 4	EX	VX	MIN	MAX		
		0	1 to 3	N	MIN	MAX			
Z	100	1	2 to 4	$EX(e^{2x})$	$VX(e^{2x})$	MIN	MAX		
		0	2 to 4	M	N	MIN	MAX		
GAMMA & ERLANG	110	1	3 to 5	EX	VX	γ	MIN	MAX	
		0	3 to 5	α	β	γ	MIN	MAX	
PARETO	120	1	2 to 4	EX	VX	MIN	MAX		
		0	2 to 4	a	k	MIN	MAX		

CTIME PROGRAM RESULTS

INPUT N=100								OUTPUT	
KODE	L	K	P(1)	P(2)	P(3)	P(4)	P(5)	EX	VX
10	1	1	6.0					5.113	20.445
10	1	3	6.0	1.0	1.2			5.261	14.198
10	1	4	6.0	0.0	-0.35	10.0		6.853	6.299
10	1	4	6.0	1.0	1.2	12.4		4.828	8.908
10	1	1	76.9					65.536	3358.492
10	1	3	76.9	1.0	20.0			68.492	1838.770
10	1	4	76.9	1.0	20.0	99.0		50.775	479.238
10	1	4	76.9	0.0	10.5	85.4		78.096	37.398
10	0	1	0.85					1.003	00.786
10	0	3	0.85	1.0	1.2			1.150	0.018
10	0	4	0.85	1.0	1.2	2.76		1.150	0.018
10	0	1	0.012					71.019	3943.934
10	0	3	0.012	1.0	10.5			72.019	3054.223
10	0	4	0.012	0.0	10.5	85.4		83.579	1.734

CTIME PROGRAM RESULTS

INPUT N=100									OUTPUT	
KODE	L	K	P(1)	P(2)	P(3)	P(4)	P(5)	P(6)	EX	VX
20	1	2	0.0	1.0					-0.112	1.159
20	1	3	0.0	1.0	-2.55				-0.087	1.083
20	1	4	0.0	1.0	-2.55	2.76			-0.117	0.990
30	0	3	0.0	4.3					1.897	0.831
30	0	4	0.0	4.3	0.88				2.212	0.535
30	0	5	0.0	4.3	0.88	4.2			2.212	0.535
40	0	4	0.0	1.0	4.3	6.5			2.4	2.938
40	0	5	0.0	1.0	4.3	6.5	0.6		2.746	2.616
40	0	6	0.0	1.0	4.3	6.5	0.6	6.0	2.736	2.558
50	1	2	0.0	1.0					1.586	5.207
50	1	3	0.0	1.0	0.6				2.396	7.164
50	1	4	0.0	1.0	0.6	6.5			1.924	1.624
50	0	2	3.2	0.8					3.122	0.879
50	0	3	3.2	0.8	2.10				3.307	0.753
50	0	4	3.2	0.8	2.10	4.0			3.307	0.753

CTIME PROGRAM RESULTS

INPUT								N=100		OUTPUT	
KODE	L	K	P(1)	P(2)	P(3)	P(4)	p(5)	EX	VX		
60	1	1	4.2					4.098	0.955		
60	1	2	4.2	2.88				4.252	0.801		
60	1	3	4.2	2.88	4.2			3.582	0.153		
70	1	2	0.835	0.012				0.855	0.012		
70	1	3	0.835	0.012	0.6			0.856	0.007		
70	1	4	0.835	0.012	0.6	0.88		0.795	0.004		
70	0	2	12.54	3.0				0.835	0.012		
70	0	3	12.54	3.0	0.6			0.842	0.010		
70	0	4	12.54	3.0	0.6	0.88		0.765	0.006		
80	1	2	1.5	3.38				1.479	3.333		
80	1	3	1.5	3.38	0.6			1.787	3.262		
80	1	4	1.5	3.38	0.6	2.76		1.303	0.247		
80	0	2	3.2	12.4				1.178	1.220		
80	0	3	3.2	12.4	0.6			1.602	1.158		
80	0	4	3.2	12.4	0.6	6.5		1.534	0.916		

CTIME PROGRAM RESULTS

INPUT N=100								OUTPUT	
KODE	L	K	P(1)	P(2)	P(3)	P(4)	P(5)	EX	VX
90	1	2	0.0	4.466				-0.177	4.315
90	1	3	0.0	4.466	-2.55			0.177	4.315
90	1	4	0.0	4.466	2.55	4.0		-0.169	1.126
90	0	1	4.2					0.096	1.490
90	0	2	4.2	0.012				0.945	0.929
90	0	3	4.2	0.012	4.3			0.890	0.697
100	1	2	1.5	2.38				-0.007	0.131
100	1	3	1.5	2.38	0.0			0.366	0.077
100	1	4	1.5	2.38	0.0	0.7		0.291	0.038
100	0	2	7.0	4.466				-0.038	0.234
100	0	3	7.0	4.466	-0.35			0.231	0.214
100	0	4	7.0	4.466	-0.35	0.098		0.098	0.107
110	1	3	3.2	0.8	0.0			3.863	0.472
110	1	4	3.2	0.8	0.0	2.76		3.935	0.485

CTIME PROGRAM RESULTS

INPUT N=100								OUTPUT	
KODE	L	K	P(1)	P(2)	P(3)	P(4)	P(5)	EX	VX
110	1	5	3.2	0.8	0.0	2.76	5.1	3.830	0.326
110	0	3	0.83	0.8	0.0			1.129	1.033
110	0	4	0.83	0.8	0.0	0.7		1.862	1.553
110	0	5	0.83	0.8	0.0	0.7	0.6	1.738	1.008
110	0	3	4.2	0.8	0.0			4.883	5.929
110	0	4	4.2	0.8	0.0	1.0		4.883	5.929
110	0	5	4.2	0.8	0.0	1.0	5.0	3.189	0.973
120	1	2	3.2	0.4				3.103	0.175
120	1	3	3.2	0.4	3.1			3.628	0.323
120	1	4	3.2	0.4	3.1	4.0		3.409	0.059
120	0	2	7.0	4.47				5.076	0.353
120	0	3	7.0	4.47	4.2			5.076	0.353
120	0	4	7.0	4.47	4.2	5.9		4.927	0.128

APPENDIX A

```

C*****DTIM 1
C      TO GENERATE A RANDOM NUMBER OF ONE OF THE DISTRIBUTION BELOW:  *DTIM 2
C      DISTRIBUTIONS:          CODE:  *DTIM 3
C      ANY DISCRETE          10  *DTIM 4
C      HYPERGEOMETRIC      20  *DTIM 5
C      GEOMETRIC          30  *DTIM 6
C      PASCAL            40  *DTIM 7
C      BINOMIAL          50  *DTIM 8
C      POISSON           60  *DTIM 9
C      SUBROUTINE REQUIRED: NORM, ERROR AND RAN(NSEED)  *DTIM 10
C      KODE = THE CODE OF THE DISTRIBUTION  *DTIM 11
C      L = 0 OR 1 TO SPECIFY THE NATURE OF VECTOR P  *DTIM 12
C      K = DIMENSION OF P  *DTIM 13
C      P = VECTOR WITH VALUES OF MEAN, VARIANCE, OR OTHER PARAMETERERS OF  *DTIM 14
C          THE DISTRIBUTION  *DTIM 15
C      VAL = THE RETURNED VARIATE  *DTIM 16
C*****DTIM 17
C      SUBROUTINE DTIME(KODE,L,K,P,VAL)  *DTIM 18
C      DIMENSION P(50)  *DTIM 19

```

```

      SUBROUTINE DTIME(KODE,L,K,P,VAL)
      DIMENSION P(50)

```

DATA NSEED/567801/

IF(KODE.EQ.10) GO TO 10

IF(KODE.EQ.20) GO TO 20

IF(KODE.EQ.30) GO TO 30

IF(KODE.EQ.40) GO TO 40

IF(KODE.EQ.50) GO TO 50

IF(KODE.EQ.60) GO TO 60

ANY DISCRETE

10 M=0

IF(L.EQ.0) IK=K

TET=(-1.0)**K

IF(L.EQ.1.AND.TET.EQ.1.0) IK=K-2

IF(L.EQ.1.AND.TET.EQ.-1.0) IK=K-1

11 K=RAN(NSEED)

LL=K

M=M+1

IF(M.EQ.100) CALL ERROR(10)

COMP=0.0

DO 12 I=2,IK,2

J=I-1

COMP=P(I)+COMP

IF(COMP.GT.1.0) CALL ERROR(11)

IF(R.LT.COMP) GO TO 13

12 CONTINUE

13 X=P(J)

IF(LL.EQ.IK) GO TO 15

IF(TET.EQ.-1.0) GO TO 14

IF(X.GT.P(LL)) GO TO 11

LL=K-1

14 IF(X.LT.P(LL)) GO TO 11

15 VAL=X

GO TO 1

HYPERGEOMETRIC

DTIM 36

DTIM 37

DTIM 38

DTIM 39

DTIM 40

DTIM 41

DTIM 42

DTIM 43

DTIM 44

DTIM 45

DTIM 46

DTIM 47

DTIM 48

DTIM 49

DTIM 50

DTIM 51

DTIM 52

DTIM 53

DTIM 54

DTIM 55

DTIM 56

DTIM 57

DTIM 58

DTIM 59

DTIM 60

DTIM 61

DTIM 62

DTIM 63

DTIM 64

DTIM 65

DTIM 66

DTIM 67

DTIM 68

DTIM 69

DTIM 70

DTIM 71

```

C
20 M=0
   IF(L.EQ.0) GO TO 21
   NC=P(1)/P(3)
   IF(NC.LT.1) CALL ERROR(21)
   NN=(P(2)-P(1)*(1.-P(3))*NC)/(P(2)-P(1)*(1.-P(3)))
   TN=NN
   IF(TN.LT.1) CALL ERROR(22)
   IF(NN.LT.NC) CALL ERROR(23)
   GO TO 22
21 NC=P(2)
   TN=P(1)
22 PX=P(3)
   PTN=TN
   X=0.0
   M=M+1
   IF(M.EQ.100) CALL ERROR(20)
   DO 26 I=1,NC
   R=РАН(NSEED)
   IF(R-PX) 23,23,24
23 S=1.0
   X=X+1.0
   GO TO 25
24 S=0.0
25 PX=(TN*PX-S)/(TN-1.0)
   TN=TN-1.0
26 CONTINUE
   IF(K.EQ.3) GO TO 28
   IF(K.EQ.4) GO TO 27
   TN=PTN
   IF(X.GT.P(5)) GO TO 22
27 TN=PTN
   IF(X.LT.P(4)) GO TO 22
28 VAL=X
   GO TO 1

```

```

DTIM 72
DTIM 73
DTIM 74
DTIM 75
DTIM 76
DTIM 77
DTIM 78
DTIM 79
DTIM 80
DTIM 81
DTIM 82
DTIM 83
DTIM 84
DTIM 85
DTIM 86
DTIM 87
DTIM 88
DTIM 89
DTIM 90
DTIM 91
DTIM 92
DTIM 93
DTIM 94
DTIM 95
DTIM 96
DTIM 97
DTIM 98
DTIM 99
DTIM100
DTIM101
DTIM102
DTIM103
DTIM104
DTIM105
DTIM106
DTIM107

```

GEOMETRIC

```
C
C
30 M=0
   IF(L.EQ.0) GO TO 35
   PROB=P(1)/P(2)
31 M=M+1
   IF(M.EQ.100) CALL ERROR(30)
   X=0.0
32 R=RAN(NSEED)
   IF(R.LT.PROB) GO TO 33
   X=X+1.
   GO TO 32
33 X=X+1.0
   IF(K.EQ.2) GO TO 39
   IF(K.EQ.3) GO TO 34
   IF(X.GT.P(4)) GO TO 31
34 IF(X.LT.P(3)) GO TO 31
   GO TO 39
35 M=M+1
   IF(M.EQ.100) CALL ERROR(30)
   X=0.0
36 R=RAN(NSEED)
   IF(R.LT.P(1)) GO TO 37
   X=X+1.0
   GO TO 36
37 X=X+1.0
   IF(K.EQ.1) GO TO 39
   IF(K.EQ.2) GO TO 38
   IF(X.GT.P(3)) GO TO 35
38 IF(X.LT.P(2)) GO TO 35
39 VAL=X
   IF(X.EQ.0.0) GO TO 35
   GO TO 1
```

PASCAL

```
DTIM108
DTIM109
DTIM110
DTIM111
DTIM112
DTIM113
DTIM114
DTIM115
DTIM116
DTIM117
DTIM118
DTIM119
DTIM120
DTIM121
DTIM122
DTIM123
DTIM124
DTIM125
DTIM126
DTIM127
DTIM128
DTIM129
DTIM130
DTIM131
DTIM132
DTIM133
DTIM134
DTIM135
DTIM136
DTIM137
DTIM138
DTIM139
DTIM140
DTIM141
DTIM142
DTIM143
```



```
40 M=0
C   TO TEST IF THE NUMBER OF SUCCESSES IS AN INTEGER OR NOT
    IF(L.EQ.0) GO TO 41
    PROB=P(1)/P(2)
    IF(PROB.GT.1.0) CALL ERROR(41)
    TNS=P(1)*PROB/(1.0-PROB)
    IF(TNS.LE.1.0) CALL ERROR(42)
    GO TO 42
41 PROB=P(2)
    TNS=P(1)
    IF(TNS.LE.1.0) CALL ERROR(42)
42 T=1.0
43 T=T+1.0
    M=M+1
    IF(M.EQ.100) CALL ERROR(40)
    IF(TNS.EQ.T) GO TO 44
    IF(TNS.GT.T) GO TO 47
    GO TO 43
44 TN=0.0
C   NUMBER OF SUCCESSES IS AN INTEGER
    TM=0.0
45 R=RAN(NSEED)
    IF(R.LE.PROB) TM=TM+1.0
    TN=TN+1.0
    IF(TM.EQ.TNS) GO TO 46
    GO TO 45
46 X=TN-TM
    GO TO 444
C   NUMBER OF SUCCESSES IS NOT AN INTEGER
47 NT=TNS
    TN=NT
    NK=(TNS-NT)*10
    SUM=0.0
    SN=0.0
    SI=0.0
    SL=TN+1.0
```

DTIM144
DTIM145
DTIM146
DTIM147
DTIM148
DTIM149
DTIM150
DTIM151
DTIM152
DTIM153
DTIM154
DTIM155
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DTIM157
DTIM158
DTIM159
DTIM160
DTIM161
DTIM162
DTIM163
DTIM164
DTIM165
DTIM166
DTIM167
DTIM168
DTIM169
DTIM170
DTIM171
DTIM172
DTIM173
DTIM174
DTIM175
DTIM176
DTIM177
DTIM178
DTIM179

```
DO 440 I=1,NK
48 R=RAN(NSEED)
   IF(R.LE.PROB) SI=SI+1.0
   SN=SN+1.0
   IF(SI.EQ.SL) GO TO 49
   GO TO 48
49 SUM=SUM+(SN-SI)
   SN=0.0
   SI=0.0
440 CONTINUE
   KK=10-NK
   DO 443 I=1,KK
441 R=RAN(NSEED)
   IF(R.LE.PROB) SI=SI+1.0
   SN=SN+1.0
   IF(SI.EQ.TN) GO TO 442
   GO TO 441
442 SUM=SUM+(SN-SI)
   SN=0.0
   SI=0.0
443 CONTINUE
   X=SUM/(NK+KK)
444 IF(K.EQ.2) GO TO 446
   IF(K.EQ.3) GO TO 445
   IF(X.GT.P(4)) GO TO 42
445 IF(X.LT.P(3)) GO TO 42
446 VAL=X
   GO TO 1
```

C
C
C
BINOMIAL

```
50 M=0
   IF(L.EQ.C) GO TO 51
   PROB=1.0-P(2)/P(1)
   IF(PROB.GT.1.0) CALL ERROR(51)
   NT=P(1)/PROB
```

DTIM180
DTIM181
DTIM182
DTIM183
DTIM184
DTIM185
DTIM186
DTIM187
DTIM188
DTIM189
DTIM190
DTIM191
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DTIM193
DTIM194
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DTIM196
DTIM197
DTIM198
DTIM199
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DTIM202
DTIM203
DTIM204
DTIM205
DTIM206
DTIM207
DTIM208
DTIM209
DTIM210
DTIM211
DTIM212
DTIM213
DTIM214
DTIM215

```
GO TO 52
51 NT=P(1)
  IF(NT.LT.1.0) CALL ERROR(52)
  PROB=P(2)
52 M=M+1
  IF(M.EQ.100) CALL ERROR(50)
  IF(NT.GT.15) GO TO 55
  X=0.0
  DO 54 I=1,NT
    R=RAN(NSEED)
    IF(R-PROB) 53,53,54
53 X=X+1.0
54 CONTINUE
  GO TO 56
55 A=PROB*NT
  B=PROB*NT*(1.-PROB)
  CALL NORM(A,B,X)
56 IF(K.EQ.2) GO TO 58
  IF(K.EQ.3) GO TO 57
  IF(X.GT.P(4)) GO TO 52
57 IF(X.LT.P(3)) GO TO 52
58 VAL=X
  GO TO 1
```

C
C
C
POISSON

```
60 M=0
  N=P(2)
61 M=M+1
  IF(M.EQ.100) CALL ERROR(60)
  X=0.0
  DO 64 I=1,N
    TEST=EXP(-P(1))
    TR=1.0
62 TR=TR*RAN(NSEED)
  IF(TR-TEST) 64,63,63
```

```
DTIM216
DTIM217
DTIM218
DTIM219
DTIM220
DTIM221
DTIM222
DTIM223
DTIM224
DTIM225
DTIM226
DTIM227
DTIM228
DTIM229
DTIM230
DTIM231
DTIM232
DTIM233
DTIM234
DTIM235
DTIM236
DTIM237
DTIM238
DTIM239
DTIM240
DTIM241
DTIM242
DTIM243
DTIM244
DTIM245
DTIM246
DTIM247
DTIM248
DTIM249
DTIM250
DTIM251
```

```
63 X=X+1.0
   GO TO 62
64 CONTINUE
   IF(K.EQ.2) GO TO 66
   IF(K.EQ.3) GO TO 65
   IF(X.GT.P(4)) GO TO 61
65 IF(X.LT.P(3)) GO TO 61
66 VAL=X
   GO TO 1
1 RETURN
  END
```

PAGE

DTIM252
DTIM253
DTIM254
DTIM255
DTIM256
DTIM257
DTIM258
DTIM259
DTIM260
DTIM261
DTIM262

```

C*****ERRO 1
C                                     *ERRO 2
C      TO GIVE THE K-TH ERROR MESSAGE      *ERRO 3
C                                     *ERRO 4
C      K = THE NUMBER FOR THE ERROR      *ERRO 5
C                                     *ERRO 6
C*****ERRO 7
C                                     ERRO 8
C      SUBROUTINE ERROR(K)                ERRO 9
C      IF(K.EQ.10) WRITE(6,10)            ERRO 10
C      IF(K.EQ.11) WRITE(6,11)            ERRO 11
C      IF(K.EQ.20) WRITE(6,20)            ERRO 12
C      IF(K.EQ.21) WRITE(6,21)            ERRO 13
C      IF(K.EQ.22) WRITE(6,22)            ERRO 14
C      IF(K.EQ.23) WRITE(6,23)            ERRO 15
C      IF(K.EQ.30) WRITE(6,30)            ERRO 16
C      IF(K.EQ.40) WRITE(6,40)            ERRO 17
C      IF(K.EQ.41) WRITE(6,41)            ERRO 18
C      IF(K.EQ.42) WRITE(6,42)            ERRO 19
C      IF(K.EQ.50) WRITE(6,50)            ERRO 20
C      IF(K.EQ.51) WRITE(6,51)            ERRO 21
C      IF(K.EQ.52) WRITE(6,52)            ERRO 22
C      IF(K.EQ.60) WRITE(6,60)            ERRO 23
10 FORMAT(' ERROR IN ANY DISCRETE DISTRIBUTION ') ERRO 24
11 FORMAT(' SUM OF ANY DISCRETE PROBABILITY IS GREATER THAN ONE ') ERRO 25
20 FORMAT(' ERROR IN HYPERGEOMETRIC DISTRIBUTION ') ERRO 26
21 FORMAT(' NUMBER OF ITEMS CHOSEN IN HYPERGEOMETRIC DIST IS LESS ERRO 27
+THAN ONE ') ERRO 28
22 FORMAT(' TOTAL NUMBER OF ITEMS IN HYPERGEOMETRIC DISTRIBUTION IS ERRO 29
+ LESS THAN ONE ') ERRO 30
23 FORMAT(' TOTAL NUMBER OF ITEMS IN HYPERGEOMETRIC DISTRIBUTION IS ERRO 31
+ LESS THAN THE NUMBER OF ITEMS CHOSEN ') ERRO 32
30 FORMAT(' ERROR IN GEOMETRIC DISTRIBUTION ') ERRO 33
40 FORMAT(' ERROR IN PASCAL DISTRIBUTION ') ERRO 34
41 FORMAT(' THE PROBABILITY P OF PASCAL DISTRIBUTION IS GREATER THANERRO 35

```

	PAGE
+ ONE ')	ERRO 36
42 FORMAT(' THE TOTAL NUMBER OF SUCCESSES IN PASCAL DISTRIBUTION IS	ERRO 37
+ LESS THAN ONE ')	ERRO 38
50 FORMAT(' ERROR IN BINOMIAL DISTRIBUTION ')	ERRO 39
51 FORMAT(' THE PROBABILITY OF OCCURANCE OF BINOMIAL DISTRIBUTION IS	ERRO 40
+ GREATER THAN ONE ')	ERRO 41
52 FORMAT(' THE NUMBER OF TRIALS OF BINOMIAL DISTRIBUTION IS LESS	ERRO 42
+ THAN ONE ')	ERRO 43
60 FORMAT(' ERROR IN POISSON DISTRIBUTION')	ERRO 44
RETURN	ERRO 45
END	ERRO 46

```
C*****NORM 1
C                                          *NORM 2
C      TO RETURN A RANDOM NORMAL VARIATE      *NORM 3
C                                          *NORM 4
C      EX = THE MEAN OF THE DISTRIBUTION      *NORM 5
C                                          *NORM 6
C      VX = THE VARIANCE OF THE DISTRIBUTION  *NORM 7
C                                          *NORM 8
C      VAL = THE RETURNED NORMAL VARIATE      *NORM 9
C                                          *NORM 10
C*****NORM 11
C                                          NORM 12
C      SUBROUTINE NORM(EX,VX,VAL)              NORM 13
C      DATA NSEED/221133/                     NORM 14
C      SUM=0.0                                  NORM 15
C      DO 25 I=1,12                             NORM 16
C      SUM=SUM+RAN(NSEED)                       NORM 17
C25 CONTINUE                                   NORM 18
C      VAL=(SUM-6.0)*SQRT(VX)+EX                NORM 19
C      RETURN                                   NORM 20
C      END                                       NORM 21
```

```
C*****RAN
C                                             *RAN
C   TO RETURN A RANDOM UNIFORM VARIATE IN (0,1) *RAN
C                                             *RAN
C   NSEED = A 5 DIGITS ODD INTEGER SEED *RAN
C                                             *RAN
C*****RAN
C                                             RAN
C   FUNCTION RAN(NSEED) *RAN
C   NSEED=IABS(NSEED*655393) *RAN
C   RAN=FLOAT(MOD(NSEED,33554432))/FLOAT(33554432) *RAN
C   RETURN *RAN
C   END *RAN
```


APPENDIX B

```

C*****CTIM 1
C      *CTIM 2
C      TO GENERATE A RANDOM NUMBER OF ONE OF THE DISTRIBUTION BELOW: *CTIM 3
C      DISTRIBUTIONS: CODE: *CTIM 4
C      EXPONENTIAL 10 *CTIM 5
C      NORMAL 20 *CTIM 6
C      TRIANGULAR 30 *CTIM 7
C      TRAPEZODIAL 40 *CTIM 8
C      LOGNORMAL 50 *CTIM 9
C      CHI-SQUARE 60 *CTIM 10
C      BETA 70 *CTIM 11
C      F 80 *CTIM 12
C      T 90 *CTIM 13
C      Z 100 *CTIM 14
C      GAMMA 110 *CTIM 15
C      PARETO 120 *CTIM 16
C      SUBROUTINE REQUIRED: NORM, ERROR AND RAN(NSEED) *CTIM 17
C      KODE = THE CODE OF THE DISTRIBUTION *CTIM 18
C      L = 0 OR 1 TO SPECIFY THE NATURE OF VECTOR P *CTIM 19

```

C		*CTIM 36
C	K = DIMENSION OF P	*CTIM 37
C		*CTIM 38
C	P = VECTOR WITH VALUES OF MEAN, VARIANCE, OR OTHER PARAMETERES OF	*CTIM 39
C	THE DISTRIBUTION	*CTIM 40
C		*CTIM 41
C	VAL = THE RETURNED VARIATE	*CTIM 42
C		*CTIM 43
C	*****	*CTIM 44
C		CTIM 45
	SUBROUTINE CTIME(KODE,L,K,P,VAL)	CTIM 46
	DATA NSEED/567801/	CTIM 47
	DIMENSION P(50)	CTIM 48
	IF(KODE.EQ.10) GO TO 10	CTIM 49
	IF(KODE.EQ.20) GO TO 20	CTIM 50
	IF(KODE.EQ.30) GO TO 30	CTIM 51
	IF(KODE.EQ.40) GO TO 40	CTIM 52
	IF(KODE.EQ.50) GO TO 50	CTIM 53
	IF(KODE.EQ.60) GO TO 60	CTIM 54
	IF(KODE.EQ.70) GO TO 70	CTIM 55
	IF(KODE.EQ.80) GO TO 80	CTIM 56
	IF(KODE.EQ.90) GO TO 90	CTIM 57
	IF(KODE.EQ.100) GO TO 100	CTIM 58
	IF(KODE.EQ.110) GO TO 110	CTIM 59
	IF(KODE.EQ.120) GO TO 120	CTIM 60
		CTIM 61
	EXPONENTIAL	CTIM 62
		CTIM 63
	10 M=0	CTIM 64
	IF(K.EQ.1) P(2)=1.0	CTIM 65
	MIN=0.0	CTIM 66
	IF(L.EQ.0) EX=1.0/P(1)	CTIM 67
	IF(L.EQ.1) EX=P(1)	CTIM 68
	IF(K.EQ.1.OR.K.EQ.2) GO TO 11	CTIM 69
	MIN=P(3)	CTIM 70
	IF(K.EQ.4) MAX=P(4)	CTIM 71

11	IF(P(2).EQ.1.0) EX=EX-MIN	CTIM 72
	IF(P(2).EQ.0.0) EX=MAX-EX	CTIM 73
12	R=RAN(NSEED)	CTIM 74
	M=M+1	CTIM 75
	IF(M.EQ.100) CALL ERROR(10)	CTIM 76
	X=-EX*ALOG(R)	CTIM 77
	IF(P(2).EQ.1.0) X=X+MIN	CTIM 78
	IF(P(2).EQ.0.0) X=MAX-X	CTIM 79
	IF(K.EQ.1) GO TO 14	CTIM 80
	IF(K.EQ.3) GO TO 13	CTIM 81
	IF(X.GT.MAX) GO TO 12	CTIM 82
13	IF(X.LT.MIN) GO TO 12	CTIM 83
14	VAL=X	CTIM 84
	GO TO 1	CTIM 85
C	NORMAL	CTIM 86
C		CTIM 87
C		CTIM 88
20	M=0	CTIM 89
21	R=RAN(NSEED)/2.0	CTIM 90
	M=M+1	CTIM 91
	IF(M.EQ.100) CALL ERROR(20)	CTIM 92
	IF(R.EQ.0.0) GO TO 20	CTIM 93
	CALL NORM(P(1),P(2),Y)	CTIM 94
	X=Y	CTIM 95
	IF(K.EQ.2) GO TO 23	CTIM 96
	IF(K.EQ.3) GO TO 22	CTIM 97
	IF(X.GT.P(4)) GO TO 21	CTIM 98
22	IF(X.LT.P(3)) GO TO 21	CTIM 99
23	VAL=X	CTIM100
	GO TO 1	CTIM101
C	TRIANGULAR	CTIM102
C		CTIM103
C		CTIM104
30	M=0	CTIM105
	FP2=(P(2)-P(1))/(P(3)-P(1))	CTIM106
31	R=RAN(NSEED)	CTIM107

M=M+1

IF(M.EQ.100) CALL ERROR(30)

IF(R.GT.FP2) GO TO 32

X=P(1)+SQRT(R*(P(3)-P(1))*(P(2)-P(1)))

GO TO 33

32 X=P(1)+SQRT((1.-R)*(P(3)-P(1))*(P(3)-P(2)))

33 IF(K.EQ.3) GO TO 35

IF(K.EQ.4) GO TO 34

IF(X.GT.P(5)) GO TO 31

34 IF(X.LT.P(4)) GO TO 31

35 VAL=X

GO TO 1

TRAPEZODIAL

40 M=0

H=2./(-P(1)-P(2)+P(3)+P(4))

F2=H*0.5*(P(2)-P(1))

F3=H*0.5*(2.*P(3)-P(1)-P(2))

41 R=RAN(NSEED)

M=M+1

IF(M.EQ.100) CALL ERROR(40)

IF(R.GT.F3) X=P(4)-SQRT((P(3)-P(4))*(P(1)+P(2)-P(3)-P(4))*(1.-R))

IF(R.LE.F3) X=(P(1)-P(2))/2.+R/H

IF(R.LE.F2) X=P(1)+SQRT(2.*R*(P(2)-P(1))/H)

IF(K.EQ.4) GO TO 43

IF(K.EQ.5) GO TO 42

IF(X.GT.P(6)) GO TO 41

42 IF(X.LT.P(5)) GO TO 41

43 VAL=X

GO TO 1

LOGNORMAL

50 M=0

IF(L.EQ.1) GO TO 51

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SIGMA=ALOG(P(2)/P(1)**2.+1.)
GMU=ALOG(P(1))-0.5*ALOG(P(2)/P(1)**2.+1.)
EX=GMU
VX=SIGMA
GO TO 53
51 EX=P(1)
VX=P(2)
52 M=M+1
IF(M.EQ.100) CALL ERROR(50)
53 CALL NORM(0.0,1.0,Z)
X=EXP(SQRT(VX)*Z+EX)
IF(K.EQ.2) GO TO 55
IF(K.EQ.3) GO TO 54
IF(X.GT.P(4)) GO TO 52
54 IF(X.LT.P(3)) GO TO 52
55 VAL=X
GO TO 1

C
C      CHI-SQUARE
C
60 M=0
61 M=M+1
IF(M.EQ.100) CALL ERROR(60)
IA=P(1)
IB=IA+1
LB=P(1)*10.0-IA*10
LA=10-IB
Z=0.0
Y=0.0
IF(IA.EQ.0.0) GO TO 64
DO 63 I=1,LA
DO 62 J=1,IA
CALL NORM(0.0,1.0,XY)
Z=Z+XY**2
62 CONTINUE
63 CONTINUE
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```
      IF(IB.EQ.0.0) GO TO 666
64  DO 66 J=1,LB
      DO 65 I=1,IB
      CALL NORM(0.0,1.0,XY)
      Y=Y+XY**2
65  CONTINUE
66  CONTINUE
666 X=(Z+Y)/10.0
      IF(K.EQ.1) GO TO 68
      IF(K.EQ.2) GO TO 67
      IF(X.GT.P(3)) GO TO 61
67  IF(X.LT.P(2)) GO TO 61
68  VAL=X
      GO TO 1
```

C
C
C

BETA

```
70  M=0
      IF(L.EQ.0.AND.M.EQ.0) GO TO 72
      MM=2.0*((P(1)**2.-P(1)**3)/P(2)-P(1))
      N=2.0*(1.0-P(1))*(P(1)-P(1)**2-P(2))/P(2)
      GO TO 72
71  MM=P(1)
      N=P(2)
72  IF(M.LT.1) GO TO 78
      IF(N.LT.1) GO TO 79
73  M=M+1
      IF(M.EQ.100) CALL ERROR(71)
      SUMM=0.0
      DO 74 I=1,MM
      CALL NORM(0.0,1.0,U)
      SUMM=SUMM+U**2
74  CONTINUE
      SUMN=0.0
      DO 75 J=1,N
      CALL NORM(0.0,1.0,V)
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```
SUMN=SUMN+V**2
75 CONTINUE
  X=SUMM/(SUMM+SUMN)
  IF(K.EQ.2) GO TO 77
  IF(K.EQ.3) GO TO 76
  IF(X.GT.P(4)) GO TO 73
76 IF(X.LT.P(3)) GO TO 73
77 VAL=X
  GO TO 1
78 CALL ERROR(71)
  GO TO 1
79 CALL ERROR(72)
  GO TO 1

      F

80 M=0
  IF(L.EQ.0) GO TO 81
  N=2.*P(1)/(P(1)-1.0)
  IF(N.LT.4.0) CALL ERROR(83)
  MM=2.*P(1)**2*(N-2.)/((N-4.)*P(2)-2.*P(1)**2)
  GO TO 82
81 MM=P(1)
  N=P(2)
82 IF(M.LT.1) CALL ERROR(81)
  IF(N.LT.1) CALL ERROR(82)
83 M=M+1
  IF(M.EQ.100) CALL ERROR(80)
  SUMM=0.0
  DO 84 I=1,MM
    CALL NORM(0.0,1.0,U)
    SUMM=SUMM+U**2
84 CONTINUE
  SUMN=0.0
  DO 85 J=1,N
    CALL NORM(0.0,1.0,V)
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SUMN=SUMN+V**2
85 CONTINUE
X=(SUMM*N)/(SUMN*M)
IF(K.EQ.2) GO TO 87
IF(K.EQ.3) GO TO 86
IF(X.GT.P(4)) GO TO 83
86 IF(X.LT.P(3)) GO TO 83
87 VAL=X
GO TO 1

C
C      T-DISTRIBUTION
C

90 M=0
IF(L.EQ.0) GO TO 91
IF(P(1).GT.0.0.OR.P(1).LT.0.0) CALL ERROR(90)
N=2.0*P(2)/(P(2)-1.0)
GO TO 92
91 N=P(1)
92 IF(N.LT.1) CALL ERROR(91)
93 M=M+1
IF(M.EQ.100) CALL ERROR(92)
T=0.0
DO 94 I=1,N
CALL NORM(0.0,1.0,Y)
T=T+Y**2
94 CONTINUE
CALL NORM(0.0,1.0,Y)
X=Y/(SQRT(T/N))
IF(L.EQ.1.AND.K.EQ.2) GO TO 97
IF(L.EQ.0.AND.K.EQ.1) GO TO 97
IF(L.EQ.1.AND.K.EQ.3) GO TO 95
IF(L.EQ.0.AND.K.EQ.2) GO TO 96
IF(L.EQ.1) LL=4
IF(L.EQ.0) LL=3
IF(X.GT.P(LL)) GO TO 93
IF(L.EQ.0) GO TO 96
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95 LL=3
   IF(X.LT.P(LL)) GO TO 93
   GO TO 97
96 LL=2
   IF(X.LT.P(LL)) GO TO 93
97 VAL=X

C
C       Z-DISTRIBUTION
C
100 M=0
   IF(L.EQ.0) GO TO 103
101 M=M+1
   IF(M.EQ.100) CALL ERROR(101)
   IF(L.EQ.0.AND.M.GE.1) GO TO 104
   N=2.*P(1)/(P(1)-1.0)
   MM=2.*P(1)**2*(N-2.)/((N-4.)*P(2)-2.*P(1)**2)
   SUMN=0.0
   SUMM=0.0
   DO 102 I=1,N
     CALL NORM(0.0,1.0,U)
     SUMN=SUMN+U**2
102 CONTINUE
   DO 1100 J=1,MM
     CALL NORM(0.0,1.0,V)
     SUMM=SUMM+V**2
1100 CONTINUE
   GO TO 107
103 N=P(2)
   MM=P(1)
   IF(P(2).LT.1.0) CALL ERROR(101)
   IF(P(1).LT.1.0) CALL ERROR(102)
104 SUMM=0.0
   SUMN=0.0
   DO 105 I=1,MM
     CALL NORM(0.0,1.0,U)
     SUMM=SUMM+U**2
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```
105 CONTINUE
    DO 106 I=1,N
    CALL NORM(0.0,1.0,V)
    SUMN=SUMN+V**2
106 CONTINUE
    Q=P(1)
    S=P(2)
107 X=ALOG((N*SUMM)/(MM*SUMN))*0.5
    IF(K.EQ.2) GO TO 109
    IF(K.EQ.3) GO TO 108
    IF(X.GT.P(4)) GO TO 101
108 IF(X.LT.P(3)) GO TO 101
109 VAL=X
    GO TO 1
```

C
C
C

GAMMA

```
110 M=0
    IF(L.EQ.0) GO TO 111
    SCALE=P(1)/P(2)
    ALPHA=P(1)*SCALE
    GO TO 112
111 SCALE=P(2)
    ALPHA=P(1)
112 IAL=ALPHA
    IFRAC=(ALPHA-IAL)*10.
113 M=M+1
    IF(M.EQ.100) CALL ERROR(110)
    IF(IFRAC.EQ.0) GO TO 118
    IF(IAL.GT.C) GO TO 118
    Z=0.0
    B=1.-ALPHA
    IA=1000000./ALPHA
    IB=1000000./B
    C=IA/1000000.
    D=IB/1000000.
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      IF(IAL)116,116,114
114 DO 115 I=1,IAL
      Z=Z-ALOG(RAN(NSEED))
115 CONTINUE
116 Y=RAN(NSEED)**C
      W=RAN(NSEED)**D
      SUM=Y+W
      IF(SUM.LE.1.0) GO TO 117
      GO TO 116
117 BETA=Y/SUM
      V=-ALOG(RAN(NSEED))
      X=(Z+BETA*V)
      GO TO 1113
118 X=0.0
      PR1=IFRAC/10.
      R=RAN(NSEED)
      IF(R.GT.PR1) N=IAL
      IF(R.LE.PR1) N=IAL+1
      Z1=0.0
      DO 119 J=1,N
      Z1=Z1-ALOG(RAN(NSEED))/SCALE
119 CONTINUE
      X=Z1+P(3)
      GO TO 1117
1113 X=X/SCALE+P(3)
1117 IF(K.EQ.3) GO TO 1116
      IF(K.EQ.4) GO TO 1115
      IF(X.GT.P(5)) GO TO 113
1115 IF(X.LT.P(4)) GO TO 113
1116 VAL=X
      GO TO 1
C
C      PARETO
C
120 M=0
      IF(L.EQ.0) GO TO 121
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      A=1.0+SQRT(4.0+4.0*P(1)**2/P(2))/2.0
      PK=P(1)*(A-1.0)/A
      GO TO 122
121  A=P(1)
      PK=P(2)
122  IF(A.LE.0.0) GO TO 120
123  M=M+1
      IF(M.EQ.100) CALL ERROR(121)
      R=RAN(NSEED)
      X=EXP(ALOG(PK)-ALOG(R)/A)
      IF(X.LT.PK) CALL ERROR(122)
      IF(K.EQ.2) GO TO 125
      IF(K.EQ.3) GO TO 124
      IF(X.GT.P(4)) GO TO 123
124  IF(X.LT.P(3)) GO TO 123
125  VAL=X
      GO TO 1
1  RETURN
END

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C*****ERRO 1
C      TO GIVE THE K-TH ERROR MESSAGE      *ERRO 2
C                                           *ERRO 3
C      K = THE NUMBER FOR THE ERROR      *ERRO 4
C                                           *ERRO 5
C*****ERRO 6
C                                           ERRO 7
      SUBROUTINE ERROR(K)                  ERRO 8
      IF(K.EQ.10) WRITE(6,10)              ERRO 9
      IF(K.EQ.20) WRITE(6,20)              ERRO 10
      IF(K.EQ.30) WRITE(6,30)              ERRO 11
      IF(K.EQ.40) WRITE(6,40)              ERRO 12
      IF(K.EQ.50) WRITE(6,50)              ERRO 13
      IF(K.EQ.60) WRITE(6,60)              ERRO 14
      IF(K.EQ.70) WRITE(6,70)              ERRO 15
      IF(K.EQ.71) WRITE(6,71)              ERRO 16
      IF(K.EQ.72) WRITE(6,72)              ERRO 17
      IF(K.EQ.80) WRITE(6,80)              ERRO 18
      IF(K.EQ.81) WRITE(6,81)              ERRO 19
      IF(K.EQ.82) WRITE(6,82)              ERRO 20
      IF(K.EQ.83) WRITE(6,83)              ERRO 21
      IF(K.EQ.90) WRITE(6,90)              ERRO 22
      IF(K.EQ.100) WRITE(6,100)            ERRO 23
      IF(K.EQ.101) WRITE(6,101)            ERRO 24
      IF(K.EQ.102) WRITE(6,102)            ERRO 25
      IF(K.EQ.110) WRITE(6,110)            ERRO 26
      IF(K.EQ.120) WRITE(6,120)            ERRO 27
      IF(K.EQ.130) WRITE(6,130)            ERRO 28
10  FORMAT('  ERROR IN EXPONENTIAL DISTRIBUTION  ' ) ERRO 29
20  FORMAT('  ERROR IN NORMAL DISTRIBUTION  ' ) ERRO 30
30  FORMAT('  ERROR IN TRIANGULAR DISTRIBUTION  ' ) ERRO 31
40  FORMAT('  ERROR IN TRAPEZODIAL DISTRIBUTION  ' ) ERRO 32
50  FORMAT('  ERROR IN LOGNORMAL DISTRIBUTION  ' ) ERRO 33
60  FORMAT('  ERROR IN CHI SQUARE DISTRIBUTION  ' ) ERRO 34
70  FORMAT('  ERROR IN BETA DISTRIBUTION  ' ) ERRO 35

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		PAGE
71	FORMAT(' DEGREE OF FREEDOM M OF BETA DISTRIBUTION IS LESS THAN + ONE ')	ERRO 36 ERRO 37
72	FORMAT(' DEGREE OF FREEDOM N OF BETA DISTRIBUTION IS LESS THAN + ONE ')	ERRO 38 ERRO 39
80	FORMAT(' ERROR IN F DISTRIBUTION ')	ERRO 40
81	FORMAT(' THE DEGREE OF FREEDOM M OF F DISTRIBUTION IS LESS THAN + THAN ONE ')	ERRO 41 ERRO 42
82	FORMAT(' THE DEGREE OF FREEDOM N OF F DISTRIBUTION IS LESS THAN + THAN ONE ')	ERRO 43 ERRO 44
83	FORMAT(' DEGREE OF FREEDOM N OF F DISTRIBUTION IS LESS THAN 4 ')	ERRO 45
90	FORMAT(' ERROR IN T DISTRIBUTION ')	ERRO 46
100	FORMAT(' ERROR IN Z DISTRIBUTION ')	ERRO 47
101	FORMAT(' DEGREE OF FREEDOM N OF Z DISTRIBUTION IS LESS THAN 1 ')	ERRO 48
102	FORMAT(' DEGREE OF FREEDOM M OF Z DISTRIBUTION IS LESS THAN 1 ')	ERRO 49
110	FORMAT(' ERROR IN GAMMA DISTRIBUTION ')	ERRO 50
120	FORMAT(' ERROR IN ERLANG DISTRIBUTION ')	ERRO 51
130	FORMAT(' ERROR IN PARETO DISTRIBUTION ')	ERRO 52
	RETURN	ERRO 53
	END	ERRO 54

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C*****NORM
C                                          *NORM
C      TO RETURN A RANDOM NORMAL VARIATE      *NORM
C                                          *NORM
C      EX = THE MEAN OF THE DISTRIBUTION      *NORM
C                                          *NORM
C      VX = THE VARIANCE OF THE DISTRIBUTION  *NORM
C                                          *NORM
C      VAL = THE RETURNED NORMAL VARIATE      *NORM
C                                          *NORM
C*****NORM 1
C                                          NORM 1
C      SUBROUTINE NORM(EX,VX,VAL)              NORM 1
C      DATA NSEED/123123/                     NORM 1
3  SUM=0.0                                     NORM 1
C      DO 1 I=1,12                             NORM 1
C      SUM=SUM+RAN(NSEED)                       NORM 1
1  CONTINUE                                   NORM 1
C      IF(SUM.LT.2.0.AND.SUM.GT.10.0) GO TO 3   NORM 1
C      R=(SUM-6.)/4.                             NORM 2
C      VAL=3.949846138*R+0.252408784*R**3+0.076542812*R**5+0.008355968*R  NORM 2
C      Q**7+0.029899776*R**9                     NORM 2
C      VAL=VAL*SQRT(VX)+EX                       NORM 2
C      RETURN                                    NORM 2
C      END                                        NORM 2

```


BIBLIOGRAPHY

1. Abramowitz and Stegun Handbook of Mathematical Functions, New York: Dover Pub. Co. (1970).
2. Blom, Gunnar Statistical Estimates and Transformed Beta Variables, New York: John Wiley and Sons (1958).
3. Bofinger, Eve and V.J. "On a Periodic Property of Pseudo-Random Sequences," Asso. for Comp. Mach. J., Vol.5 (1958), pp.261-5.
4. Box, G.E.P. and Muller, M.E. "A Note on the Generation of Random Normal Deviates," Ann. Math. Stat., Vol.28 (1958), pp.610-11.
5. Brown, Robert G. Statistical Forecasting for Inventory Control, New York: McCraw-Hill Book Co. (1959).
6. Busch, R.R. and Mosteller, F. Stochastic Models for Learning. New York: John Wiley and Sons (1955).
7. Butler, James W. "Machine Sampling from Given Probability Distribution" in Symposium on Monte Carlo Methods, Edited by H.A. Meyer, New York: John Wiley and Sons (1956).
8. Cacoullos, Theophilos "A Relation Between t and F -Distributions," Amer. Stat. Asso. J., Vol.60 (1965), pp.528-31.
9. Cannon, L.E. III Pseudo Random Number Generators for Statistical Applications, Masters Thesis, University of Georgia, Athens, Georgia (1971).
10. Certaine, J. "On Sequences of Pseudo-Random Numbers of Maximal Length," Asso. for Comp. Mach. J., Vol.5 (1958), pp.353-56.
11. Feller, William An Introduction to Probability Theory and Its Applications, New York: John Wiley and Sons (1950).
12. Greenberger, Martin "Notes on a New Pseudo-Random Number Generator," Asso. for Comp. Mach. J., Vol.8 (1961), pp.163-7.
13. Gorenstein, S. "Testing a Random Number Generator," Comm. of the ACM, Vol.10 (1967), pp.111-8.

14. Hald, A. Statistical Theory with Engineering Applications, New York: John Wiley and Sons, Inc. (1952).
15. Hasting, C. Jr. Approximations for Digital Computers, Princeton: University Press (1955).
16. Hoefding, Wassily "On the Distribution of the Expected Values of the Order Statistics," Ann. of Math. Stat., Vol.24 (1953), pp.93-100.
17. Knuth, Donald E. The Art of Computer Programming 2d ed., Mass: Addison Wesley Pub. Co. (1973).
18. Kotz, Samuel Distribution in Statistics, Boston: Houghton Mifflin (1969).
19. Lehmer, D.H. "Mathematical Methods in Large Scale Computing Units," Symposium on Large-Scale Digital Calculating Machinery, Harvard University (1949), pp. 141-6.
20. MacLaren, M. Donald and Marsaglia, George "Uniform Random Number Generators," Asso. for Comp. Mach. J., Vol.12 (1965), pp.83-9.
21. Marsaglia, G., MacLaren, M.D. and Bray, T.A. "A Fast Procedure for Generating Normal Random Variables," Comm. of the ACM, Vol.7 (1964), pp.4-10.
22. Marsaglia, G. and Bray, T.A. "One-Line Random Number Generators and Their Use in Combinations," Comm. of the ACM, Vol.11 (1968), pp.757-9.
23. Marsaglia, G. "Generating Discrete Random Variables in a Computer," Comm. of the ACM, Vol. (1963), pp.37-8.
24. Marsaglia, G. "Generating Exponential Random Variables," Ann. Math. Stat., Vol.32 (1961), pp.899.
25. Mathur, R.K. "A Note on the Wilson-Hilferty Transformation of Chi-square," Bull. Calcutta Stat. Assoc., Vol.10 (1961), pp.103-5.
26. Metropolis, N. "Phase Shifts-Middle Squares-Wave Equation" in Symposium on Monte Carlo Methods, Edited by H.A. Meyer, New York: John Wiley and Sons (1956).

27. Meyer, P.L. Introductory Probability and Statistical Applications, Addison Wesley Pub. Co. (1970).
28. Muller, Mervin E. "A Comparison of Methods for Generating Normal Deviates on Digital Computers," Asso. for Comp. Mach. J., Vol.6 (1959), pp.376-83.
29. Muller, Mervin E. "An Inverse of the Generation of Random Normal Deviates on Large-Scale Computers," Math. Tables Aids Comp. 12 (1958), pp.167-174.
30. Naylor, Balintfy, Burdick and Chu Computer Simulation Techniques, New York: John Wiley and Sons (1966).
31. Newmann, T.G. and Odell P.L. The Generation of Random Variates, New York: Hafner Pub. Co. (1971).
32. Norman, J.E. and Cannon L.E. "A Computer Program for the Generation of Random Variables from any Discrete Distribution," J. Stat. Comp. Simul. Vol.1 (1972), pp.331-48.
33. Parzen, E. Modern Probability Theory and Its Applications, New York: John Wiley and Sons (1960).
34. Sacks, A.J. A Program to Evaluate Uniform Random Number Generators, Masters Thesis, University of Houston (1974).
35. Rothlf, F.J. and Sokal, R.R. Statistical Tables, W.H. Freeman and Co. (1969).
36. Phillips, Don T. and Beightler, Charles S. "Procedures for Generating Gamma Variates with Non-integer Parameters Sets," J. Stat. Comput. Simul., Vol.1 (1972), pp.197-208.
37. Smith, Marvin B. Probability Models for Petroleum Investment Decisions, Paper SPE 2587 of Society of Petroleum Engineers of AIME (1969).
38. Taussky, O. and Todd, J. "Generation of Pseudo Random Numbers" in Symposium on Monte Carlo Methods, Edited by H.A. Meyer, New York: John Wiley and Sons (1956).
39. Tocher, K.D. The Art of Simulation, Princeton: D. Van Nostrand Co. Inc. (1963).

40. Teichroew, Dan Distribution Sampling with high Speed Computers, Ph.D. Thesis, University of North Carolina (1953).
41. Von Neumann, John Various Techniques used in Connection with Random Digits, paper No.13 in "Monte Carlo Methods," NBS Applied Mathematics Series No.12, U.S. Govt. Printing Office (1951).
42. Votaw, D.F. Jr. and Rafferty, J.A. "High Speed Sampling" MTAC, Vol.5 (1951), pp.1-8.
43. Wagner, Harvey M. Principles of Operations Research with Applications to Managerial Decisions, New Jersey: Prentice-Hall Inc. (1969).
44. Johnk, M.D. "Erzeugung von Betaverteilten and Gamma Verteilten Zufallszahlen", Metrika, Vol.8, No.1, (1964) PP.5-15.